UCLA

UCLA Previously Published Works

Title

Political Orientation Predicts Credulity Regarding Putative Hazards

Permalink

https://escholarship.org/uc/item/707173p5

Authors

Fessler, Daniel M.T. Pisor, Anne C. Holbrook, Colin

Publication Date

2017-01-28

Data Availability

The data associated with this publication are available at: http://osf.io/qqq82/

Peer reviewed

Political Orientation Predicts Credulity Regarding Putative Hazards
Daniel M.T. Fessler ^{a,1} , Anne C. Pisor, b,c and Colin Holbrook ^a
^a Department of Anthropology and Center for Behavior, Evolution, & Culture
University of California, Los Angeles
Los Angeles, CA 90095-1553 USA
^b Department of Anthropology
University of California, Santa Barbara
Santa Barbara, CA 93106-3210 USA
^c Department of Human Behavior, Ecology, and Culture
Max Planck Institute for Evolutionary Anthropology
04103 Leipzig, Germany
To whom correspondence should be addressed: Department of Anthropology 341 Haines Hall University of California, Los Angeles Los Angeles, CA 90095-1553 USA Tel.: 310 794-9252 Fax: 310 206-7833 E-mail: dfessler@anthro.ucla.edu Accepted for publication in Psychological Science

The instruments described in this paper are included in the Supplementary Online Materials, and are also archived at osf.io/qqq82. The complete datasets, lists of variables, and analytic code are archived at osf.io/qqq82 and http://escholarship.org/uc/item/82j5p9r3

Abstract

To benefit from information provided by others, people must be somewhat credulous. However, credulity entails risks. The optimal level of credulity depends on the relative costs of believing misinformation versus failing to attend to accurate information. When information concerns hazards, erroneous incredulity is often more costly than erroneous credulity, as disregarding accurate warnings is more harmful than adopting unnecessary precautions. Because no equivalent asymmetry characterizes information concerning benefits, people should generally be more credulous of hazard information than of benefit information. This adaptive negatively-biased credulity is linked to negativity bias in general, and is more prominent among those who believe the world to be dangerous. Because both threat sensitivity and dangerous-world beliefs differ between conservatives and liberals, we predicted that conservatism would positively correlate with negatively-biased credulity. Two online studies of Americans support this prediction, potentially illuminating the impact of politicians' alarmist claims on different portions of the electorate.

Keywords: threat sensitivity; negativity bias; negatively-biased credulity; political orientation

In 2012, a liberal professor wrote that the Obama Administration was stockpiling ammunition, preparing for totalitarian rule. This idea was ignored by liberals. In 2015, conservative bloggers asserted that a military exercise aimed to occupy Texas and impose martial law. Conservatives became so concerned that the Texas Governor ordered the State Guard to monitor the exercise.

The different fates of these two conspiracy theories might simply reflect their historical particulars. Whereas in 2012 liberal Americans largely approved of the Obama Administration, in 2015 most conservative Americans did not. Perhaps the first theory died while the second prospered simply because the latter resonated with the views of a substantial audience while the former did not. However, two bodies of research suggest that psychological differences related to political orientation may also have been at work. First, a sizeable literature documents that, in the U.S., responsiveness to negative stimuli correlates with political orientation, with conservatives displaying more responsiveness, and liberals displaying less. Second, recent studies indicate that people are more credulous of information concerning hazards than of information concerning benefits – and individuals differ in this regard. Here, we combine these approaches, testing the hypothesis that political orientation is correlated with differences in credulity toward hazard information. If correct, this thesis potentially illuminates the differential impacts that politicians' alarmist claims have on liberal and conservative constituencies.

We employ the terms "liberal" and "conservative" recognizing that these are heterogeneous categories, and that self-identifying members of each may hold internally incompatible positions on various issues; we view these features as a source of noise, hence any differences found despite them constitute foundational orientations shared by core groups of category members (Weeden & Kurzban, 2016). Research has revealed psychological differences

between liberals and conservatives, including both broad features of personality (Carney, Jost, Gosling, & Potter, 2008) and the priority given to different moral principles (Graham, Haidt, & Nosek, 2009). Reviewing a large number of studies, Hibbing, Smith, and Alford (2014) concluded that conservatives display greater "negativity bias" than do liberals (or, perhaps more precisely, "threat bias" [Lilienfeld & Latzman, 2014], i.e., sensitivity to the possibility of danger). Subsequent research has largely bolstered this conclusion (Ahn et al., 2014; Mills, Smith, Hibbing, & Dodd, 2014; Mills et al., 2016; but see Knoll, O'Daniel, & Cusato, 2015). Like other animals, humans exhibit negativity bias – compared to positive events, negative events capture attention and information processing more readily, elicit strong emotions more easily, and are more memorable (Rozin & Royzman, 2001; Baumeister, Bratslavsky, Finkenauer, & Vohs, 2001). If valence indexes the biological fitness implications that a class of events would have had in ancestral environments, then this pattern is explicable in evolutionary terms as stemming from the generally greater detrimental fitness consequences of failing to immediately attend to, address, and learn from fitness-reducing events compared to failing to do so for fitness-enhancing events, as threats frequently both are more imminent than, and preclude, opportunities (Rozin & Royzman, 2001; Baumeister et al., 2001). Within a species, the optimal level of negativity bias will depend on the interaction of features of both the individual and the environment (e.g., a vulnerable individual in a hazardous environment should be guided by greater negativity bias than a robust individual in a safe environment, etc.) – there is no invariantly "correct" degree of negativity bias. Consonant with this, there are substantial individual differences in negativity bias. If a core dimension of political orientation is that liberals value the opportunities afforded by change and cultural heterogeneity, whereas conservatives value the safety of tradition and cultural homogeneity, then conservatism is more

69

70

71

72

73

74

75

76

77

78

79

80

81

82

83

84

85

86

87

88

89

90

91

consonant with pronounced negativity bias than is liberalism, as conservatives will often see pitfalls where liberals see promise (Hibbing et al., 2014).

92

93

94

95

96

97

98

99

100

101

102

103

104

105

106

107

108

109

110

111

112

113

114

While the evolutionary considerations underlying negativity bias apply across species, in humans they intersect with our reliance on cultural information. Our species uniquely exploits cumulative cultural evolution and the technological and organizational advantages that it provides – we are culture-dependent, a characteristic likely undergirded by specific psychological mechanisms for acquiring cultural information (Fessler, 2006). Relying on cultural information necessitates credulity, as the utility of a given practice is frequently not evident to the learner, and is often opaque even to teachers (Boyd & Richerson, 2006). However, those who are overly credulous risk acquiring erroneous information and/or being exploited (Kurzban, 2007). The trade-off between the benefits of credulity and its costs varies as a function of information type, such that the optimal level of credulity differs across different messages. With regard to information concerning hazards, the costs of erroneous credulity will often be lower than the costs of erroneous incredulity: while the former results in unnecessary precautions, the latter can result in injury or death. (As these possibilities indicate, the extent of the asymmetry in costs depends on the magnitude of the consequences should the information prove accurate.) Because no equivalently overarching asymmetry exists with regard to information concerning benefits, people should exhibit negatively-biased credulity, i.e., ceteris paribus, they should more readily view as true information concerning hazards relative to information concerning benefits (Fessler, Pisor, & Navarrete, 2014). Experimental results confirm this – when statements are framed as being about hazards they are judged more likely to be true than when they are framed as involving benefits (Fessler et al., 2014; see also Hilbig, 2009; Hilbig, 2012a; Hilbig, 2012b).

At the proximate level, negatively-biased credulity is explained by the greater processing fluency attending negative information, thus linking negatively-biased credulity to negativity bias in general (Hilbig, 2009; Hilbig, 2012a; Hilbig, 2012b). Given that conservatives display greater threat sensitivity, and may display greater negativity bias, than do liberals, this proximate pathway generates the prediction that conservatives will exhibit greater negatively-biased credulity than liberals. This prediction is reinforced by additional conceptual and empirical considerations.

