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Authors

Watson, Hunna J
Yilmaz, Zeynep
Thornton, Laura M
[et al.](#)

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Genome-wide Association Study Identifies Eight Risk Loci and Implicates Metabo-Psychiatric Origins for Anorexia Nervosa

A full list of authors and affiliations appears at the end of the article.

Abstract

Characterized primarily by low BMI, anorexia nervosa is a complex and serious illness¹, affecting 0.9-4% of women and 0.3% of men²⁻⁴, with twin-based heritability estimates of 50-60%⁵.

*Correspondence and requests for materials should be addressed to C.M.B. or G.B. cbulik@med.unc.edu, cynthia_bulik@med.unc.edu.

Author contributions

C.M.B. and P.F.S. conceived and designed the study. L.T., C.M.B., and G.B. performed overall study coordination. C.M.B. was lead PI of ANGI. P.F.S. was Co-Investigator of ANGI. N.G.M., M.L., and P.B.M. were site PIs of ANGI. H.J.W., Z.Y., J.R.I.C., C.H., J.B., H.A.G., S.Y., V.M.L., M.M., P.G.R. and S.E.M. performed the statistical analyses. H.J.W., Z.Y., C.H., J.R.I.C., H.A.G., J.B., A.H., P.G.R., P.F.S., G.B. and C.M.B. comprised the writing group. C.M.B. and G.B. were PGC-ED co-chairs. S.R. provided statistical consultation. A.H. assisted with data interpretation. A.W.B., C.M.B., J.J., M.K., K.M.K., P.L., G.M., C.N., R.P., L.T., and T.D.W. collected and managed the ANGI samples at sites and assisted with site-specific study co-ordination. A.W.B., J.M.B., H.B., S.C., K.A.H., L.J.H., C.J., A.S.K., W.K., J.M., C.M.O., J.F.P., N.L.P., M.S., T.W., D.C.W., and D.B.W. provided ANGI controls and extra samples. L.E.D provided data expertise. S.G., J.G., A.K.H., A.J., K.M.K., J.T.L., R.P., and L.P. contributed to the ANGI study. S.G., J.G., K.K., J.T.L., M.M., S.M., and L.P. were ANGI site analysts. K.B.H. and K.L.P. provided additional analysis for some secondary analyses. G.W.M., T.D.W., A.B., P.L., and C.N. were ANGI investigators. J.J. and M.K. assisted with ANGI recruitment in NZ. C.M.B., G.B., and P.F.S. supervised the study. H.J.W., C.M.B., Z.Y., C.H., G.B., J.R.I.C., H.A.G., S.Y., J.B., P.F.S., and P.G. wrote the manuscript. PGC-ED members and other individuals contributed to sample acquisition and made individual data from subjects available: R.A.H.A., L.A. T.A., O.A.A., J.H.B., A.W.B., W.H.B., A.B., I.B., C.B., J.M.B., H.B., G.B., K. B., C.M.B., R.B., M.C., S.C., M.C., J.R.I.C., R.D.C., P.C., S.C., S.C., J.C., U.N.D., O.S.P.D., M.D., G.D., D.D., J.E.D., D.M.D., D.D., C.D., M.D., E.D.M., K.E., S.E., G.E., T.E., X.E., A.F., A.F., F.F., M.M.F., K.F., M.F., L.F., A.J.F., M.F., S.G., I.G., J.G., F.G., S.G., P.G., M.G.M., J.G., S.G., K.A.H., K.H., J.H., J.H., S.G.H., A.K.H., S.H., B.H., W.H., A.H., L.J.H., J.I.H., H.I., H.L., V.J., S.J., C.J., J.J., A.J., A.J., G.K., D.K., A.S.K., J.K., L.K., A.K., M.J.H.K., W.K., J.L.K., M.K., A.K., K.K., Y.K., L.K., G.S.K., M.C.L., M.L., S.L., R.D.L., P.L., L.L., B.L., J.L., J.L., P.M., M.M., K.M., S.M., C.M., N.G.M., M.M., S.M., P.M., A.M., I.M., N.M., J.M., A.M.M., P.M., P.M., M.A.M., B.N., M.N., C.N., I.N., C.M.O., J.K.O., R.A.O., L.P., A.P., J.P., H.P., N.L.P., J.F.P., D.P., R.R., A.R., N.R., T.R., V.R., S.R., F.R., M.R., A.R., D.R., F.R., P.S., S.W.S., U.S., A.S., J.S., L.S., P.E.S., M.C.T.S.L., A.S., S.S., M.S., P.F.S., B., J.P.S., I.T., E.T., A.T., F.T., J.T., A.T., M.T., K.T., A.A.V., E.F.V., T.D.W., G.W., E.W., H.J.W., T.W., D.C.W., E.W., D.B.W., G.S., S.Z., and S.Z. All authors critically reviewed the manuscript.

^aThese authors contributed equally to this work

^bThe members of this consortium are listed in the Supplementary Note

^cThese authors jointly directed this project

Competing interests

The authors report the following potential competing interests. O.A.A. received a speaker's honorarium from Lundbeck. G.B. received grant funding and consultancy fees in preclinical genetics from Eli Lilly, consultancy fees from Otsuka and has received honoraria from Illumina. C.M.B. is a grant recipient from Shire Pharmaceuticals and served on Shire Scientific Advisory Board; she receives author royalties from Pearson. D.D. served as a speaker and on advisory boards, and has received consultancy fees for participation in research from various pharmaceutical industry companies including: AstraZeneca, Boehringer, Bristol Myers Squibb, Eli Lilly, Genesis Pharma, GlaxoSmithKline, Janssen, Lundbeck, Organon, Sanofi, UniPharma, and Wyeth; he has received unrestricted grants from Lilly and AstraZeneca as director of the Sleep Research Unit of Eginition Hospital (National and Kapodistrian University of Athens, Greece). J.I.H. has received grant support from Shire and Sunovion, and has received consulting fees from DiaMentis, Shire, and Sunovion. A.S.K. is a member of the Shire Canadian BED Advisory Board and is on the steering committee for the Shire B/educated Educational Symposium: June 15-16, 2018. J.L.K. served as an unpaid member of the scientific advisory board of AssurexHealth Inc. M.L. declares that, over the past 36 months, he has received lecture honoraria from Lundbeck and served as scientific consultant for EPID Research Oy. No other equity ownership, profit-sharing agreements, royalties, or patent. P.F.S. is on the Lundbeck advisory committee and is a Lundbeck grant recipient; he has served on the scientific advisory board for Pfizer, has received a consultation fee from Element Genomics, and a speaker reimbursement fee from Roche. J.T. has received an honorarium for participation in an EAP meeting and has received royalties from several books from Routledge, Wiley, and Oxford University press. T.W. has acted as a lecturer and scientific advisor to H. Lundbeck A/S. All other authors have no conflicts of interest to disclose.

URLs. GCTA, <http://cnsngenomics.com/software/gcta>; GSMR, <http://cnsngenomics.com/software/gsmr>; LDSC, <https://github.com/bulik/ldsc>; MAGMA, <http://ctg.cncr.nl/software/magma>.

Mortality rates are higher than other psychiatric disorders⁶, and outcomes are unacceptably poor⁷. Combining data from the Anorexia Nervosa Genetics Initiative (ANGI)^{8,9} and the Eating Disorders Working Group of the Psychiatric Genomics Consortium (PGC-ED), we conducted a genome-wide association study (GWAS) of 16,992 anorexia nervosa cases and 55,525 controls, identifying eight significant loci. The genetic architecture of anorexia nervosa mirrors its clinical presentation showing significant genetic correlations with psychiatric disorders, physical activity, metabolic (including glycaemic), lipid, and anthropometric traits, independent of the effects of common variants associated with BMI. Results further encourage a reconceptualization of anorexia nervosa as a metabo-psychiatric disorder. Explicating the metabolic component is a critical direction, and attention to both psychiatric and metabolic components may be key to improving outcomes.

The first PGC-ED GWAS (3,495 cases, 10,982 controls) estimated the common genetic variant-based heritability of anorexia nervosa as ~20%, identified the first genome-wide significant locus, and reported significant genetic correlations (r_g) between anorexia nervosa and psychiatric and metabolic/anthropometric phenotypes¹⁰. These r_g pointed toward metabolic etiological factors, as they are robust to reverse causation although they could be mediated associations¹¹ or reflect confounding processes¹². To advance genomic discovery in anorexia nervosa and further explore genetic correlations, we combined samples from ANGI^{8,9}, the Genetic Consortium for Anorexia Nervosa (GCAN)/Wellcome Trust Case Control Consortium-3 (WTCCC-3)¹³, and the UK Biobank¹⁴, quadrupling our sample size.

Our GWAS meta-analysis included 33 datasets comprising 16,992 cases and 55,525 controls of European ancestry from 17 countries (Supplementary Tables 1-4). We had 80% power to detect an odds ratio (OR) of 1.09-1.19 (additive model, 0.9% lifetime risk, $\alpha = 5 \times 10^{-8}$, MAF 0.05–0.5). Typical of complex trait GWAS, we observed test statistic inflation ($\lambda = 1.22$) consistent with polygenicity, with no evidence of significant population stratification according to the LD intercept and attenuation ratio (Supplementary Results; Supplementary Fig. 1). Meta-analysis results were completed for autosomes and the X chromosome. We identified eight loci exceeding genome-wide significance ($P < 5 \times 10^{-8}$; Table 1 for loci; Fig. 1 for the Manhattan plot; Supplementary Figs. 2a-h and 3a-h for the forest and region plots). Many were near the threshold for significance, and no significant heterogeneity of SNP associations across cohorts was detected ($P = 0.15$ -0.64; Supplementary Figs. 2a-h). Conditional and joint analysis (GCTA-COJO)¹⁵ confirmed independence of the lead SNPs within the significant loci (Supplementary Table 5). The eight loci were annotated to identify known protein-coding genes (Supplementary Table 6; Supplementary Table 7 reports a gene look-up restricted to the single-gene loci). The previously reported PGC-ED genome-wide significant variant (rs4622308)¹⁰ on 12q13.2 did not reach genome-wide significance ($P = 7.02 \times 10^{-5}$); however, between-cohort heterogeneity was apparent ($I^2 = 53.7$; Supplementary Fig. 4 and Supplementary Results). The OR was in the same direction in 22 (67%) of the cohorts ($z = 2.00$, $P = 0.05$, 2-tailed).