115

116

117

118

119

120

121

122

123

124

125

126

127

128

129

130

131

132

133

134

135

136

Because newly-identified hazards often share features, and therefore co-occur, with previously-known hazards, the more dangerous the world in which one lives, the more likely that one will encounter additional hazards, and thus the greater the asymmetry between the costs of erroneous credulity and those of erroneous incredulity when assessing information concerning hazards. Accordingly, individuals who know (or believe they know) of the existence of many hazards should display elevated negatively-biased credulity. This functionality is reinforced at the proximate level, as congruence between a message and prior beliefs enhances biased credulity (White, Pahl, Buehner, & Haye, 2003). Consonant with the above, belief that the world is dangerous correlates positively with negatively-biased credulity (Fessler et al., 2014). Importantly, in keeping with conservatives' view of tradition and cultural homogeneity as buffers against an uncertain world, conservatism is linked with dangerous-world beliefs, both directly and via associations with authoritarianism (Federico, Hunt, & Ergun, 2009; relatedly, see Altemeyer, 1998; Crowson, Thoma, & Hestevold, 2005; Duckitt, 2001; Duckitt, Wagner, Du Plessis, & Birum, 2002; van Leeuwen & Park, 2009; Lilienfeld & Latzman, 2014). Hence, if conservatives view the world as more dangerous than do liberals, then conservatives should

display more negatively-biased credulity than liberals. To test this prediction, we measured negatively-biased credulity and assessed political orientation in two U.S. samples.

Study 1 Methods

Participants

On the basis of variance observed in Fessler et al. (2014) Study 2, an approximate final sample size of 450 was targeted. Expecting attrition and exclusions, in early October of 2015, 540 U.S. participants were recruited via MechanicalTurk.com in exchange for \$0.50. Data were pre-screened for minimal completeness (see below), repeat participation, taking at least 3 minutes to complete the study, speaking English as a first language, and answering "catch questions" (descriptive statistics in Table S2a; predictors of exclusion reported in Table S3). The final sample consisted of 472 adults (48% female; 81% White) ranging in age from 19 to 65 (M = 36.03, SD = 11.81).

Materials and Procedure

We created a credulity scale consisting of fourteen plausible but false statements, and two true statements included to preclude deception (participants were informed that some of the statements were factual). For each of eight domains, one statement concerned a benefit and one concerned a hazard (e.g., "Eating carrots results in significantly improved vision," "Kale contains thallium, a toxic heavy metal, that the plant absorbs from soil"; see SOM for complete instrument). Participants reported judgments of truthfulness using 1-7 scales (1 = I'm absolutely certain this statement is FALSE; 7 = I'm absolutely certain this statement is TRUE). As noted earlier, the magnitude of the phenomenon addressed by a message should color credulity toward

it, as any asymmetry between the costs of erroneous credulity and erroneous incredulity will be a function of the significance of the benefit or hazard at issue. Statements were therefore selected so that, for a given domain, the presumed magnitudes of the benefit or hazard were approximately equal; additionally, participants were asked to judge these magnitudes using a 1-7 scale (1 = The benefit [hazard] described in this statement is SMALL; 7 = The benefit [hazard] described in this statement is LARGE). In cases of incomplete responses, if a participant left fewer than 10% of the items unanswered, missing responses were imputed (see SOM, Appendix 3; see Table S5 for model fits without imputation). Statements were presented in truly random order. To measure bias in credulity regarding hazard information relative to benefit information, in the models reported in the main text we examine the difference between hazard credulity and benefit credulity; the SOM presents complementary models respectively examining only hazard credulity or only benefit credulity as the response (Tables S6a-b).

Next, political orientation was assessed using four measures. First, participants completed a slightly updated form of Dodd et al.'s (2012) version of a Wilson and Patterson (1968) issues index (see SOM) in which participants indicate whether they agree, disagree, or are uncertain regarding 28 contemporary issues, half of which are favored by conservatives (e.g., "Biblical truth," "tax cuts"), and half of which are favored by liberals (e.g., "abortion rights," "socialism"). For each conservative topic, agreement was scored as +1 and disagreement as -1, with reverse scoring for liberal topics; "uncertain" was scored as 0. With three exceptions (see SOM Appendix 1), responses to all topics were summed such that increasingly positive values indicate greater conservatism (α = .88). Second, using Dodd et al.'s social principles index (minus one item concerning danger – see SOM), participants selected one of two completions of the stem "Society works best when..." (e.g., "people are rewarded according to merit" versus

"people are rewarded according to need"). The choices are intended to capture preferences for traditional social order, in-group favoritism, obedience to authority, and punishment of transgressions. Typically conservative responses were coded as "1," typically liberal responses were coded as "-1", then responses were summed such that larger values indicate greater conservatism ($\alpha = .72$). All items and stem-completion options were presented in truly random order. Any missing values were imputed if participants failed to answer less than 10% of these measures (see Table S5 for fit without imputation). Third, participants indicated their political position on a 9-point scale ("strongly liberal" = 1, "strongly conservative" = 9). Lastly, participants reported their political party affiliation, scored as +1 for traditionally conservative parties ("Republican", "Tea Party"), -1 for traditionally liberal parties ("Democrat", "Green"), and 0 for Libertarians or unaffiliated individuals. Demographic items followed, including parenthood status, as previous research (see Fessler, Holbrook, Pollack, & Hahn-Holbrook, 2014) suggests that parents may be more sensitive to the presence of hazards than non-parents. Additionally, participant height and self-assessed fighting ability were collected for a future study; exploratory analyses indicate these have no bearing on the results of interest here, hence they are not reported. See SOM for complete survey.

199

200

201

202

203

204

183

184

185

186

187

188

189

190

191

192

193

194

195

196

197

198

Study 1 Results

To facilitate participant comprehension, in our credulity measure, for each item the low end of the Likert-type scale is anchored by 1 ("I'm absolutely certain this statement is FALSE"). Our weighting procedure involves multiplying the participant's response on this scale by the participant's assessment of the magnitude of the given hazard or benefit. Accordingly, to

preclude assigning a positive multiplicative product to items deemed entirely false by a participant, we began by subtracting 1 from all credulity responses.

Because our four measures of political orientation had disparate ranges, we z-scored each measure, performed a principal components analysis, and extracted the first component (summarizing 72.65% of the variance, each measure having a loading of 0.80 or higher) as a summary of political orientation, where higher values indicate greater conservatism. (An alternative variable created by summing the four measures together produced similar results when included in our models – see SOM Table S7.)

Employing the R statistical program version 3.3.1 (R Core Team, 2016), linear models were fit with the difference between hazard credulity (weighted by the participant's perceived magnitude for each respective item) and benefit credulity (similarly weighted) as the response. Variables that exhibited skewness were rounded down to the 97.5th percentile if negatively skewed, and up to the 2.5th percentile if positively skewed (see SOM Appendix 2). No models exhibited collinearity, i.e., none exhibited a variance inflation factor greater than 3.

Although not significant, participants tended to find our (almost entirely false) weighted credulity-scale items more believable if they concerned a hazard rather than a benefit ($M_{hazard} = 12.28$, $M_{benefit} = 11.96$, t(934.51) = 1.02, p = .31). A participant's average credulity toward benefits was correlated with the participant's average credulity toward hazards, r = .41.

Addressing the key prediction at issue, participants who were more conservative were significantly more likely to exhibit greater credulity for information about hazards relative to information about benefits (Table 1), an effect independent of controls (Table S4). Treating hazard credulity separately from benefit credulity confirms these results: conservatism has a positive effect on hazard credulity, but no effect on benefit credulity (Tables S6a-b); this is true

even if we do not weight credulity by the participant's perceived magnitude of the hazard or benefit described in each item (Table S8) or if we treat credulity for each item as a separate response (and include a random intercept for each participant and item; Tables S9a-b). Likewise, this effect is robust to the exclusion of any single item (see Figure S1). The relationship between conservatism and negatively-biased credulity was driven predominantly by participants' responses to the Wilson-Patterson issues index (Table 2a). More specifically, items from this index addressing social conservatism predicted negatively-biased credulity; the effect of conservative views on the military, obedience to authority, and punishment was in the same direction, albeit not significant, while there was no effect of fiscal conservatism (Table 3; Figure 1a; see SOM Appendix 1 for the Wilson-Patterson issues index items by category).

(TABLES 1-3 APPEAR ON THE FOLLOWING PAGES)

Table 1. Unstandardized Parameter Estimates, Standardized Parameter Estimates, 95% Confidence Intervals for Unstandardized Parameter Estimates, and P Values for Models with Political Summary Measure as a Predictor of the Difference between Weighted Hazard Credulity and Weighted Benefit Credulity.

			Study 1	Study 2						
Variable	Parm. Est.	Std. Est.	Lower 95%	Upper 95%	p	Parm. Est.	Std. Est.	Lower 95%	Upper 95%	p
Intercept	27	.00	-2.09	1.54	.77	.63	.00	-1.25	2.52	.51
Polit. summ.	.36	.12	.08	.65	.01	.54	.19	.28	.81	.00

Study 1: N = 472. Adjusted $R^2 = .01$, F(10, 461) = 1.66, p = .09. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 476. Adjusted $R^2 = .03$, F(12, 463) = 2.09, p = .02. Women, "other" ethnicity, some high school/high school diploma, and median general reasoning ability held at zero. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes (see Table S10 for regression on the subset for which parenthood status was available, Study 2).

Table 2a.

Study 1: Unstandardized Parameter Estimates, Standardized Parameter Estimates, 95% Confidence Intervals for Unstandardized Parameter Estimates, and P Values for Models with Distinct Political Measures as Predictors of the Difference between Weighted Hazard Credulity and Weighted Benefit Credulity.

	Wilson-Patterson Issues			S	Society Works			Political Likert				Political Category*							
	Parm	Std	5%	95%		Parm	Std	5%	95%		Parm	Std	5%	95%		Parm	5%	95%	
Variable	Est	Est	CI	CI	p	Est	Est	CI	CI	p	Est	Est	CI	CI	p	Est	CI	CI	p
Intercept	.07	.00	-1.76	1.90	.94	.02	.00	-1.85	1.89	.98	-1.00	.00	-2.95	.95	.32	81	-2.69	1.07	.40
Issues	.09	.16	.04	.14	.00														
Society						.09	.09	.00	.18	.05									
Likert Category:											.17	.07	05	.39	.13				
Libert/Unaff																.74	34	1.83	.18
Conservat																.97	34	2.27	.15

N = 472. Wilson-Patterson Issues model: adjusted $R^2 = .02$, F(10, 461) = 2.14, p = .02. Society Works model: adjusted $R^2 = .01$, F(10, 461) = 1.43, p = .17. Political Likert model: adjusted $R^2 = .01$, F(10, 461) = 1.27, p = .25. Political Category model: adjusted $R^2 = .01$, F(11, 460) = 1.20, p = .28. *Standardized betas not provided for categorical variables.

Table 2b

Study 2: Unstandardized Parameter Estimates, Standardized Parameter Estimates, 95% Confidence Intervals for Unstandardized

Parameter Estimates, and P Values for Models with Distinct Political Measures as Predictors of the Difference between Weighted

Hazard Credulity and Weighted Benefit Credulity.

	Wils	son-P	atters	on Issi	ies		Socie	ety Wor	rks			Pol	itical Li	kert		Poli	itical C	'ategor	<i>y</i> *
	Parm	Std	5%	95%		Parm	Std	5%	95%		Parm	Std	5%	95%		Parm	5%	95%	
Variable	Est	Est	CI	CI	p	Est	Est	CI	CI	p	Est	Est	CI	CI	p	Est	CI	CI	p
Intercept	.89	.00	99	2.77	.35	.82	.00	-1.09	2.74	.40	-1.03	.00	-3.11	1.06	.34	35	-2.33	1.64	.73
Issues	.10	.22	.06	.15	.00														
Society						.12	.15	.05	.20	.00									
Likert											.33	.14	.11	.54	.00				
Category:																			
Libert/Unaff																.77	31	1.84	.16
Conservat																1.74	.60	2.88	.00

N = 476. Wilson-Patterson Issues model: adjusted $R^2 = .04$, F(12, 463) = 2.55, p = .003. Society Works model: adjusted $R^2 = .01$, F(12, 463) = 1.54, p = .11. Political Likert model: adjusted $R^2 = .01$, F(12, 463) = 1.49, p = .13. Political Category model: adjusted $R^2 = .01$, F(13, 462) = 1.37, p = .17. *Standardized betas not provided for categorical variables.

Table 3. Unstandardized Parameter Estimates, Standardized Parameter Estimates, 95% Confidence Intervals for Unstandardized Parameter Estimates, and P Values for Models with Social Conservatism, Fiscal Conservatism, and Military/Obedience/Punishment Conservatism as Predictors of the Difference between Weighted Hazard Credulity and Weighted Benefit Credulity.

			Study 1		Study 2					
Variable	Parm. Est.	Std. Est.	5% CI	95% CI	p	Parm. Est.	Std. Est.	5% CI	95% CI	p
Intercept	44	.00	-2.25	1.37	.63	.51	.00	-1.38	2.41	.60
Social	.29	.11	.00	.58	.05	.33	.14	.07	.59	.01
Fiscal	.00	.00	38	.38	.99	.13	.04	22	.48	.46
Military	.23	.07	10	.56	.18	.24	.09	05	.53	.11

Study 1: N = 472. Adjusted $R^2 = .02$, F(12, 459) = 1.78, p = .046. Sub-scales of the Wilson-Patterson issues index (modified from Dodd et al., 2012), summarized by first principal component; see SOM for details.

Study 2: N = 476. Adjusted $R^2 = .04$, F(14, 461) = 2.24, p = .006. Sub-scales of the Wilson-Patterson issues index (modified from Dodd et al., 2012), summarized by first principal component; see SOM for details.

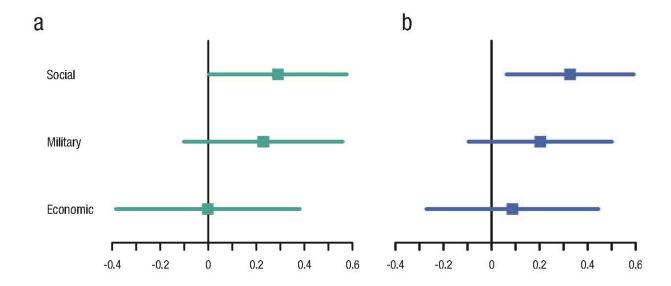


Fig. 1. Unstandardized parameter estimates with 95% confidence intervals for social, military, and fiscal conservatism for (a) Study 1 and (b) Study 2.

Discussion

Study 1 documented the predicted association between political orientation and negatively-biased credulity. However, likely reflecting shortcomings of MechanicalTurk, the sample suffered nontrivial data loss, and was not balanced as regards political orientation, being skewed left. We therefore conducted a second study, recruiting participants via Prolific Academic, an arguably superior online platform (Peer, Samat, Brandimarte, & Acquisti, 2015). Study 2 also improved on Study 1 by replacing outdated military items ("Patriot Act", "Iraq war") with contemporary topics (e.g., "Drone strikes," "Bomb cities controlled by terrorists"). To rule out the possibility that the pattern documented in Study 1 derives from differences in general reasoning abilities (Kemmelmeier, 2008), we added short measures of problem-solving and abstract reasoning (see SOM).

Study 2 Methods

Participants

In Study 2, in early September of 2016, 738 U.S. participants were recruited via Prolific Academic in exchange for \$2.31. Data were pre-screened for completeness, repeat participation, taking at least 10 minutes to complete the study (the cutoff was extended from Study 1 due to the addition of time-consuming measures of reasoning and problem-solving), speaking English as a first language, and correctly answering "catch questions" (descriptive statistics in Table S2b; predictors of exclusion reported in Table S3). As the sample evinced a left-skewed political orientation, we randomly excluded participants who self-identified as more liberal (i.e., a 2 or lower) on the 9-point political orientation scale until our sample approximated the distribution of political orientations in the U.S. as documented in a Gallup poll conducted a few months prior to

our study (Jones & Saad, 2016). Results are robust to the exclusion or inclusion of these individuals (see Table S11). The final sample consisted of 476 adults (40% female; 79% White) ranging in age from 18 to 73 (M = 34.32, SD = 12.56).

Materials and Procedures

Participants were presented with the same credulity scales described in Study 1. Statements were presented in truly random order. Political orientation was assessed using the four measures described in Study 1, with some minor modifications. As noted above, items concerning U.S. military policy in Dodd et al.'s (2012) version of a Wilson and Patterson (1968) issues index were updated (see SOM). With two exceptions (see SOM Appendix 1), responses to all topics were summed; the scale had a high degree of internal consistency ($\alpha = .82$). Dodd et al.'s social principles index (minus one item concerning danger – see SOM) again had high internal consistency ($\alpha = .74$). This was followed by demographic items and measures of general reasoning ability (see SOM for complete survey). Many participants failed to indicate whether they were parents, so parenthood status is excluded from all models unless otherwise stated.

Study 2 Results

Because our four measures of political orientation had disparate ranges, we z-scored each measure, performed a principal components analysis, and extracted the first component (summarizing 73.90% of the variance, each having a loading of .77 or higher) as a summary of political orientation, where higher values indicate greater conservatism. (An alternative variable created by summing the four measures together produced similar results when included in our models – see SOM Table S6.)

Linear models were fit with the difference between weighted hazard credulity and weighted benefit credulity as the response. Variables that exhibited skewness were rounded down to the 97.5th percentile if negatively skewed (see SOM Appendix 2). No models exhibited collinearity.

286

287

288

289

290

291

292

293

294

295

296

297

298

299

300

301

302

303

304

305

306

307

308

Participants found weighted credulity-scale items significantly more believable if they concerned a hazard rather than a benefit ($M_{hazard} = 12.82$, $M_{benefit} = 11.48$, t = 4.03, p < .001). A participant's average credulity toward benefits was correlated with the participant's average credulity toward hazards, r = .48.