Although GWAS findings are informative genome-wide, identifying strong hypotheses about their connections to specific genes is not straightforward. We evaluated three ways to “connect” anorexia nervosa GWAS loci to genes: regulatory chromatin interactions;

relationship to brain expression QTLs (eQTLs; using a superset of CommonMind¹⁶ and GTE_x¹⁷) and the standard approach of gene location within a GWAS locus. The significant anorexia nervosa loci implicated 121 brain-expressed genes, 74% by location, 55% by adult brain eQTL, 93% by regulatory chromatin interaction, and 58 genes by all three methods. Supplementary Figs. 5a-h show the eight GWAS loci, GENCODE gene models, adult brain regulatory chromatin interactions, brain eQTLs, and functional genomic annotations.

Four single-gene loci were confirmed by eQTL, chromatin interaction, or both. These were the locus-intersecting genes *CADMI* (locus 2 chr11:114.9-115.4 Mb, Supplementary Fig. 5b), *MGMT* (locus 4, chr10:131.2-131.4 Mb, Supplementary Fig. 5d), *FOXPI* (locus 5, chr3:70.6-71.0 Mb, Supplementary Fig. 5e) and *PTBP2* (locus 6, chr1:96.6-97.2 Mb, Supplementary Fig. 5f). For locus 5, eQTL data implicated a distal gene, *GPR27*. One intergenic locus (locus 7, chr5:24.9-25.3 Mb, Supplementary Fig. 5g) had no eQTL or chromatin interactions whereas the other intergenic locus (locus 8, chr3:93.9-95.0 Mb, Supplementary Fig. 5h) had eQTL connections to *PROS1* and *ARL13B*. Two complex multigenic loci had many brain-expressed genes and dense chromatin and eQTL interactions that precluded identification of any single gene (locus 1, chr3:47.5-51.3 Mb; locus 3, chr2:53.8-54.3 Mb, Supplementary Figs. 5a and 5c). The clearest evidence and connections were for the single-gene loci intersecting *CADMI*, *MGMT*, *FOXPI*, and *PTBP2* and we conclude these genes may play a role in anorexia nervosa etiology (Supplementary Results).

Supplementary Table 8 presents multi-trait analysis (GCTA-mtCOJO¹⁸ conditioning our genome-wide significant SNPs on associated variants in GWAS of BMI, type 2 diabetes, education years, HDL cholesterol, neuroticism, and schizophrenia. Seven loci appear to be independent. Locus 2 on chr11 may not be unique to anorexia nervosa and may be driven by genetic variation also associated with type 2 diabetes.

Liability-scale SNP heritability (SNP- h^2) was estimated with LD score regression (LDSC)^{19,20}. Assuming a lifetime prevalence of 0.9-4%²⁻⁴, SNP- h^2 was 11-17% (s.e. = 1%), supporting the polygenic nature of anorexia nervosa. Polygenic risk score (PRS) analyses using a leave-one-out approach indicated that the PRS captures ~1.7% of the phenotypic variance on the liability scale for discovery $P=0.5$. We did not observe differences in polygenic architecture between anorexia nervosa subtypes with binge eating (2,381 cases, 10,249 controls) or without (2,262 cases, 10,254 controls) or between males (447 cases, 20,347 controls) and females (14,898 cases, 27,545 controls) (Methods, Supplementary Results, Supplementary Fig. 6, Supplementary Table 9). Similar to females, males in the highest PRS decile had 4.13 (95% CI: 2.58-6.62) times the odds of anorexia nervosa than those in the lowest decile. Confirmation of these results requires larger samples.

We tested SNP-based genetic correlations (SNP- r_g) with external traits using bivariate LDSC^{19,20}. Bonferroni-significant SNP- r_g assorted into five trait categories: psychiatric and personality; physical activity; anthropometric; metabolic; and educational attainment (Supplementary Table 10). Fig. 2 presents Bonferroni-corrected positive SNP- r_g with OCD (SNP- $r_g \pm$ s.e. = 0.45 ± 0.08 ; $P=4.97 \times 10^{-9}$), MDD (0.28 ± 0.07 ; $P=8.95 \times 10^{-5}$), anxiety disorders (0.25 ± 0.05 ; $P=8.90 \times 10^{-8}$), and schizophrenia (0.25 ± 0.03 ; $P=4.61 \times 10^{-18}$). This pattern reflects observed comorbidities in clinical and epidemiological studies^{21,22}. The

newly-identified positive SNP- r_g with physical activity (0.17 ± 0.05 ; $P = 1.00 \times 10^{-4}$) encourages further exploration of the refractory symptom of pathologically elevated activity in anorexia nervosa²³. We note that the significant SNP- r_g of anorexia nervosa with educational attainment (0.25 ± 0.03 ; $P = 1.69 \times 10^{-15}$) and related constructs was not seen for IQ²⁴.

Expanding our previous observations¹⁰, we present a palette of metabolic and anthropometric r_g with anorexia nervosa more pronounced than in other psychiatric disorders. We observed significant negative SNP- r_g with fat mass (-0.33 ± 0.03 ; $P = 7.23 \times 10^{-25}$), fat-free mass (-0.12 ± 0.03 ; $P = 4.65 \times 10^{-5}$), BMI (-0.32 ± 0.03 ; $P = 8.93 \times 10^{-25}$), obesity (-0.22 ± 0.03 ; $P = 2.96 \times 10^{-11}$), type 2 diabetes (-0.22 ± 0.05 ; $P = 3.82 \times 10^{-5}$), fasting insulin (-0.24 ± 0.06 ; $P = 2.31 \times 10^{-5}$), insulin resistance (-0.29 ± 0.07 ; $P = 2.83 \times 10^{-5}$), and leptin (-0.26 ± 0.06 ; $P = 4.98 \times 10^{-5}$), and a significant positive SNP- r_g with HDL cholesterol (0.21 ± 0.04 ; $P = 3.08 \times 10^{-7}$).

Systems biology analyses of our results revealed preliminarily interesting results (Methods, Supplementary Tables 11-13, Supplementary Figs. 7-15). Gene-wise analysis with MAGMA prioritized 79 Bonferroni-significant genes, most within the multigenic locus on chr3 (Supplementary Table 11). MAGMA indicated an association with *NCAMI* (Supplementary Table 11) the expression of which increases in response to food restriction in a rodent activity-based anorexia nervosa model²⁵. Partitioned heritability analysis showed, as with other GWAS²⁶, considerable enrichment of SNP- h^2 in conserved regions (fold enrichment = 24.97, s.e. = 3.29, $P = 3.32 \times 10^{-11}$; Supplementary Fig. 7)²⁷. Cell type group-specific annotations revealed that the overall SNP- h^2 is significantly enriched for CNS tissue (Supplementary Fig. 8). One biological pathway was significant:

GO:positive_regulation_of_embryonic_development (32 genes, $P = 1.39 \times 10^{-7}$; Supplementary Table 12), which contains two Bonferroni-significant genes on chr3, *CTNNB1* and *DAG1*. *CTNNB1* encodes catenin beta-1, which is part of adherens junctions, and *DAG1* encodes dystroglycan, a receptor which binds extracellular matrix proteins²⁸. *DAG1* falls within locus 1 (47.5-51.3 Mb). This pathway points to a potential role of developmental processes in the etiology of this complex phenotype (although this is currently speculative). Genes associated with anorexia nervosa were enriched for expression in most brain tissues, particularly the cerebellum, which has a notably high proportion of neurons²⁹ (Supplementary Fig. 9). Among 24 brain cell types from mouse brain, significant enrichment was found for medium spiny neurons and pyramidal neurons from hippocampal CA1 (Supplementary Fig. 10). Both medium spiny and pyramidal neurons are linked to feeding behaviors including food motivation and reward^{30,31} (Supplementary Results). Using PrediXcan (Supplementary Methods), 36 genes were predicted to be differentially expressed in GTEx tissues or blood (Supplementary Table 13) with the expression of *MGMT* predicted to be downregulated in the caudate. We cautiously note that these results represent the first indications of specific pathways, tissues, and cell types that may mediate genetic risk for anorexia nervosa.

Because low BMI is pathognomonic of anorexia nervosa, we investigated the extent to which genetic variants associated with BMI accounted for genetic correlations with metabolic and anthropometric traits. First, covarying for the genetic associations of BMI

(Methods) led to a mild but statistically non-significant attenuation of the SNP- r_g between anorexia nervosa and fasting insulin, leptin, insulin resistance, type 2 diabetes, and HDL cholesterol (Supplementary Tables 14-15), suggesting that anorexia nervosa shares genetic variation with these metabolic phenotypes that may be independent of BMI. Second, we investigated bidirectional causality using generalized summary data-based Mendelian randomization¹⁸. GSMR analyses indicate a significant bidirectional causal relationship such that anorexia nervosa risk-increasing alleles may increase risk for low BMI and BMI-lowering alleles may increase the risk of anorexia nervosa (Supplementary Table 16). It is important to note that having only eight genome-wide significant loci for anorexia nervosa render this analysis marginally powered in the direction of anorexia nervosa to BMI, although this analysis is well powered in the direction of BMI to anorexia nervosa.

Replication is challenging with GWAS of low prevalence conditions like anorexia nervosa, as replication samples must be sufficiently powered to detect the initial findings. We included all available samples in our analysis to maximize chances of reaching the GWAS inflection point, after which there might be a linear increase in “hits”³². The PRS leave-one-out analyses provide evidence of replication by demonstrating a higher burden of anorexia nervosa common risk variants in cases, compared with controls, across all the cohorts (Supplementary Fig. 16).