Addressing the key prediction at issue, participants who were more conservative were again significantly more likely to exhibit greater credulity for information about hazards relative to information about benefits (Table 1), an effect independent of the effects of controls (Table S4); the same is true of the entire sample (i.e., when no highly liberal individuals are excluded) – see Table S11. One item (concerning terrorism) had a large influence on hazard credulity. Although exclusion of this item diminished the magnitude of the effect below significance, the effect consistently remained in the same direction across multiple iterations of the model, varying only slightly as a function of the set of liberals excluded (see Figure S1b). Treating hazard credulity separately from benefit credulity corroborates the predicted relationship: conservatism has a positive effect on hazard credulity, but no effect on benefit credulity (Tables S6a-b; see Figure S2b for the varied effect of excluding the terrorism item); this is true even if we do not weight credulity by the participant's perceived magnitude of the hazard or benefit described in each item (Table S8) or if we treat credulity for each item as a separate response (and include a random intercept for each participant and item; Tables S9a-b). As in Study 1, the relationship between conservatism and negatively-biased credulity was driven predominantly by

participants' responses to the Wilson-Patterson issues index (Table 2b). Also as in Study 1, items from this index addressing social conservatism predicted negatively-biased credulity, and, once again, the effect of conservative views on the military, obedience to authority, and punishment, was in the same direction though not significant, while fiscal conservatism again made no notable contribution in this regard (Table 3; Figure 1b; see SOM Appendix 1 for the Wilson-Patterson issues index items by category).

General Discussion

Because liberals and conservatives differ in responsiveness to negative information, particularly concerning threats, and similarly differ in how dangerous they perceive the world to be, we predicted, and found, that political orientation correlates with the tendency to believe information about hazards relative to the tendency to believe information about benefits, with liberals displaying less of this propensity and conservatives displaying more of it. This effect was driven by political orientation as defined by views on social issues. These results contribute to a corpus suggesting that, due to the intersection of variance in environments and variance in individual capabilities, a variety of potentially viable strategies emerge, with some individuals being more sensitive to the possibility of threats, and, correspondingly, paying higher precautionary costs, and others being less sensitive to this possibility, and paying higher costs when hazards are encountered.

While the predicted relationships are evident in our results, these findings should be considered preliminary given that ours were not representative nationwide samples, and our credulity measure consists of a small number of items. Indeed, its limited scope likely explains why, although in Study 2 our novel measure produced the previously documented overarching

pattern of negatively-biased credulity, in Study 1 this pattern was nonsignificant, albeit in the predicted direction. The same limitation may account for the outsized influence of one item on the key results of Study 2.

Because older individuals display less negativity bias than younger individuals (Reed, Chan, & Mikels, 2014), yet are generally more conservative (Cornelis, Van Hiel, Roets, & Kossowska, 2009), some have questioned the relationship between negativity bias and conservatism (Sedek, Kossowska, & Rydzewska, 2014). While our data do not resolve this, examining wide age ranges, we find no interaction between political orientation and age in predicting negatively-biased credulity (SOM Table S13; Figure S2). Rather, we find an effect of political orientation even when age is controlled for (Tables 1-3, Table S4), suggesting independent effects.

Social conservatism, but not fiscal conservatism, predicts increased negatively-biased credulity. Whereas fiscal conservatism is orthogonal to individuals' exposure to hazards, adherence to what are seen as tried-and-true rules for social organization and personal comportment – the foundations of social conservatism – is, for its proponents, a defense against disorder and danger; correspondingly, social conservatism correlates with threat-relevant personality features differentiating liberals and conservatives, but fiscal conservatism does not (Carney, Jost, Gosling, & Potter, 2008). Although in our models negatively-biased credulity is not predicted by conservative views on the military, obedience to authority, and endorsement of punishment (all of which concern avenues for enhancing stability and safety), consonant with the above reasoning, the magnitude of the association between this characteristic and negatively-biased credulity does not differ greatly from that of social conservatism (see Figure 1). Future work should therefore further examine the impact of this attribute on negatively-biased credulity.

The difference in negatively-biased credulity that we document likely interdigitates with related phenomena. Consonant with negatively-biased credulity, people judge those providing information about hazards as more competent than those providing other information (Boyer & Parren, 2015); our findings suggest that conservatives will display this pattern more than liberals. A parallel bias exists in information transmission, as people are more likely to transmit messages concerning hazards than messages concerning benefits (Altshteyn, 2014; Bebbington, MacLeod, Ellison, & Fay, in press; but see Stubbersfield, Tehrani, & Flynn, 2015). Political orientation likely shapes this bias also, potentially influencing the speed and breadth of dissemination of messages as a function of the political composition of a social network. A variety of phenomena thus link to negatively-biased credulity in a manner suggesting that politicians' alarmist claims will differentially impact liberals and conservatives.

In the 2016 U.S. election, President-elect Donald Trump enjoyed support from social conservatives despite being a recent convert to their positions; displaying limited familiarity with their scriptures; and having boasted of violating one of their commandments. While this support may have largely derived from, for example, Mr. Trump's opposition to abortion, the relationship between political orientation and negatively-biased credulity suggests that social conservatives may also have been influenced by his alarmist rhetoric, finding plausible such readily falsifiable claims as his August 29, 2016 tweet that "Inner-city crime is reaching record levels". Similarly, while it is difficult to gauge the effect of fake news on the election, the credence given by social conservatives to bogus reports of nefarious conspiracies apparently explains why profit-minded purveyors of fake news disproportionately targeted conservative audiences (Sydell, 2016). More broadly, although distinguishing between Chicken Little and Cassandra is frequently difficult – with grave perils attending mistakes on both sides – it seems

378 that social conservatives may be more apt to follow the former into the fox's den than they are to 379 disregard the latter and witness the fall of Troy. 380 381 Author Contributions: D.M.T. Fessler and C. Holbrook conceived of the study. D.M.T. 382 Fessler developed the methods with input from A.C. Pisor and C. Holbrook. C. Holbrook 383 oversaw data collection. A.C. Pisor conducted all analyses, with input from C. Holbrook and 384 D.M.T. Fessler, D.M.T. Fessler drafted the manuscript with critical revisions from A.C. Pisor 385 and C. Holbrook. All authors approved the final version of the manuscript for submission. 386 387 **Acknowledgments** 388 We thank Scott Lilienfeld, John Hibbing, and Eddie Harmon-Jones for helpful feedback. C. 389 Holbrook was supported by the U.S. Air Force Office of Scientific 390 Research under Award #FA9550-115-1-0469. 391 392 References 393 Ahn, W.-Y., Kishida, K. T., Gu, X., Lohrenz, T., Harvey, A., Alford, J. R., Smith, K. B., Yaffe, 394 G., Hibbing, J. R., and Dayan, P. (2014). Nonpolitical images evoke neural predictors of 395 political ideology. Current Biology, 24(22), 2693-2699. (doi: 10.1016/j.cub.2014.09.050) 396 Altemeyer, B. (1998). The other "authoritarian personality". Advances in Experimental Social 397 Psychology, 30, 47-92. 398 Altshteyn, I. (2014). Evidence for a warning bias in information transmission in social networks. 399 M.A. University of California, Los Angeles, Los Angeles.