In conclusion, we report multiple genetic loci alongside promising clinical and functional analyses and enrichments. The increased sample size in the present GWAS has allowed us to characterize more fully the metabolic contribution to anorexia nervosa than our previous report¹⁰ by revealing significant r_g with metabolism related phenotypes including glycemic and anthropometric traits and by demonstrating that the effect is robust to correction for the effects of common variants significantly associated with BMI. Low BMI has traditionally been viewed as a consequence of the psychological features of anorexia nervosa (i.e., drive for thinness and body dissatisfaction). This perspective has failed to yield interventions that reliably lead to sustained weight gain and psychological recovery⁷. Fundamental metabolic dysregulation may contribute to the exceptional difficulty that individuals with anorexia nervosa have in maintaining a healthy BMI (even after therapeutic renourishment). Our results encourage consideration of both metabolic and psychological drivers of anorexia nervosa when exploring new avenues for treating this frequently lethal illness.

Methods

Samples and study design.

Thirty-three datasets with 16,992 anorexia nervosa cases and 55,525 controls were included in the primary GWAS. We included individuals from the Eating Disorders Working Group of the Psychiatric Genomics Consortium (PGC-ED) Freeze 1¹⁰; newly collected samples from the Anorexia Nervosa Genetics Initiative (ANGI)^{8,9}; archived samples from the Genetic Consortium for Anorexia Nervosa (GCAN)/Wellcome Trust Case Control Consortium-3 (WTCCC3)¹³; anorexia nervosa samples from UK Biobank¹⁴; and additional controls from Poland. Case definitions established a lifetime diagnosis of anorexia nervosa via hospital or register records, structured clinical interviews, or on-line questionnaires based on standardized criteria (DSM-III-R, DSM-IV, ICD-8, ICD-9, or ICD-10), whereas in the UK

Biobank cases self-reported a diagnosis of anorexia nervosa. Controls were carefully matched for ancestry, and some, but not all control cohorts were screened for lifetime eating and/or some or all psychiatric disorders. Given the relative rarity of anorexia nervosa, large unscreened control cohorts were deemed appropriate for inclusion³³.

The cohorts are detailed in the Supplement. Ethical approvals and consent forms were reviewed and archived for all participating cohorts (see Supplementary Methods ANGI-DK for Danish methods). Summary details about ascertainment (Supplementary Table 2), the genotyping platforms used (Supplementary Table 3), and genotype availability (Supplementary Table 4) can be accessed in the Supplement.

Statistical analysis.

Data processing and analysis were done on the Lisa Compute Cluster hosted by SURFsara (<http://www.surfsara.nl>) and the GenomeDK high-performance computing cluster (<http://genome.au.dk>).

Meta-analysis of genome-wide association data.—Quality control (QC), imputation, GWAS, and meta-analysis followed the standardized pipeline of the PGC, Ricopili (Rapid Imputation Consortium Pipeline). Ricopili versions used were 2017_Oct_11.002 and 2017_Nov_30.003. QC included SNP and sample QC, population stratification and ancestry outliers, and familial and cryptic relatedness. Further information about the Ricopili pipeline is available from the website (<https://sites.google.com/a/broadinstitute.org/ricopili>) and GitHub repository (https://github.com/Nealelab/ricopili/tree/master/rp_bin). Further details of the QC procedures can be found in the Supplementary Methods.

Imputation.—Imputation of SNPs and insertions-deletions was based on the 1000 Genomes Phase 3 (<http://www.internationalgenome.org>) data³⁴.

GWAS.—GWASs were conducted separately for each cohort using imputed variant dosages and an additive model. Covariates nominally associated with the phenotype in univariate analysis ($P < 0.05$) and five ancestry PCs were included in GWAS (Supplementary Table 18). These analyses used the tests and methods programmed in the Ricopili pipeline. Genomic inflation factors (λ) of the final datasets indicated no evidence of inflation of the test statistics due to population stratification or other sources (Supplementary Table 1). The 33 cohorts were meta-analyzed with the Ricopili pipeline which uses an inverse-variance weighted fixed-effect model. We filtered our GWAS results with minor allele frequency (MAF) > 0.01 and INFO score > 0.70 (indicating “high-quality”).

Analysis of chrX.—Several cohorts in the primary GWAS did not have X chromosome variant data, specifically, some GCAN-based cohorts (*fre1*, *ukd1*, *usa1*, *gns2*) and were excluded. Imputation was performed separately from the autosome³⁵. ChrX variants in the pseudoautosomal regions were excluded prior to imputation. SNPs exceeding MAF and INFO score thresholds of 0.01 and 0.70 were retained and analysis was performed with PLINK v1.9 (<https://www.cog-genomics.org/plink2>) and Ricopili.

Female-only GWAS.—A supplementary GWAS analysis was conducted on females only to determine the similarity of the results to the primary GWAS analysis which included both females and males. The cohorts that did not have chrX variants to verify sex could not be included (*fre1*, *ukd1*, *usa1*, *gns2*).

Distance- and LD-based clumping.—The GWAS results implicate genomic regions (“loci”). To define a locus, (1) SNPs that met the genome-wide significant threshold of $P < 5 \times 10^{-8}$ were identified; (2) clumping was used to convert significant SNPs to regions. The SNP with the smallest P value in a genomic window was kept as the index SNP and SNPs in high linkage disequilibrium (LD) with the index SNP defined the left and right end of the region (SNPs with $P < 0.0001$ and $r^2 > 0.1$ within 3 Mb windows); (3) partially or wholly overlapping clumps within 50 Kb were identified and merged into one region; (4) only loci with additional evidence of association from variants in high LD as depicted by regional plots were retained; further, forest plots needed to confirm the associations based on the majority of cohorts; and (5) conditional analyses were conducted to identify SNPs with associations independent of the top SNP within the genomic chunk of interest.

Annotation.—Genome-wide significant loci were annotated with RegionAnnotator (<https://github.com/ivankosmos/RegionAnnotator>) to identify known protein-coding genes within loci (Supplementary Table 6).

Conditional and joint analysis.—Conditional and joint analysis was conducted using GCTA-COJO¹⁵. GCTA-COJO investigates every locus with a joint combination of independent markers via a genome-wide SNP selection procedure. It takes into account the LD correlations between SNPs and runs a conditional and joint analysis on the basis of conditional P values. After a model optimizing process, the joint effects of all selected SNPs are calculated. The largest subsample from our GWAS (*sedk*) was used to approximate the underlying LD structure of the investigated lead SNPs. The conditional regression was performed in a stepwise manner using the GCTA software³⁶. We analyzed SNPs that had a $P < 5 \times 10^{-8}$ (Supplementary Table 5).

Multi-trait-based conditional and joint analysis.—To separate marginal effects from conditional effects (i.e., the effect of a risk factor on an outcome controlling for the effect of another risk factor), we performed a multi-trait-based conditional and joint analysis (GCTA-mtCOJO)¹⁸ using an extension of the GCTA software³⁶ (Supplementary Table 8). This method uses summary-level data to perform the conditional analysis. We conditioned the results of our anorexia nervosa GWAS on GWAS results for education years³⁷, type 2 diabetes³⁸, HDL cholesterol³⁹, BMI (Hübel, Gaspar, Coleman, Hanscombe, Purves...Breen, unpublished report), schizophrenia⁴⁰, and neuroticism⁴¹. We again used the individual-level genotype data from our largest cohort (*sedk*) to approximate the underlying LD structure. As a first step, the method performs a generalized summary data-based Mendelian randomization (GSMR) analysis to test for causal association between the outcome (i.e., anorexia nervosa) and the risk factor (e.g., schizophrenia). We removed potentially pleiotropic SNPs from this analysis by the heterogeneity in dependent instruments (HEIDI) outlier method¹⁸. Pleiotropy is the phenomenon when a single locus directly affects several

phenotypes. The power of the HEIDI-outlier method is dependent on sample size of the GWAS. Pleiotropic SNPs are defined as the SNPs that show an effect on the outcome that significantly diverges from that expected under a causal model. Second, the GCTA-mtCOJO method calculates the genetic correlation between the exposure and the outcome using linkage disequilibrium score regression (LDSC) to adjust for genetic overlap^{19,20}. It also uses the intercept of the bivariate LDSC to account for potential sample overlap^{19,20}. As a result, GCTA-mtCOJO calculates conditional betas, conditional standard errors, and conditional *P* values. Subsequently, we clumped the conditional GWAS results using the standard PLINK v1.9⁴² algorithm (SNPs with $P < 0.0001$ and $r^2 > 0.1$ within 3 Mb windows) to investigate if any of the genome-wide significant loci showed dependency on genetic variation associated with other phenotypes. As stated in Zhu et al.¹⁸, the GCTA-mtCOJO analysis requires the estimates of b_{xy} of the covariate risk factors on the target risk factor and disease, r_g of the covariate risk factors, heritability (h^2_{snp}) for the covariate risk factors, and the sampling covariance between SNP effects estimated from potentially overlapping samples.

eQTL and Hi-C interactions.—Although GWAS findings are informative genome-wide, identifying strong hypotheses about their connections to specific genes is not straightforward. The lack of direct connections to genes constrains subsequent experimental modeling and efforts to develop improved therapeutics. Genomic location is often used to connect significant SNPs to genes, but this is problematic because GWAS loci usually contain many correlated and highly significant SNP associations over hundreds of Kb. Moreover, the three-dimensional (3D) arrangement of chromosomes in cell nuclei enables regulatory interactions between genomic regions located far apart⁴³. Chromosome conformation capture methods like Hi-C enable identification of 3D interactions *in vivo*^{44,45} and can clarify GWAS findings. For example, an intergenic region associated with multiple cancers was shown to be an enhancer for *MYC* via a long-range chromatin loop^{46,47}, and intronic *FTO* variants are robustly associated with body mass but influence expression of distal genes via long-range interactions⁴⁸. The *Nature* paper of Won et al⁴⁹ used Hi-C to assess the 3D chromatin interactome in fetal brain, and asserted connections of some schizophrenia associations to specific genes.

To gain further understanding of 3D chromatin organization of the brain and to evaluate disease relevance, we applied “easy Hi-C”⁵⁰ to postmortem samples ($N = 3$ adult temporal cortex). Library quality and yield from eHi-C are comparable to conventional Hi-C but requires much less starting material. Please refer to the following pre-print for details on methodology, data processing, quality control and statistical models used for these analyses⁵¹. We generated sufficient reads to enable a kilobase resolution map of the chromatin interactome from adult human brain. To our knowledge, these are the deepest Hi-C data on any human tissue (excluding cell lines) as they generated 22.5X as many *cis*-contacts as for the next largest datasets (DLPFC and hippocampus). We generated tissue RNA-seq, total-stranded RNA-seq, ChIP-seq (H3K27ac, H3K4me3, and CTCF), and open chromatin data (ATAC-seq) for adult brain to help interpret the eHi-C results. We also integrated brain expression and eQTL data from GTEx to aid these analyses. The Hi-C analysis is unbiased in that all chromatin interactions that pass a confidence threshold are

considered when evaluating the associations between SNPs and genes (i.e., it is not a capture experiment where only “candidate” SNP-to-gene associations are evaluated).

Similar to the work by Won *et al.*⁴⁹, we used Hi-C data generated from human adult brain to identify genes implicated by three-dimensional functional interactomics (Supplementary Figs. 5 a-h). These Hi-C data ($N=3$, anterior temporal cortex) contain more than 103K high-confidence, regulatory chromatin interactions⁵¹. These interactions capture the physical proximity of two regions of the genome in brain nuclei (“anchors”, 10 Kb resolution) although they are separated by 20 Kb to 2 Mb in genomic distance. We focused on the regulatory subset of E-P or P-P (E = enhancer, P = promoter) chromatin interactions (with P defined by location of an open chromatin anchor near the transcription start site of an adult brain-expressed transcript and E defined by overlap with open chromatin in adult brain plus either H3K27ac or H3K4me3 histone marks). The presence of a regulatory chromatin interaction from a GWAS locus to a gene provides a strong hypothesis about SNP-to-gene regulatory functional interactions.

SNP-based heritability estimation.—LDSC software (<https://github.com/bulik/ldsc>) and method were used to estimate SNP-based heritabilities for each cohort and overall^{19,20}. We used precomputed LD scores based on the 1000 Genomes Project European ancestry samples³⁴ directly downloaded from <https://github.com/bulik/ldsc>. The liability scale estimate assumed a population prevalence of 0.9%-4% for anorexia nervosa^{2,3}.

Within-trait prediction: polygenic risk scoring. Polygenic leave-one-dataset-out analysis, using PRSice v2.1.3⁵², was conducted in the first instance to identify any extreme outlying datasets. In addition, it enabled the evaluation of the association between anorexia nervosa polygenic risk score (PRS) and anorexia nervosa risk in an independent cohort as a means of replication of the GWAS results. We derived a PRS for anorexia nervosa from the meta-analysis of all datasets except for the target cohort, then applied the PRS to the target cohort to predict affected status (Supplementary Fig. 16). Logistic regression was performed, including as covariates the first five ancestry components and any other PCs significantly associated with the phenotype in the target cohort, and the target cohort was split into deciles based on anorexia nervosa PRS, with decile 1 comprised of those with the lowest anorexia nervosa PRS serving as the referent.

Anorexia nervosa subtype analysis.—PRS analyses were conducted with anorexia nervosa subgroups to investigate prediction of case status across the subtypes. For this, we split the anorexia nervosa cases to two groups based on whether binge eating was present. First, GWAS meta-analyses were conducted for (a) anorexia nervosa with binge eating vs controls (2,381 cases and 10,249 controls; $k=3$ datasets: *aunz*, *chop*, *usa2*) and (b) anorexia nervosa with no binge eating vs controls (2,262 cases and 10,254 controls; $k=3$ datasets: *aunz*, *chop*, *usa2*). Controls were randomly split between analyses to maintain independence (Supplementary Fig. 6). Genetic correlation analysis using LDSC^{19,20} was conducted to examine the potential genetic overlap of the two anorexia nervosa subtypes (Supplementary Table 9). Second, using PRSice⁵², we calculated PRS for each anorexia nervosa subtype separately in the three target cohorts for which anorexia nervosa subtype data were available.

Finally, mean PRS scores were estimated for each subtype by cohort after accounting for covariates in R. Subtype phenotyping is described in the Supplementary Methods.

Males.—In order to assess whether sex-specific differences in anorexia nervosa genetic risk load exist, we calculated PRS, using PRSice⁵², from a GWAS meta-analysis performed on females only (14,898 cases and 27,545 controls) and applied it to a male-only target cohort (447 cases and 20,347 controls) to predict affected status.

Cross-trait analysis: genetic correlations.—Common variant-based genetic correlation (SNP- r_g) measures the extent to which two traits or disorders share common genetic variation. SNP- r_g between anorexia nervosa and 447 traits (422 from an internally curated dataset and 25 from LDHub⁵³) were tested using GWAS summary statistics via an analytical extension of LDSC^{19,20}. The sources of the summary statistics files (PMID, DOI, or unpublished results) used in the LDSC are provided in Supplementary Table 10. When there were multiple summary statistics files available for a trait, significant SNP- r_g reported in the main text were chosen based on the largest sample size and/or matching ancestry with our sample (i.e., European ancestry).

Genetic correlations with anorexia nervosa corrected for BMI were carried out to investigate whether the observed genetic correlations between anorexia nervosa and metabolic phenotypes were attributable to BMI or partially independent. We used GCTA-mtCOJO¹⁸ to perform a GWAS analysis for anorexia nervosa conditioning on BMI using BMI summary data from our UK Biobank analysis (described in the next section) to derive anorexia nervosa GWAS summary statistics corrected for the common variants genetic component of BMI (Supplementary Tables 14 and 15).

GWAS of related traits in UK Biobank.—Several GWAS analyses were carried out for traits in UK Biobank to allow us to investigate body composition genetics in healthy individuals without a psychiatric disorder, a weight-altering disorder, or who were taking weight-altering medication. We also used UK Biobank to carry out GWAS of physical activity level, anxiety, and neuroticism. For details see the Supplementary Methods.

Generalized summary data-based Mendelian randomization (GSMR).—We performed two bidirectional GSMR analyses¹⁸ to test for the causal association between first, BMI and anorexia nervosa, and second, Type 2 diabetes and anorexia nervosa, using an extension of the GCTA software³⁶ (Supplementary Table 16). We used the individual-level genotype data from our largest cohort (*sedk*) to approximate the underlying LD structure. We removed potentially pleiotropic SNPs from this analysis by the HEIDI outlier method¹⁸. Pleiotropic SNPs are defined as the SNPs which show an effect on the outcome that significantly diverges from the one expected under a causal model. The method uses the intercept of the bivariate LD score regression to account for potential sample overlap^{19,20}. As a rule of thumb GSMR requires GWAS to have at least ten genome-wide significant hits. We lowered the threshold for this requirement to eight SNPs in our analyses of anorexia nervosa as an exposure and BMI or Type 2 diabetes as an outcome. Results, therefore, should be interpreted cautiously. We, furthermore, investigated bidirectional conditional effects between BMI or Type 2 diabetes and anorexia nervosa. We used GCTA-mtCOJO to

perform a GWAS analysis for anorexia nervosa conditioning on (1) BMI using summary data from our UK Biobank analysis and (2) Type 2 diabetes using summary data³⁸. Our anorexia nervosa GWAS and the BMI and Type 2 diabetes GWASs are based on independent samples. For BMI, we also re-ran the GSMR analysis using the BMI-adjusted anorexia nervosa GWAS summary data from the GCTA-mtCOJO analysis.

Gene-wise analysis.—MAGMA v1.06⁵⁴ was used to perform a gene-wise test of association with anorexia nervosa based on GWAS summary statistics. MAGMA generates gene-based P values by combining SNP-based P values within a gene while accounting for LD. In order to include regulatory regions, SNPs are mapped to genes within a 35 kb upstream and 10 kb downstream window, and the gene P value is obtained using the “multi=snp-wise” model, which aggregates mean and top SNP association models. We tested 19,846 ENSEMBL genes, including the X chromosome (Supplementary Table 11). As reference panel for the underlying LD structure we used 1000 Genomes European data phase 3³⁴.

Pathway analysis.—MAGMA v1.06⁵⁴ was used to perform a competitive pathway analysis, testing whether genes associated with anorexia nervosa were more enriched in a given pathway than all other pathways. The analysis included chrX. Biological pathways were defined using gene ontology pathways and canonical pathways from MSigDB v6.1⁵⁵, and psychiatric pathways mined from the literature. A total 7,268 pathways were tested (Supplementary Table 12).

Partitioned heritability.—Partitioned heritability was investigated using stratified LDSC²⁶ which estimates the per-SNP contribution to overall SNP-heritability (SNP- h^2) across various functional annotation categories of the genome (Supplementary Fig. 7). It accounts for linked markers and uses a ‘full baseline model’ of 24 annotations that are not specific to any cell type. We excluded the MHC region in our analysis. SNP- h^2 can be partitioned in two different ways: a non-cell type-specific and a cell type-specific manner. Partitioned heritability analysis was used to test for cell type-specific enrichment in the GWAS of anorexia nervosa among 10 cell type groups; adrenal and pancreas, cardiovascular, central nervous system (CNS), connective and bone, gastrointestinal, immune and hematopoietic, kidney, liver, skeletal muscle, and other tissue, which includes adipose tissue (Supplementary Fig. 8).