Baumeister, R. F., Bratslavsky, E., Finkenauer, C., & Vohs, K. D. (2001). Bad is stronger than

400

101	good. Review of General Psychology, 5(4), 323-370. (doi: 10.1037/1089-2680.5.4.323)
102	Bebbington, K., MacLeod, C., Ellison, T. M., & Fay, N. (in press). The sky is falling: evidence
103	of a negativity bias in the social transmission of information. Evolution and Human
104	Behavior. (doi: 10.1016/j.evolhumbehav.2016.07.004) Retrieved from
105	http://www.ehbonline.org/article/S1090-5138(16)30166-0/fulltext
106	Boyd, R., & Richerson, P. J. (2006). Culture and the evolution of the human social instincts. In
107	S. Levinson & N. Enfield (Eds.), Roots of human sociality (pp. 453-477). Oxford: Berg.
108	Boyer, P., & Parren, N. (2015). Threat-related information suggests competence: A possible
109	factor in the spread of rumors. PloS ONE, 10(6), e0128421. (doi:
110	10.1371/journal.pone.0128421)
4 11	Carney, D. R., Jost, J. T., Gosling, S. D., & Potter, J. (2008). The secret lives of liberals and
112	conservatives: Personality profiles, interaction styles, and the things they leave behind.
113	Political Psychology, 29(6), 807-840. (doi: 10.1111/j.1467-9221.2008.00668.x)
114	Cornelis, I., Van Hiel, A., Roets, A., & Kossowska, M. (2009). Age differences in conservatism
115	Evidence on the mediating effects of personality and cognitive style. Journal of
116	Personality, 77(1), 51-88. (doi: 10.1111/j.1467-6494.2008.00538.x)
117	Crowson, H. M., Thoma, S. J., & Hestevold, N. (2005). Is political conservatism synonymous
118	with authoritarianism? The Journal of Social Psychology, 145(5), 571-592. (doi:
119	10.3200/SOCP.145.5.571-592)
120	Dodd, M. D., Balzer, A., Jacobs, C. M., Gruszczynski, M. W., Smith, K. B., & Hibbing, J. R.
121	(2012). The political left rolls with the good and the political right confronts the bad:
122	connecting physiology and cognition to preferences. Philosophical Transactions of the
123	Royal Society of London B: Biological Sciences, 367(1589), 640-649. (doi:

- 424 10.1098/rstb.2011.0268)
- Duckitt, J. (2001). A dual-process cognitive-motivational theory of ideology and prejudice.
- 426 Advances in Experimental Social Psychology, 33, 41-114.
- Duckitt, J., Wagner, C., Du Plessis, I., & Birum, I. (2002). The psychological bases of ideology
- and prejudice: testing a dual process model. *Journal of Personality and Social Psychology*,
- 429 83(1), 75. (doi: 10.1037//0022-3514.83.1.75)
- Fessler, D. M. T., Pisor, A. C., & Navarrete, C. D. (2014). Negatively-biased credulity and the
- cultural evolution of beliefs. *PLoS ONE*, *9*(4), e95167. (doi:
- 432 10.1371/journal.pone.0095167)
- Fessler, D. M. T. (2006). Steps toward the evolutionary psychology of a culture-dependent
- species. In P. Carruthers, S. Laurence, & S. Stich (Eds.), *The Innate Mind: Culture and*
- 435 *Cognition Vol. II* (pp. 91-117). New York: Oxford University Press.
- 436 Fessler, D. M. T., Holbrook, C., Pollack, J. S., & Hahn-Holbrook, J. (2014). Stranger danger:
- Parenthood increases the envisioned bodily formidability of menacing men. *Evolution and*
- 438 *Human Behavior*, 35(2), 109-117. (doi: 10.1016/j.evolhumbehav.2013.11.004)
- Graham, J., Haidt, J., & Nosek, B. A. (2009). Liberals and conservatives rely on different sets of
- 440 moral foundations. *Journal of Personality and Social Psychology*, 96(5), 1029. (doi:
- 441 10.1037/a0015141)
- 442 Hibbing, J. R., Smith, K. B., & Alford, J. R. (2014). Differences in negativity bias underlie
- variations in political ideology. *Behavioral and Brain Sciences*, 37(03), 297-307.
- 444 (doi:10.1017/S0140525X13001192)
- 445 Hilbig, B. E. (2009). Sad, thus true: Negativity bias in judgments of truth. *Journal of*
- Experimental Social Psychology, 45(4), 983-986. (doi:10.1016/j.jesp.2009.04.012)

447 Hilbig, B. E. (2012a). Good things don't come easy (to mind): Explaining framing effects in judgments of truth. Experimental Psychology, 59(1), 38-46. (doi: 10.1027/1618-448 449 3169/a000124) 450 Hilbig, B. E. (2012b). How framing statistical statements affects subjective veracity: Validation 451 and application of a multinomial model for judgments of truth. Cognition, 125(1), 37-48. 452 (doi:10.1016/j.cognition.2012.06.009) 453 Jones, J., & Saad, L. (2016). Gallup Poll Social Series: Values and Beliefs. Retrieved October 454 20, 2016 from http://www.gallup.com/poll/191741/democrats-liberal-social-issues-455 economic-ones.aspx?g source=liberal&g medium=search&g campaign=tiles 456 Federico, C. M., Hunt, C. V., & Ergun, D. (2009). Political expertise, social worldviews, and 457 ideology: Translating "competitive jungles" and "dangerous worlds" into ideological 458 reality. Social Justice Research, 22(2), 259-279. (doi: 10.1007/s11211-009-0097-0) 459 Kemmelmeier, M. (2008). Is there a relationship between political orientation and cognitive 460 ability? A test of three hypotheses in two studies. Personality and Individual Differences, 45(8), 461 767-772. (doi: 10.1016/j.paid.2008.08.003) 462 Knoll, B. R., O'Daniel, T. J., & Cusato, B. (2015). Physiological responses and political 463 behavior: Three reproductions using a novel dataset. Research & Politics, 2(4), 464 2053168015621328. (doi: 10.1177/2053168015621328) 465 Kurzban, R. (2007). Representational epidemiology: Skepticism and gullibility. In S. W. 466 Gangestad & J. A. Simpson (Eds.), The evolution of mind: fundamental questions and 467 controversies (pp. 357-362). New York: The Guilford Press. 468 Lilienfeld, S. O., & Latzman, R. D. (2014). Threat bias, not negativity bias, underpins 469 differences in political ideology. Behavioral and Brain Sciences, 37(03), 318-319. (doi:

- 470 10.1017/S0140525X1300263X)
- 471 Mills, M., Gonzalez, F. J., Giuseffi, K., Sievert, B., Smith, K. B., Hibbing, J. R., & Dodd, M. D.
- 472 (2016). Political conservatism predicts asymmetries in emotional scene memory.
- 473 *Behavioural and Brain Research*, 306(1), 84-90. (doi: 10.1016/j.bbr.2016.03.025)
- 474 Mills, M., Smith, K. B., Hibbing, J. R., & Dodd, M. D. (2014). The politics of the face-in-the-
- 475 crowd. Journal of Experimental Psychology: General, 143(3), 1199. (doi:
- 476 10.1037/a0035177)
- Peer, E., Samat, S., Brandimarte, L., & Acquisti, A. (2015). Beyond the Turk: An empirical
- 478 comparison of alternative platforms for crowdsourcing online behavioral research.
- Retrieved July 26, 2016 from http://papers.ssrn.com/sol3/Papers.cfm?abstract_id=2594183
- 480 R Core Team (2016). R: A language and environment for statistical computing. Retrieved
- September 15, 2016 from http://www.r-project.org/
- Reed, A. E., Chan, L., & Mikels, J. A. (2014). Meta-analysis of the age-related positivity effect:
- Age differences in preferences for positive over negative information. *Psychology and*
- 484 *Aging*, 29(1), 1-15. (doi: 10.1037/a0035194)
- Rozin, P., & Royzman, E. B. (2001). Negativity bias, negativity dominance, and contagion.
- 486 Personality & Social Psychology Review, 5(4), 296-320. (doi:
- 487 10.1207/S15327957PSPR0504 2)
- Sedek, G., Kossowska, M., & Rydzewska, K. (2014). The importance of adult life-span
- perspective in explaining variations in political ideology. *Behavioral and Brain Sciences*,
- 490 37(03), 329-330. (doi:10.1017/S0140525X13002732)
- 491 Stubbersfield, J. M., Tehrani, J. J., & Flynn, E. G. (2015). Serial killers, spiders and cybersex:
- Social and survival information bias in the transmission of urban legends. *British Journal*

493	of Psychology, 106(2), 288-307. (doi: 10.1111/bjop.12073)
494	Sydell, L. (2016, November 23). We tracked down a fake-news creator in the suburbs. Here's
495	what we learned. National Public Radio. Retrieved January 2, 2017 from
496	http://www.npr.org/sections/all techconsidered/2016/11/23/503146770/npr-finds-the-head-considered/2016/11/23/504040/npr-finds-the-head-considered/2016/11/23/5040/npr-finds-the-head-considered/2016/11/2
497	of-a-covert-fake-news-operation-in-the-suburbs.
498	van Leeuwen, F., & Park, J. H. (2009). Perceptions of social dangers, moral foundations, and
499	political orientation. Personality and Individual Differences, 47(3), 169-173. (doi:
500	10.1016/j.paid.2009.02.017)
501	Weeden, J., & Kurzban, R. (2016). Do people naturally cluster into liberals and conservatives?
502	Evolutionary Psychological Science, 2(1), 47-57. (doi: 10.1007/s40806-015-0036-2)
503	White, M. P., Pahl, S., Buehner, M., & Haye, A. (2003). Trust in risky messages: The role of
504	prior attitudes. Risk Analysis, 23(4), 717-726. (doi: 10.1111/1539-6924.00350)
505	Wilson, G. D., & Patterson, J. R. (1968). A new measure of conservatism. British Journal of
506	Social and Clinical Psychology, 7(4), 264-269. (doi: 10.1111/j.2044-8260.1968.tb00568.x)

Supplementary Online Materials

to accompany

Political Orientation Predicts Credulity Regarding Putative Hazards

Daniel M.T. Fessler, Anne C. Pisor, and Colin Holbrook

The complete dataset, list of variables, and analytic code employed in this project are archived at osf.io/qqq82 and http://escholarship.org/uc/item/82j5p9r3

Table of contents

Survey instrument

Credulity Index

Modified versions of Dodd et al.'s (2012) Wilson-Patterson Issues Index

Modified version of Dodd et al.'s (2012) Social Principles Index

Demographics

Note: Study 2 contained items taken from the Raven's Progressive Matrices (Raven,

Raven, & Court, 1998; 16 items) and the Wonderlic Cognitive Ability Test (1992; 10

items). Because the authors of this paper do not have permission to republish these

instruments, readers who wish to know which items from these instruments were

employed in Study 2 should contact the authors directly.