Gene expression.—We conducted a series of gene expression analyses as detailed in the Supplementary Methods.

Reporting summary

Further information on research design is available in the Life Science Reporting Summary linked to this article.

Data availability

The Psychiatric Genomics Consortium’s (PGC) policy is to make genome-wide summary results public. Genome-wide summary statistics for the meta-analysis are freely

downloadable from PGCs download website (<http://www.med.unc.edu/pgc/results-and-downloads>). Individual-level data are deposited in dbGaP (<http://www.ncbi.nlm.nih.gov/gap>) for ANGI-ANZ/SE/US (accession number phs001541.v1.p1) and CHOP/PFCG (accession number phs000679.v1.p1). ANGI-DK individual-level data are not available in dbGaP owing to Danish laws, but are available via collaboration with PIs. GCAN/WTCCC3 individual-level data are deposited in EGA (<https://www.ebi.ac.uk/ega>) (accession number EGAS00001000913) with the exception of Netherlands and US/Canada, which are available via collaboration with PIs. UK Biobank individual-level data can be applied for on the UK Biobank website (<http://www.ukbiobank.ac.uk/register-apply>).

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Authors

Hunna J Watson^{1,2,3}, Zeynep Yilmaz^{1,4,a}, Laura M Thornton^{1,a}, Christopher Hübel^{5,6,a}, Jonathan RI Coleman^{5,7,a}, Héléna A Gaspar^{5,7}, Julien Bryois⁶, Anke Hinney⁸, Virpi M Leppä⁶, Manuel Mattheisen^{9,10,11,12}, Sarah E Medland¹³, Stephan Ripke^{14,15,16}, Shuyang Yao⁶, Paola Giusti-Rodríguez⁴, Anorexia Nervosa Genetics Initiative^b, Ken B Hanscombe¹⁷, Kirstin L Purves⁵, Eating Disorders Working Group of the Psychiatric Genomics Consortium^b, Roger AH Adan^{18,19,20}, Lars Alfredsson²¹, Tetsuya Ando²², Ole A Andreassen²³, Jessica H Baker¹, Wade H Berrettini²⁴, Ilka Boehm²⁵, Claudette Boni²⁶, Vesna Boraska Perica^{27,28}, Katharina Buehren²⁹, Roland Burghardt³⁰, Matteo Cassina³¹, Sven Cichon³², Maurizio Clementi³¹, Roger D Cone³³, Philippe Courtet³⁴, Scott Crow³⁵, James J Crowley^{4,10}, Unna N Danner¹⁹, Oliver SP Davis^{36,37}, Martina de Zwaan³⁸, George Dedoussis³⁹, Daniela Degortes⁴⁰, Janiece E DeSocio⁴¹, Danielle M Dick⁴², Dimitris Dikeos⁴³, Christian Dina⁴⁴, Monika Dmitrzak-Weglaz⁴⁵, Elisa Docampo^{46,47,48}, Laramie E Duncan⁴⁹, Karin Egberts⁵⁰, Stefan Ehrlich²⁵, Geòrgia Escaramís^{46,47,48}, Tõnu Esko^{51,52}, Xavier Estivill^{46,47,48,53}, Anne Farmer⁵, Angela Favaro⁴⁰, Fernando Fernández-Aranda^{54,55}, Manfred M Fichter^{56,57}, Krista Fischer⁵¹, Manuel Föcker⁸, Lenka Foretova⁵⁸, Andreas J Forstner^{32,59,60,61,62}, Monica Forzan³¹, Christopher S Franklin²⁷, Steven Gallinger⁶³, Ina Giegling⁶⁴, Johanna Giuranna⁸, Fragiskos Gonidakis⁶⁵, Philip Gorwood^{26,66}, Monica Gratacos Mayora^{46,47,48}, Sébastien Guillaume³⁴, Yiran Guo⁶⁷, Hakon Hakonarson^{67,68}, Konstantinos Hatzikotoulas^{27,69}, Joanna Hauser⁷⁰, Johannes Hebebrand⁸, Sietske G Helder^{5,71}, Stefan Herms^{32,60,62}, Beate Herpertz-Dahlmann²⁹, Wolfgang Herzog⁷², Laura M Huckins^{27,73}, James I Hudson⁷⁴, Hartmut Imgart⁷⁵, Hidetoshi Inoko⁷⁶, Vladimir Janout⁷⁷, Susana Jiménez-Murcia^{54,55}, Antonio Julià⁷⁸, Gursharan Kalsi⁵, Deborah Kaminská⁷⁹, Jaakko Kaprio^{80,81}, Leila Karhunen⁸², Andreas Karwautz⁸³, Martien JH Kas^{18,84}, James L Kennedy^{85,86,87}, Anna Keski-Rahkonen⁸⁰, Kirsty Kiezebrink⁸⁸, Youl-Ri Kim⁸⁹, Lars Klareskog⁹⁰, Kelly L Klump⁹¹, Gun Peggy S Knudsen⁹², Maria C La Via¹, Stephanie Le Hellard^{93,94,95}, Robert D Levitan^{85,86,87}, Dong Li⁶⁷, Lisa Lilienfeld⁹⁶, Bochao Danae Lin¹⁸, Jolanta Lissowska⁹⁷, Jurjen Luykx¹⁸, Pierre J Magistretti^{98,99}, Mario Maj¹⁰⁰, Katrin Mannik^{51,101}, Sara Marsal⁷⁸,

Christian R Marshall¹⁰², Morten Mattingsdal²³, Sara McDevitt^{103,104}, Peter McGuffin⁵, Andres Metspalu^{51,105}, Ingrid Meulenbelt¹⁰⁶, Nadia Micali^{107,108,109}, Karen Mitchell¹¹⁰, Alessio Maria Monteleone¹⁰⁰, Palmiero Monteleone¹¹¹, Melissa A Munn-Chernoff¹, Benedetta Nacmias¹¹², Marie Navratilova⁵⁸, Ioanna Ntalla³⁹, Julie K O'Toole¹¹³, Roel A Ophoff^{18,114}, Leonid Padyukov⁹⁰, Aarno Palotie^{52,81,115}, Jacques Pantel²⁶, Hana Papezova⁷⁹, Dalila Pinto⁷³, Raquel Rabionet^{116,117,118}, Anu Raevuori⁸⁰, Nicolas Ramoz²⁶, Ted Reichborn-Kjennerud^{92,119}, Valdo Ricca^{112,120}, Samuli Ripatti^{52,80,121}, Franziska Ritschel^{25,122}, Marion Roberts^{5,123,124}, Alessandro Rotondo¹²⁵, Dan Rujescu^{56,64}, Filip Rybakowski¹²⁶, Paolo Santonastaso¹²⁷, André Scherag¹²⁸, Stephen W Scherer¹²⁹, Ulrike Schmidt^{7,130}, Nicholas J Schork¹³¹, Alexandra Schosser¹³², Jochen Seitz²⁹, Lenka Slachtova¹³³, P. Eline Slagboom¹⁰⁶, Margarita CT Slof-Op 't Landt^{134,135}, Agnieszka Slopian¹³⁶, Sandro Sorbi^{112,137}, Beata Wi tkowska¹³⁸, Jin P Szatkiewicz⁴, Ioanna Tachmazidou²⁷, Elena Tenconi⁴⁰, Alfonso Tortorella^{139,140}, Federica Tozzi¹⁴¹, Janet Treasure^{7,130}, Artemis Tsitsika¹⁴², Marta Tyszkiewicz-Nwafor¹³⁶, Konstantinos Tziouvas¹⁴³, Annemarie A van Elburg^{19,144}, Eric F van Furth^{134,135}, Gudrun Wagner⁸³, Esther Walton²⁵, Elisabeth Widen⁸¹, Eleftheria Zeggini^{27,69}, Stephanie Zerwas¹, Stephan Zipfel¹⁴⁵, Andrew W Bergen^{146,147}, Joseph M Boden¹⁴⁸, Harry Brandt¹⁴⁹, Steven Crawford¹⁴⁹, Katherine A Halmi¹⁵⁰, L. John Horwood¹⁴⁸, Craig Johnson¹⁵¹, Allan S Kaplan^{85,86,87}, Walter H Kaye¹⁵², James Mitchell¹⁵³, Catherine M Olsen¹³, John F Pearson¹⁵⁴, Nancy L Pedersen⁶, Michael Strober^{155,156}, Thomas Werge¹⁵⁷, David C Whiteman¹³, D. Blake Woodside^{86,87,158,159}, Garret D Stuber^{1,160}, Scott Gordon¹³, Jakob Grove^{9,161,162,163}, Anjali K Henders¹⁶⁴, Anders Juréus⁶, Katherine M Kirk¹³, Janne T Larsen^{161,165,166}, Richard Parker¹³, Liselotte Petersen^{161,165,166}, Jennifer Jordan^{123,167}, Martin Kennedy¹⁶⁸, Grant W Montgomery^{13,164,169}, Tracey D Wade¹⁷⁰, Andreas Birgegård^{10,11}, Paul Lichtenstein⁶, Claes Noring^{10,11}, Mikael Landén^{6,171,a}, Nicholas G Martin^{13,a}, Preben Bo Mortensen^{161,165,166,a}, Patrick F Sullivan^{1,4,6,a}, Gerome Breen^{5,7,c}, Cynthia M Bulik^{1,6,172,c}

Affiliations

¹Department of Psychiatry, University of North Carolina at Chapel Hill, Chapel Hill, North Carolina, US ²School of Psychology, Curtin University, Perth, Australia ³School of Paediatrics and Child Health, University of Western Australia, Perth, Australia ⁴Department of Genetics, University of North Carolina at Chapel Hill, Chapel Hill, North Carolina, US ⁵Institute of Psychiatry, Psychology and Neuroscience, Social, Genetic and Developmental Psychiatry (SGDP) Centre, King's College London, London, UK ⁶Department of Medical Epidemiology and Biostatistics, Karolinska Institutet, Stockholm, Sweden ⁷National Institute for Health Research Biomedical Research Centre, King's College London and South London and Maudsley National Health Service Foundation Trust, London, UK ⁸Department of Child and Adolescent Psychiatry, University Hospital Essen, University of Duisburg-Essen, Essen, Germany ⁹Department of Biomedicine, Aarhus University, Aarhus, Denmark ¹⁰Department of Clinical Neuroscience, Karolinska Institutet, Stockholm, Sweden ¹¹Center for Psychiatry Research, Stockholm Health Care

Services, Stockholm City Council, Stockholm, Sweden ¹²Department of Psychiatry, Psychosomatics and Psychotherapy, University of Würzburg, Würzburg, Germany ¹³QIMR Berghofer Medical Research Institute, Brisbane, Australia ¹⁴Analytic and Translational Genetics Unit, Massachusetts General Hospital, Boston, Massachusetts, US ¹⁵Stanley Center for Psychiatric Research, Broad Institute of the Massachusetts Institute of Technology and Harvard University, Cambridge, Massachusetts, US ¹⁶Department of Psychiatry and Psychotherapy, Charité - Universitätsmedizin, Berlin, Germany ¹⁷Department of Medical and Molecular Genetics, King's College London, Guy's Hospital, London, UK ¹⁸Brain Center Rudolf Magnus, Department of Translational Neuroscience, University Medical Center Utrecht, Utrecht, The Netherlands ¹⁹Center for Eating Disorders Rintveld, Altrecht Mental Health Institute, Zeist, The Netherlands ²⁰Institute of Neuroscience and Physiology, Sahlgrenska Academy at the University of Gothenburg, Gothenburg, Sweden ²¹Institute of Environmental Medicine, Karolinska Institutet, Stockholm, Sweden ²²Department of Behavioral Medicine, National Institute of Mental Health, National Center of Neurology and Psychiatry, Tokyo, Japan ²³NORMENT KG Jepsen Centre, Division of Mental Health and Addiction, University of Oslo, Oslo University Hospital, Oslo, Norway ²⁴Department of Psychiatry, Center for Neurobiology and Behavior, Perelman School of Medicine at the University of Pennsylvania, Philadelphia, Pennsylvania, US ²⁵Division of Psychological and Social Medicine and Developmental Neurosciences, Faculty of Medicine, Technische Universität Dresden, Dresden, Germany ²⁶INSERM 1266, Institute of Psychiatry and Neuroscience of Paris, Paris, France ²⁷Wellcome Sanger Institute, Wellcome Genome Campus, Hinxton, Cambridge, UK ²⁸Department of Medical Biology, School of Medicine, University of Split, Split, Croatia ²⁹Department of Child and Adolescent Psychiatry, Psychosomatics and Psychotherapy, RWTH Aachen University, Aachen, Germany ³⁰Department of Child and Adolescent Psychiatry, Klinikum Frankfurt/Oder, Frankfurt, Germany ³¹Clinical Genetics Unit, Department of Woman and Child Health, University of Padova, Padova, Italy ³²Institute of Medical Genetics and Pathology, University Hospital Basel, Basel, Switzerland ³³Life Sciences Institute and Department of Molecular and Integrative Physiology, University of Michigan, Ann Arbor, Michigan, US ³⁴Department of Emergency Psychiatry and Post-Acute Care, CHRU Montpellier, University of Montpellier, Montpellier, France ³⁵Department of Psychiatry, University of Minnesota, Minneapolis, Minnesota, US ³⁶MRC Integrative Epidemiology Unit, University of Bristol, Bristol, UK ³⁷School of Social and Community Medicine, University of Bristol, Bristol, UK ³⁸Department of Psychosomatic Medicine and Psychotherapy, Hannover Medical School, Hannover, Germany ³⁹Department of Nutrition and Dietetics, Harokopio University, Athens, Greece ⁴⁰Department of Neurosciences, University of Padova, Padova, Italy ⁴¹College of Nursing, Seattle University, Seattle, Washington, US ⁴²Department of Psychology, Virginia Commonwealth University, Richmond, Virginia, US ⁴³Department of Psychiatry, Athens University Medical School, Athens University, Athens, Greece ⁴⁴L'institut du thorax, INSERM, CNRS, UNIV Nantes, CHU Nantes, Nantes, France ⁴⁵Department of Psychiatric Genetics,

Poznan University of Medical Sciences, Poznan, Poland ⁴⁶Barcelona Institute of Science and Technology, Barcelona, Spain ⁴⁷Universitat Pompeu Fabra, Barcelona, Spain ⁴⁸Centro de Investigación Biomédica en Red en Epidemiología y Salud Pública (CIBERESP), Barcelona, Spain ⁴⁹Department of Psychiatry and Behavioral Sciences, Stanford University, Stanford, California, US ⁵⁰Department of Child and Adolescent Psychiatry, Psychosomatics and Psychotherapy, University Hospital of Würzburg, Centre for Mental Health, Würzburg, Germany ⁵¹Estonian Genome Center, University of Tartu, Tartu, Estonia ⁵²Program in Medical and Population Genetics, Broad Institute of the Massachusetts Institute of Technology and Harvard University, Cambridge, Massachusetts, US ⁵³Genomics and Disease, Bioinformatics and Genomics Programme, Centre for Genomic Regulation, Barcelona, Spain ⁵⁴Department of Psychiatry, University Hospital of Bellvitge –IDIBELL and CIBERobn, Barcelona, Spain ⁵⁵Department of Clinical Sciences, School of Medicine, University of Barcelona, Barcelona, Spain ⁵⁶Department of Psychiatry and Psychotherapy, Ludwig-Maximilians-University (LMU), Munich, Germany ⁵⁷Schön Klinik Roseneck affiliated with the Medical Faculty of the University of Munich (LMU), Munich, Germany ⁵⁸Department of Cancer, Epidemiology and Genetics, Masaryk Memorial Cancer Institute, Brno, Czech Republic ⁵⁹Institute of Human Genetics, University of Bonn, School of Medicine & University Hospital Bonn, Bonn, Germany ⁶⁰Department of Genomics, Life and Brain Center, University of Bonn, Bonn, Germany ⁶¹Department of Psychiatry (UPK), University of Basel, Basel, Switzerland ⁶²Department of Biomedicine, University of Basel, Basel, Switzerland ⁶³Department of Surgery, Faculty of Medicine, University of Toronto, Toronto, Canada ⁶⁴Department of Psychiatry, Psychotherapy and Psychosomatics, Martin Luther University of Halle-Wittenberg, Halle, Germany ⁶⁵1st Psychiatric Department, National and Kapodistrian University of Athens, Medical School, Eginition Hospital, Athens, Greece ⁶⁶CMME, Hôpital Sainte-Anne (GHU Paris Psychiatrie et Neurosciences), Paris Descartes University, Paris, France ⁶⁷Center for Applied Genomics, Children’s Hospital of Philadelphia, Philadelphia, Pennsylvania, US ⁶⁸Department of Pediatrics, Perelman School of Medicine, University of Pennsylvania, Philadelphia, Pennsylvania, US ⁶⁹Institute of Translational Genomics, Helmholtz Zentrum München, Neuherberg, Germany ⁷⁰Department of Adult Psychiatry, Poznan University of Medical Sciences, Poznan, Poland ⁷¹Zorg op Orde, Leidschendam, The Netherlands ⁷²Department of General Internal Medicine and Psychosomatics, Heidelberg University Hospital, Heidelberg University, Heidelberg, Germany ⁷³Department of Psychiatry, and Genetics and Genomics Sciences, Division of Psychiatric Genomics, Icahn School of Medicine at Mount Sinai, New York, New York, US ⁷⁴Biological Psychiatry Laboratory, McLean Hospital/Harvard Medical School, Boston, Massachusetts, US ⁷⁵Eating Disorders Unit, Parklandklinik, Bad Wildungen, Germany ⁷⁶Department of Molecular Life Science, Division of Basic Medical Science and Molecular Medicine, School of Medicine, Tokai University, Isehara, Japan ⁷⁷Faculty of Health Sciences, Palacky University, Olomouc, Czech Republic ⁷⁸Rheumatology Research Group, Vall d’Hebron Research Institute, Barcelona, Spain ⁷⁹Department of Psychiatry, First