Appendix 1. Categories of conservatism based on a modified version of Dodd et al.'s (2012)

Wilson-Patterson Issues Index

Appendix 2. Addressing outliers

Tables S1a, S1b. Descriptive statistics, Studies 1 and 2

Tables S2a, S2b. Descriptive statistics for excluded participants, Studies 1 and 2

Table S3. Parameter estimates, 95% confidence intervals, and p values for logistic model exploring predictors of being excluded for incomplete responses, not speaking English as a first language, repeat participation, and not answering catch questions

Table S4. Parameter estimates, 95% confidence intervals, and p values for models with the political summary measure as a predictor, weighted hazard credulity minus weighted benefit credulity as the outcome, full model

Table S5. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor, weighted hazard credulity minus weighted benefit credulity as the outcome, full model with no imputation

Tables S6a, S6b. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor, weighted credulity (S6a: hazard; S6b: benefit) as the outcome

Table S7. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor, weighted hazard credulity minus weighted benefit credulity as the outcome

Table S8. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor, with the unweighted difference between a participant's hazard and benefit credulity as the response.

Tables S9a, S9b. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor, with a participant's credulity for each item (S9a: hazard; S9b: benefit) as the response

Table S10. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor and parenthood status as a control, weighted hazard credulity minus weighted benefit credulity as the outcome, Study 2

Table S11. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor, weighted hazard credulity minus weighted benefit credulity as the outcome, including all liberals excluded for Study 2 analyses

Table S12. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor interacting with sex of the participant, weighted hazard credulity minus weighted benefit credulity as the outcome

Table S13. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor interacting with age of the participant, weighted hazard credulity minus weighted benefit credulity as the outcome

Figure S1. The estimated effect of political orientation on the difference between hazard and benefit credulity with the terrorism item excluded

Figure S2. The estimated effect of political orientation on hazard credulity with the terrorism item excluded

Figure S3a,b. LOESS fit of weighted hazard credulity by age

(Credulity Index)

(Each item was presented on a single web page, and the order of items was randomized)

Below are a series of statements collected from the media. Some of these statements are true, and some of them are false. For each of the statements, please indicate, by checking the appropriate box, how confident you are that the statement is true or false. Also, for each of the statements, please indicate how significant you think the things described in the statement are. Please note that your answers to each of these two questions should be independent of each other. For example, you might decide that you're absolutely certain that a statement is true, and select 7 for this question, but also feel that the risk described in the statement is small, and select 1 for this question.

1. Storing batteries in a refrigerator or freezer will improve their performance.

<i>8</i> -		<i>8</i>		r			
1	2	3	4	5	6	7	
I'm absolutely certain this statement is F .						'm absolutely certain this nent is TRUE	
1	2	3	4	5	6	7	
The benefit						The benefit	
described in th	nis				des	scribed in this	
statement is SI	MALL				stateme	nt is LARGE	

2. Cell pho	nes damage	credit card m	agnetic strips,	making them	unusable.						
1	2	3	4	5	6	7					
I'm absolutely	I				I	'm absolutely					
certain this						certain this					
statement is FALSE statement is TRUE											
1	2	3	4	5	6	7					
The risk						The risk					
described in th	nis				des	scribed in this					
statement is S	MALL				stateme	nt is LARGE					
3. Eating ca	arrots results	in significan	tly improved	vision.							
1	2	3	4	5	6	7					
I'm absolutely	I				I	'm absolutely					
certain this						certain this					
statement is F	ALSE				staten	nent is TRUE					
1	2	3	4	5	6	7					
The benefit						The benefit					
described in th	nis				des	scribed in this					
statement is S	MALL	statement is SMALL statement is LARGE									

4. Kale cor	ntains thalliu	m, a toxic hea	avy metal, tha	t the plant abso	orbs from soil										
1	2	3	4	5	6	7									
I'm absolutely	V				Ι	'm absolutely									
certain this						certain this									
statement is F	ALSE				staten	nent is TRUE									
1	2	3	4	5	6	7									
The risk						The risk									
described in the	nis				des	scribed in this									
statement is S	MALL				stateme	nt is LARGE									
5. Exercisi	ng on an emp	oty stomach b	ourns more cal	lories.											
1	2	3	4	5	6	7									
I'm absolutely	7				Ι	'm absolutely									
certain this						certain this									
statement is F	ALSE				staten	nent is TRUE									
1	2	3	4	5	6	7									
The benefit						The benefit									
described in the	his				des	scribed in this									
statement is S	MALL				stateme	statement is SMALL statement is LARGE									

6. Long-dis	stance runnin	g causes oste	oarthritis of the	ne knees.							
1	2	3	4	5	6	7					
I'm absolutely	7				I	'm absolutely					
certain this						certain this					
statement is F	ALSE				staten	nent is TRUE					
1	2	3	4	5	6	7					
The risk						The risk					
described in th	nis				des	scribed in this					
statement is S	MALL				stateme	nt is LARGE					
7. Selecting	g credit cards	that have a l	ow credit limi	it improves on	e's credit sco	re.					
1	2	3	4	5	6	7					
I'm absolutely	<i>I</i>				I	'm absolutely					
certain this						certain this					
statement is F	ALSE				staten	nent is TRUE					
1	2	3	4	5	6	7					
The benefit						The benefit					
described in th	nis				des	scribed in this					
statement is S	MALL				stateme	nt is LARGE					

8. Hotel roo	om keycards	are often end	coded with per	rsonal informa	tion that can	be read by thieves.
1	2	3	4	5	6	7
I'm absolutely	y]	I'm absolutely
certain this						certain this
statement is F	ALSE				stater	ment is TRUE
1	2	3	4	5	6	7
The risk						The risk
described in the	his				de	escribed in this
statement is S	MALL				stateme	ent is LARGE
9. People w	vho own cats	live longer the	han people wl	ho don't.		
1	2	3	4	5	6	7
I'm absolutely		3	т	3		I'm absolutely
certain this	y					certain this
statement is F	ALSE				stater	ment is TRUE
5.60.0 10 1	11202				3	
1	2	3	4	5	6	7
The benefit						The benefit
described in the	his				de	escribed in this
statement is S	MALL				stateme	ent is LARGE

10. Sharks	pose a signif	icant risk to b	beachgoers.				
1	2	3	4	5	6	7	
I'm absolutely					I	'm absolutely	
certain this						certain this	
statement is F .	ALSE				staten	nent is TRUE	
1	2	3	4	5	6	7	
The risk						The risk	
described in th	nis				de	scribed in this	
statement is Si	MALL				stateme	ent is LARGE	
11. Stockw	ood, Califor	nia is one of t	he safest citie	es in the U.S.			
1	2	3	4	5	6	7	
I'm absolutely					I	'm absolutely	
certain this						certain this	
statement is F.	ALSE				staten	nent is TRUE	
1	2	3	4	5	6	7	
The benefit						The benefit	
described in th	nis				de	scribed in this	
statement is SI	MALL				stateme	ent is LARGE	

12. Terrorist	t attacks in t	the U.S. have	increased sin	ce Sept 11, 20	001.		
1	2	3	4	5	6	7	
I'm absolutely					I	'm absolutely	
certain this						certain this	
statement is F A	ALSE				staten	nent is TRUE	
1	2	3	4	5	6	7	
The risk						The risk	
described in thi	S				des	scribed in this	
statement is SM	IALL				stateme	nt is LARGE	
13. When fly business class			ou are more li	kely to be upg	graded from ed	conomy to	
1	2	3	4	5	6	7	
I'm absolutely					I	'm absolutely	
certain this						certain this	
statement is F A	ALSE				staten	nent is TRUE	
1	2	3	4	5	6	7	
The benefit						The benefit	
described in thi	S				des	scribed in this	
statement is SN	1ALL				stateme	nt is LARGE	

			partially open gen masks to		on a commerci	al jetliner, causing
1	2	3	4	5	6	7
I'm absolutely	y				Ι	'm absolutely
certain this						certain this
statement is F	FALSE				staten	nent is TRUE
1	2	3	4	5	6	7
The risk						The risk
described in t	his				des	scribed in this
statement is S	MALL				stateme	nt is LARGE
		ı hard-topped metal inside t		protection fro	m lightning, a	s long as the
1	2	3	4	5	6	7
I'm absolutely	y				Ι	'm absolutely
certain this						certain this
statement is F	FALSE				staten	nent is TRUE
1	2	3	4	5	6	7
The benefit						The benefit
described in the	his				des	scribed in this
statement is S	MALI.				stateme	nt is LARCE

16. In the U	.S., an avera	age of 32 peop	ple are killed	by lightning ea	ach year.		
1	2	3	4	5	6	7	
I'm absolutely					I	m absolutely	
certain this						certain this	
statement is F A	ALSE				statem	nent is TRUE	
1	2	3	4	5	6	7	
The risk						The risk	
described in the	is				des	scribed in this	
statement is SN	ЛАТТ.				stateme	nt is LARGE	

In the following sections, please tell us about yourself.

Your gender:
Female
Male
Your age:
How many letters are in the English alphabet?

(Study 1: Wilson-Patterson Issues Index – modified from Dodd et al. [2012])

Please indicate whether you agree or disagree, or are uncertain, with regard to each topic listed below:

1.	school prayer: agreedisagree	uncertain
2.	school prayer: agree disagree pacifism: agree disagree unc	ertain
3.	socialism: agree disagree unc	ertain
4.	pornography:agreedisagree illegal immigration:agreedisagree	uncertain
5.	illegal immigration: agreedisagree	uncertain
6.	women's equality: agreedisagree	uncertain
7	death nenalty: agree disagree	uncertain
8.	The Patriot Act:agreedisagree premarital sex:agreedisagree gay marriage:agreedisagreedisagree	uncertain
9.	premarital sex:agreedisagree	uncertain
10.	gay marriage: agreedisagree	uncertain
11.	abortion rights: agree disagree	uncertain
12.	evolution:agreedisagreeunc	ertain
	patriotism: agreedisagreeunc	
14.	Biblical truth: agreedisagree	uncertain
15.	2003 Iraq invasion': agreedisagree	uncertain
	welfare spending:agreedisagree	
17.	tax cuts:agreedisagreeunc gun control:agreedisagreeunc	ertain
18.	gun control:agreedisagreeunc	ertain
19.	military spending: agreedisagree	uncertain
20.	warrantless searches: agree disagree globalization: agree disagree	igreeuncertain
21.	globalization: agreedisagree	uncertain
22.	pollution control: agree disagree small government: agree disagree	uncertain
23.	small government:agreedisagree	uncertain
	school standards: agreedisagree	
25.	foreign aid: agreedisagreeunc	ertain
26.	free trade:agreedisagreeunc	ertain
27.	obedience to authorities ² :agreedisa compromise with enemies ³ :agreedisa	igreeuncertain
28.	compromise with enemies : agreedisa	igreeuncertain
	charter schools ⁴ : agreedisagree	uncertain
¹ Modi	fied from Dodd et al. 's original "Iraq"	
-	fied from Dodd et al's original "obedience"	
	fied from Dodd et al.'s original "compromis	
4 Repla	aces Dodd et al.'s original "school standards	s"

(Study 2: Wilson-Patterson Issues Index – modified from Dodd et al. [2012])

Please indicate whether you agree or disagree, or are uncertain, with regard to each topic listed below:

1. school prayer: agreedisagreeuncertain
2. pacifism:agreedisagreeuncertain
3. socialism:agreedisagreeuncertain
4. pornography: agreedisagreeuncertain
5. illegal immigration: agreedisagreeuncertain
6. women's equality:agreedisagreeuncertain
7. death penalty: agreedisagreeuncertain
8. use nuclear weapons against threats to the U.S. ¹ : agreedisagreeuncertain
9. premarital sex:agreedisagreeuncertain
10. gay marriage:agreedisagreeuncertain
11. abortion rights: agreedisagreeuncertain
12. evolution: agreedisagreeuncertain
13. patriotism:agreedisagreeuncertain
14. Biblical truth:agreedisagreeuncertain
15. bomb cities controlled by terrorists ¹ :agreedisagreeuncertain
16. welfare spending: agreedisagreeuncertain
17. tax cuts: agreedisagreeuncertain
18. waterboarding terror suspects ² :agreedisagreeuncertain
19. gun control: agreedisagreeuncertain
20. military spending: agreedisagreeuncertain
21. warrantless searches: agreedisagreeuncertain
22. globalization:agreedisagreeuncertain
23. pollution control: agreedisagreeuncertain
24. small government: agreedisagreeuncertain
25. charter schools ¹ : agreedisagreeuncertain
26. foreign aid: agreedisagreeuncertain
27. free trade: agreedisagreeuncertain
28. drone strikes ² : agreedisagreeuncertain
29. obedience to authorities ³ : agreedisagreeuncertain
30. compromise with enemies ⁴ :agreedisagreeuncertain
Modified from Dodd et al.'s original to increase relevance to contemporary politics
² Added to increase relevance to contemporary politics
³ Modified from Dodd et al's original "obedience"

 $^{^{2}}$ Aa

⁴ Modified from Dodd et al.'s original "compromise"

(Social Principles Index – slightly modified* from Dodd et al. [2012])

Please tell us your opinions regarding how society works best by selecting one of the two options in each of the following statements:

Society works best when...

- 1-People live according to traditional values
- 2-People adjust their values to fit changing circumstances

Society works best when...

- 1-Behavioral expectations are based on an external code
- 2-Behavioral expectations are allowed to evolve over the decades

Society works best when...

- 1-Our leaders stick to their beliefs regardless
- 2-Our leaders change positions whenever situations change

Society works best when...

- 1-We take care of our own people first
- 2-We realize that people everywhere deserve our help

Society works best when...

- 1-Those who break the rules are punished
- 2-Those who break the rules are forgiven

Society works best when...

- 1-Every member contributes
- 2-More fortunate members sacrifice to help others

Society works best when...

- 1-People are rewarded according to merit
- 2-People are rewarded according to need

Society works best when...

- 1-People take primary responsibility for their welfare
- 2-People join together to help others

Society works best when...

- 1-People are proud they belong to the best society there is
- 2-People realize that no society is better than any other

Society works best when...

- 1-Our leaders are obeyed
- 2-Our leaders are questioned

Society works best when...

- 1-Our leaders call the shots
- 2-Our leaders are forced to listen to others

Society works best when...

- 1-People recognize the unavoidable flaws of human nature
- 2-People recognize that humans can be changed in positive ways

Society works best when...

- 1-Our leaders compromise with their opponents in order to get things done
- 2-Our leaders adhere to their principles no matter what

* Because it directly addresses belief in a dangerous world, the following item from Dodd et al.'s original measure was omitted from the survey:

Society works best when...

- 1-People realize the world is dangerous
- 2-People assume all those in faraway places are kindly

How w	ould you r	ate your o	overall po	olitical ori	entation	?		
o	0	0	o	0	o	0	o	o
Extremely	7			Moderate				Extremely
Liberal								Conservative
Please s	select the t	term that l	est desc	ribes your	political	l affiliatio	n:	
Repu	ıblican							
Dem	ocratic							
Tea l	Party							
_Libe	rtarian							
Gree	n							
_Othe	r (please i	ndicate) _		-				
None	e / not affi	liated with	n any pol	litical part	y			
Do you	consider	yourself a	n Ameri	can?				
-	Yes							
	Somewha No	t						
	110							
Is Engl	ish your fi	rst langua	ge?					
-	Yes							
-	No							
Vorm 54	hnisit							
	chnicity:							
-	African-A	merican						

- Middle Eastern - Pacific Islander
- South Asian / Indian

- Hispanic / Latin American

- White

- Asian

- More than one
- Other

Annual household income:

- under \$20,000
- \$20 \$30,000
- \$30 \$40,000
- \$40 \$50,000
- \$50 \$60,000
- \$60 \$70,000
- \$70 \$80,000
- \$80 \$90,000
- \$90 \$100,000
- \$100 \$110,000
- \$110 \$120,000
- \$120 \$130,000
- \$130 \$140,000
- \$140 \$150,000
- \$150 \$160,000
- \$170 \$180,000
- \$180 \$190,000
- \$190 \$200,000
- \$200 \$210,000
- \$210 \$220,000
- \$220 \$230,000
- \$230 \$240,000
- \$240 \$250,000
- \$250 \$260,000
- \$260 \$270,000
- \$270 \$280,000
- \$280 \$290,000
- \$290 \$300,000
- over \$300,000

Education:

- Middle school or less
- Some High School
- High School Graduate
- Some college
- AA degree
- College graduate
- Some graduate school
- Master's degree
- Advanced degree (e.g., Ph.D.)