Faculty of Medicine, Charles University, Prague, Czech Republic ⁸⁰Department of Public Health, University of Helsinki, Helsinki, Finland ⁸¹Institute for Molecular Medicine Finland, Helsinki Institute of Life Science, University of Helsinki, Helsinki, Finland ⁸²Institute of Public Health and Clinical Nutrition, Department of Clinical Nutrition, University of Eastern Finland, Kuopio, Finland ⁸³Eating Disorders Unit, Department of Child and Adolescent Psychiatry, Medical University of Vienna, Vienna, Austria ⁸⁴Groningen Institute for Evolutionary Life Sciences, University of Groningen, Groningen, The Netherlands ⁸⁵Centre for Addiction and Mental Health, Toronto, Canada ⁸⁶Institute of Medical Science, University of Toronto, Toronto, Canada ⁸⁷Department of Psychiatry, University of Toronto, Toronto, Canada ⁸⁸Institute of Applied Health Sciences, University of Aberdeen, Aberdeen, UK ⁸⁹Department of Psychiatry, Seoul Paik Hospital, Inje University, Seoul, Korea ⁹⁰Rheumatology Unit, Department of Medicine, Center for Molecular Medicine, Karolinska Institutet and Karolinska University Hospital, Stockholm, Sweden ⁹¹Department of Psychology, Michigan State University, East Lansing, Michigan, US ⁹²Department of Mental Disorders, Norwegian Institute of Public Health, Oslo, Norway ⁹³Department of Clinical Science, K.G. Jebsen Centre for Psychosis Research, Norwegian Centre for Mental Disorders Research (NORMENT), University of Bergen, Bergen, Norway ⁹⁴Dr. Einar Martens Research Group for Biological Psychiatry, Center for Medical Genetics and Molecular Medicine, Haukeland University Hospital, Bergen, Norway ⁹⁵Department of Clinical Medicine, Laboratory Building, Haukeland University Hospital, Bergen, Norway ⁹⁶American School of Professional Psychology, Argosy University, Northern Virginia, Arlington, Virginia, US ⁹⁷Department of Cancer Epidemiology and Prevention, M Skłodowska-Curie Cancer Center - Oncology Center, Warsaw, Poland ⁹⁸BESE Division, King Abdullah University of Science and Technology, Thuwal, Saudi Arabia ⁹⁹Department of Psychiatry, University of Lausanne-University Hospital of Lausanne (UNIL-CHUV), Lausanne, Switzerland ¹⁰⁰Department of Psychiatry, University of Campania "Luigi Vanvitelli", Naples, Italy ¹⁰¹Center for Integrative Genomics, University of Lausanne, Lausanne, Switzerland ¹⁰²Department of Paediatric Laboratory Medicine, The Hospital for Sick Children, Toronto, Canada ¹⁰³Department of Psychiatry, University College Cork, Cork, Ireland ¹⁰⁴HSE National Clinical Programme for Eating Disorders, Cork, Ireland ¹⁰⁵Institute of Molecular and Cell Biology, University of Tartu, Tartu, Estonia ¹⁰⁶Department of Biomedical Data Science, Leiden University Medical Centre, Leiden, The Netherlands ¹⁰⁷Department of Psychiatry, Faculty of Medicine, University of Geneva, Geneva, Switzerland ¹⁰⁸Division of Child and Adolescent Psychiatry, Geneva University Hospital, Geneva, Switzerland ¹⁰⁹Great Ormond Street Institute of Child Health, University College London, London, UK ¹¹⁰National Center for PTSD, VA Boston Healthcare System, Department of Psychiatry, Boston University School of Medicine, Boston, Massachusetts, US ¹¹¹Department of Medicine, Surgery and Dentistry "Scuola Medica Salernitana", University of Salerno, Salerno, Italy ¹¹²Department of Neuroscience, Psychology, Drug Research and Child Health (NEUROFARBA), University of Florence, Florence, Italy ¹¹³Kartini Clinic, Portland, Oregon, US

¹¹⁴Center for Neurobehavioral Genetics, Semel Institute for Neuroscience and Human Behavior, University of California Los Angeles, Los Angeles, California, US

¹¹⁵Center for Human Genome Research at the Massachusetts General Hospital, Boston, Massachusetts, US

¹¹⁶Saint Joan de Déu Research Institute, Saint Joan de Déu Barcelona Children's Hospital, Barcelona, Spain

¹¹⁷Institute of Biomedicine (IBUB), University of Barcelona, Barcelona, Spain

¹¹⁸Department of Genetics, Microbiology and Statistics, University of Barcelona, Barcelona, Spain

¹¹⁹Institute of Clinical Medicine, University of Oslo, Oslo, Norway

¹²⁰Department of Health Science, University of Florence, Florence, Italy

¹²¹Institute for Molecular Medicine Finland (FIMM), HiLIFE Unit, University of Helsinki, Helsinki, Finland

¹²²Eating Disorders Research and Treatment Center, Department of Child and Adolescent Psychiatry, Faculty of Medicine, Technische Universität Dresden, Dresden, Germany

¹²³Department of Psychological Medicine, University of Otago, Christchurch, New Zealand

¹²⁴Faculty of Medicine & Health Sciences, University of Auckland, Auckland, New Zealand

¹²⁵Department of Psychiatry, Neurobiology, Pharmacology, and Biotechnologies, University of Pisa, Pisa, Italy

¹²⁶Department of Psychiatry, Poznan University of Medical Sciences, Poznan, Poland

¹²⁷Department of Neurosciences, Padua Neuroscience Center, University of Padova, Padova, Italy

¹²⁸Institute of Medical Statistics, Computer and Data Sciences, Jena University Hospital, Jena, Germany

¹²⁹Department of Genetics and Genomic Biology, The Hospital for Sick Children, Toronto, Canada

¹³⁰Institute of Psychiatry, Psychology and Neuroscience, Department of Psychological Medicine, King's College London, London, UK

¹³¹J. Craig Venter Institute (JCVI), La Jolla, California, US

¹³²Department of Psychiatry and Psychotherapy, Medical University of Vienna, Vienna, Austria

¹³³Department of Pediatrics and Center of Applied Genomics, First Faculty of Medicine, Charles University, Prague, Czech Republic

¹³⁴Center for Eating Disorders Ursula, Rivierduinen, Leiden, The Netherlands

¹³⁵Department of Psychiatry, Leiden University Medical Centre, Leiden, The Netherlands

¹³⁶Department of Child and Adolescent Psychiatry, Poznan University of Medical Sciences, Poznan, Poland

¹³⁷IRCSS Fondazione Don Carlo Gnocchi, Florence, Italy

¹³⁸Department of Environmental Epidemiology, Nofer Institute of Occupational Medicine, Lodz, Poland

¹³⁹Department of Psychiatry, University of Naples SUN, Naples, Italy

¹⁴⁰Department of Psychiatry, University of Perugia, Perugia, Italy

¹⁴¹Brain Sciences Department, Stremble Ventures, Limassol, Cyprus

¹⁴²Adolescent Health Unit, Second Department of Pediatrics, "P. & A. Kyriakou" Children's Hospital, University of Athens, Athens, Greece

¹⁴³Pediatric Intensive Care Unit, "P. & A. Kyriakou" Children's Hospital, University of Athens, Athens, Greece

¹⁴⁴Faculty of Social and Behavioral Sciences, Utrecht University, Utrecht, The Netherlands

¹⁴⁵Department of Internal Medicine VI, Psychosomatic Medicine and Psychotherapy, University Medical Hospital Tuebingen, Tuebingen, Germany

¹⁴⁶BioRealm, LLC, Walnut, California, US

¹⁴⁷Oregon Research Institute, Eugene, Oregon, US

¹⁴⁸Christchurch Health and Development Study, University of Otago, Christchurch, New Zealand

¹⁴⁹The Center for Eating Disorders at Sheppard Pratt, Baltimore, Maryland, US

¹⁵⁰Department of Psychiatry, Weill Cornell Medical

College, New York, New York, US ¹⁵¹Eating Recovery Center, Denver, Colorado, US
¹⁵²Department of Psychiatry, University of California San Diego, San Diego, California, US ¹⁵³Department of Psychiatry and Behavioral Science, University of North Dakota School of Medicine and Health Sciences, Fargo, North Dakota, US
¹⁵⁴Biostatistics and Computational Biology Unit, University of Otago, Christchurch, New Zealand ¹⁵⁵Department of Psychiatry and Biobehavioral Science, Semel Institute for Neuroscience and Human Behavior, University of California Los Angeles, Los Angeles, California, US ¹⁵⁶David Geffen School of Medicine, University of California Los Angeles, Los Angeles, California, US ¹⁵⁷Department of Clinical Medicine, University of Copenhagen, Copenhagen, Denmark ¹⁵⁸Centre for Mental Health, University Health Network, Toronto, Canada ¹⁵⁹Program for Eating Disorders, University Health Network, Toronto, Canada ¹⁶⁰Department of Cell Biology and Physiology, University of North Carolina at Chapel Hill, Chapel Hill, North Carolina, US ¹⁶¹The Lundbeck Foundation Initiative for Integrative Psychiatric Research (iPSYCH), Aarhus, Denmark ¹⁶²Centre for Integrative Sequencing, iSEQ, Aarhus University, Aarhus, Denmark ¹⁶³Bioinformatics Research Centre, Aarhus University, Aarhus, Denmark ¹⁶⁴Institute for Molecular Bioscience, University of Queensland, Brisbane, Australia ¹⁶⁵National Centre for Register-Based Research, Aarhus BSS, Aarhus University, Aarhus, Denmark ¹⁶⁶Centre for Integrated Register-based Research (CIRRAU), Aarhus University, Aarhus, Denmark ¹⁶⁷Canterbury District Health Board, Christchurch, New Zealand ¹⁶⁸Department of Pathology and Biomedical Science, University of Otago, Christchurch, New Zealand ¹⁶⁹Queensland Brain Institute, University of Queensland, Brisbane, Australia ¹⁷⁰School of Psychology, Flinders University, Adelaide, Australia ¹⁷¹Department of Psychiatry and Neurochemistry, Institute of Neuroscience and Physiology, Sahlgrenska Academy at the University of Gothenburg, Gothenburg, Sweden ¹⁷²Department of Nutrition, University of North Carolina at Chapel Hill, Chapel Hill, North Carolina, US

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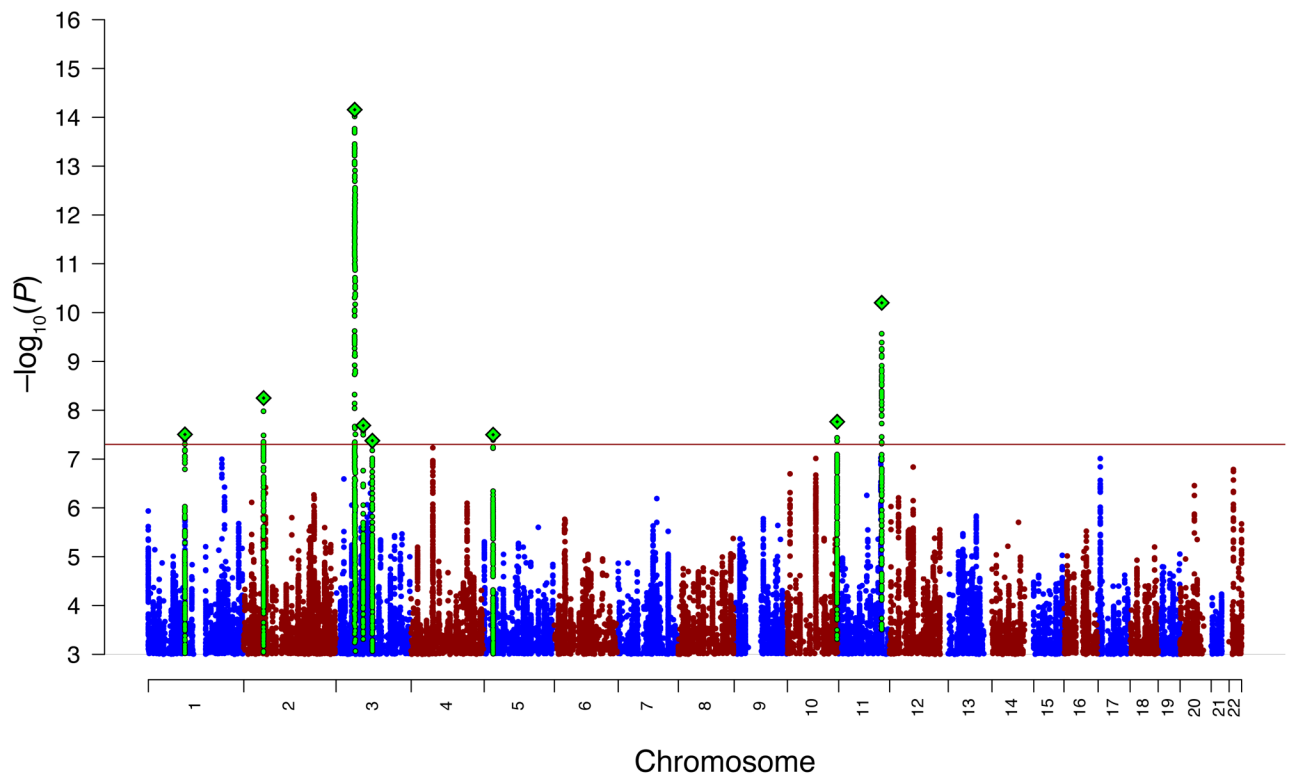


Figure 1. The Manhattan plot for the primary genome-wide association meta-analysis of anorexia nervosa with 33 case-control samples (16,992 cases and 55,525 controls of European descent).

The $-\log_{10}(P)$ values for the association tests (two-tailed) are shown on the y-axis and the chromosomes are ordered on the x-axis. Eight genetic loci surpassed genome-wide significance ($-\log_{10}(P) > 7.3$). The lead variant is indicated by a diamond and green circles show the variants in linkage-disequilibrium. The blue and red colors differentiate adjacent chromosomes.

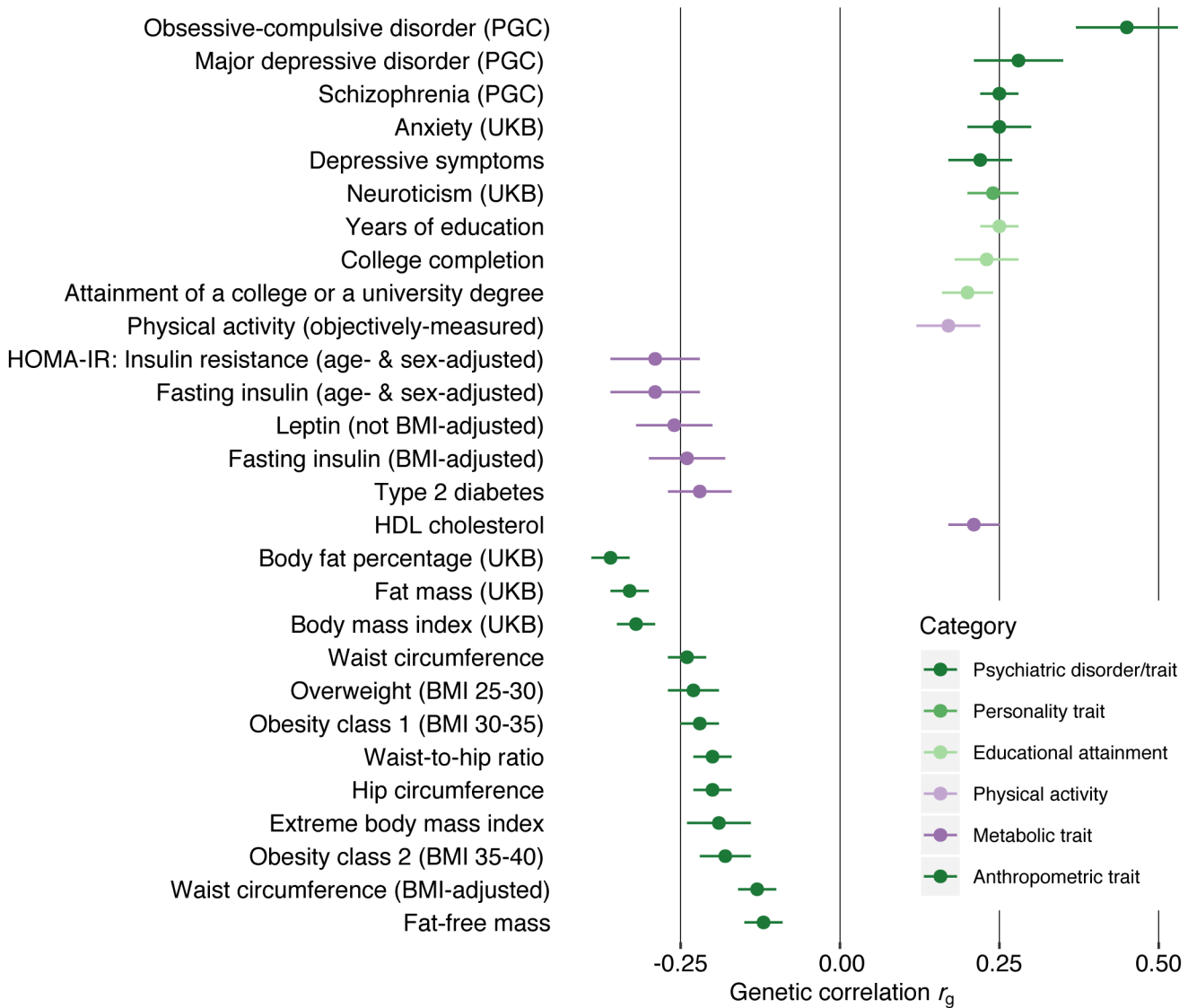


Figure 2. Bonferroni-significant genetic correlations (SNP- r_g s) and standard errors (error bars) between anorexia nervosa and other phenotypes as estimated by LD score regression. Only traits with significant P values following Bonferroni correction are shown. Correlations with 447 phenotypes were tested (Bonferroni-corrected significance threshold $P > 1.11 \times 10^{-4}$). Complete results are shown in Table S10. PGC = Psychiatric Genomics Consortium, UKB = UK Biobank, HOMA-IR = Homeostatic model assessment - insulin resistance.

Newly associated genome-wide significant loci for anorexia nervosa

Table 1.

Locus	Chr	Basepair region		Lead SNP	BP	P	A1/A2	OR	s.e.	Freq	Type	Number of genes	Nearest gene
		range left	range right										
1	3	47588253	51368253	rs9821797	48718253	6.99E-15	A/T	1.17	0.02	0.12	multigenic	111	<i>NCKIPSD</i>
2	11	114997256	115424956	rs6589488	115096956	6.31E-11	A/T	1.14	0.02	0.13	single-gene	1	<i>CADMI</i>
3	2	53881813	54362813	rs2287348	54039813	5.62E-09	T/C	1.11	0.02	0.16	multigenic	13	<i>ASB3, ERLECI</i>
4	10	131269764	131463964	rs2008387	131448764	1.73E-08	A/G	1.08	0.01	0.33	single-gene	2	<i>MGMT</i>
5	3	70670750	71074150	rs9874207	71019750	2.05E-08	C/T	1.08	0.01	0.49	single-gene	2	<i>FOXP1</i>
6	1	96699455	97284455	rs10747478	96901455	3.13E-08	T/G	1.08	0.01	0.41	single-gene	2	<i>PTBP2</i>
7	5	24945845	25372845	rs370838138	25081845	3.17E-08	G/C	1.08	0.01	0.56	intergenic	0	<i>CDH10</i>
8	3	93968107	95059107	rs13100344	94605107	4.21E-08	T/A	1.08	0.01	0.54	intergenic	2	<i>NSUN3</i>

Note. Shown are the results of the GWAS meta-analysis of anorexia nervosa (16,992 cases and 55,552 controls) which detected eight genome-wide significant loci. All of the eight loci are novel. Chr (chromosome) and Region (hg19) are shown for SNPs with $P < 1e-05$ and linkage-disequilibrium (LD) $r^2 > 0.1$ with the most associated "lead" SNP, the location of which is given in BP (basepair). A1/A2 refers to Allele 1/Allele 2 and OR and s.e. are the odds ratio and standard error for the association between A1 and the phenotype. Freq is the frequency of A1 in controls. Number of genes was determined by genomic location, adult brain eQTL, regulatory chromatin interactions, and MAGMA gene-wise analysis (see Methods). Nearest gene is the nearest gene within the region of LD "friends" of the lead variant (LD- $r^2 > 0.6$ +/- 500 Kb). The meta-analysis was restricted to variants with minor allele frequency (MAF) 0.01 and information quality (INFO) score 0.70. All loci were confirmed via forest plots based on consistent direction of effect in the majority of cohorts and via region plots whereby neighboring LD "friends" were required to show a similar effect. Chromosome X was analyzed but had no loci that reached genome-wide significance. Note that although lead variants are annotated to the nearest gene, this does not mean that the gene listed is a causal gene.