SOM: Conservatism and Credulity

How man	y letters	are in the	word "ob	oligatory"	??		_	
What is y	our heigh	ht, to the 1	nearest ha	lf-inch?				
Fe	eet:	Inc	hes:					
(Study 1)	How sur	prised wo	ould you b	oe to see s	someone (eat lunch	in the af	ternoon?
o	0	o	o	o	o	0	0	0
Not surprise	d							Extremely
at all								surprised

Are you a parent?
- Yes - No
Study 1: Yes →) Please answer the following questions about your family.
Study 1) Are you currently raising a baby in your home?
- Yes - No
Study 1) How many girls have you had?
Study 1) How many boys have you had?
Study 1) How many girls have you personally raised?
Study 1) How many boys have you personally raised?
Study 1) How old were you when had your first child?
Study 1) How old is your YOUNGEST child, in years? (If an infant, please specify that you are answering in months, e.g., "8 months"):
Study 1) What is the gender of your YOUNGEST child?
Study 1) How old is your OLDEST child, in years?
Study 1) If you have only had one child, please type "NA": What is the gender of your DLDEST child?

Appendix 1. Categories of conservatism based on a modified version of Dodd et al.'s (2012) Wilson-Patterson issues index.

For Study 1, we sorted 25 of 28 items from the modified Wilson-Patterson issues index into three types of conservatism:

Social conservatism: school prayer, pornography, illegal immigration, women's equality,
 premarital sex, gay marriage, abortion rights, evolution, biblical truth, gun control
 Economic conservatism: socialism, welfare spending, tax cuts, globalization, pollution control,
 small government, foreign aid

Military, obedience, and punishment conservatism: pacifism, death penalty, Patriot Act, patriotism, the 2003 Iraq invasion, military spending, obedience, compromise

We omitted items concerning free trade and charter schools (our modification to the school standards item), as neither discriminated between liberals and conservatives. An item concerning warrantless search was also omitted as it did not load onto any of the three categories described above.

For Study 2, we removed the Iraq invasion question as its continuing relevance is questionable, but added other items intended to gauge international military involvement. We sorted 26 of 30 items from the modified Wilson-Patterson issues index into three types of conservatism:

Social conservatism: school prayer, pornography, illegal immigration, women's equality,
 premarital sex, gay marriage, abortion rights, evolution, biblical truth, gun control
 Fiscal conservatism: socialism, welfare spending, tax cuts, globalization, pollution control, small government, foreign aid

Military, obedience, and punishment conservatism: pacifism, death penalty, Patriot Act,
patriotism, military spending, obedience, compromise, use nuclear weapons against
threats to the U.S., bomb cities controlled by terrorists, waterboarding terror suspects,
drone strikes

We omitted items concerning free trade and globalization, as neither discriminated between liberals and conservatives.

We summarized each of the three above categories using principal components analysis. For Study 1, the social conservatism principal component summarized 43.68% of the variance with variable loadings between .39-.80, the economic conservatism principal component summarized 33.42% of the variance with variable loadings between .40-.72, and the military/obedience/punishment conservatism principal component summarized 35.63% of the variance with variable loadings between .45-.68. For Study 2, the social conservatism principal component summarized 45.55% of the variance with variable loadings between .28-.82, the economic conservatism principal component summarized 35.82% of the variance with variable loadings between .41-.71, and the military/obedience/punishment conservatism principal component summarized 37.34% of the variance with variable loadings between .45-.72.

Appendix 2. Addressing outliers

When exploratory data analysis revealed outliers, these points were rounded up or down to lower their influence on model fit. In Study 1, extreme positive values for participant age, income, education, social conservatism, and the Wilson-Patterson issues index were rounded down to the 97.5th percentile (i.e., ages rounded to 65, income rounded to the 15th increment (\$160,000), advanced degrees lumped with some advanced degree study, social conservatism rounded to 5, and Wilson-Patterson rounded to 17). Very low values for education, i.e., five individuals who had not completed high school, were lumped with high school graduates. Likewise, in Study 2, 3 individuals who had not completed high school were lumped with high school graduates, and 7 individuals with a doctoral degree were lumped with master's degree recipients. We also rounded down participants with the highest incomes to the 97.5th percentile (income increment 18, or incomes larger than \$200,000 annually) and rounded up participants with the lowest Raven's matrices and Wonderlic scores to the 2.5th percentile (-2.23 and -1.90 standard deviations, respectively).

Appendix 3. Imputation, random seeds, and random culling in Study 2

Missing values were imputed for participants who failed to respond to less than 10% of the credulity items, less than 10% of the issues items, and less than 10% of the social principles index; values were also imputed for participants who failed to provide their political orientation (Study 1 n = 3, Study 2 n = 0), political category (Study 1 n = 3, Study 2 n = 6), income (Study 1 n = 1, Study 2 n = 5), or education (Study 1 n = 7, Study 1 n = 1). Imputation was performed via predictive mean matching (Van Buuren and Groothuis-Oudshoorn, 2011): in this approach, given all participants' responses, the function generates a mean prediction for one participant's missing value (Little, 1988). Imputation was performed five times for each missing value and the mean of these five imputations kept as the final value. Participants with imputed values are included in all models except in the model reported in Table S5. Predictive mean matching relies on a random number generator. We initialize the generator with five different seed values. Results reported were generated using the third seed. In Study 2, we randomly eliminate participants to achieve a sample that is approximately nationally representative in terms of social political orientation (Jones and Saad, 2016). We perform this process five times, and note where results were altered by the sample selected.

 Table S1a. Study 1: descriptive statistics.

	1.6	αD	16 1	1.6:	1.6	3.7	%	%	%	%	%	NT /
Variable	Mean	SD	Median	Min	Max	N	level 1	level 2	level	level 4	level 5	Notes
Credulity difference	.28	5.34	.13	-17.38	15.63	449	_	_				Weighted avg. hazards - weighted avg. benefits
Wtd. avg. hazard credulity	12.30	5.07	11.69	1.00	29.63	456						Weighted by centrists' perceived hazardousness
Wtd. avg. benefit credulity	11.98	4.63	11.50	1.75	27.63	463						Weighted by centrists' perceived beneficialness
Cred. difference (unweighted)	19	.85	25	-2.88	2.38	459						Avg. hazards - avg. benefits
Political summary	.00	1.70	31	-2.87	4.51	472						Principal component of the four politics measure
Pol. summary (Non-PCA)	.01	3.43	69	-5.64	8.89	444						Summary of the four politics measures
"Society works best"	-3.45	5.46	-3.00	-13.00	13.00	466						Positive values more conservative
Political Likert	3.99	2.20	4.00	1.00	9.00	471						1 = extremely liberal 9 = extremely conservative
Political category	NA	NA	.00	NA	NA	469	.49	.32	.19			1=liberal party 2=libertarian or unaffiliated 3=conservative party
Wilson- Patterson index	-5.21	9.80	-6.00	-25.00	16.68	454						Positive values more conservative
Social conservatism	01	2.06	70	-2.20	5.03	472						Principal component of sub-measure of Wilson-
Fiscal	00	1 53	- 04	-2 85	3 89	472						Patterson index Principal component of

Parenthood	NA	NA	1.00	NA	NA	472	.61	.39				1=no, 2=yes. 3=no reply
												college, 3=associate's, 4=bachelor's, 5=at least some advanced degree
Education	NA	NA	4.00	NA	NA	465	.15	.09	.36	.26	.14	1=high school, 2=some
Ethnicity	NA	NA	2.00	NA	NA	472	.19	.81				1=other, 2=white
Sex	NA	NA	2.00	NA	NA	472	.48	.52				1=female, 2=male
Income	3.78	3.57	3.00	.00	14.00	471						
Age	17.03	11.81	14.00	.00	46.00	472						Given in years
Military conservatism	.00	1.69	.00	-3.51	3.51	472						Principal component of sub-measure of Wilson- Patterson index
												Patterson index

Table S1b. Study 2: descriptive statistics for subsample excluding randomly omitted liberals.

Variable	Mean	SD	Median	Min	Max	N	% level	% level 2	% level	% level 4	% level 5	% level 6	Notes
Credulity difference	1.24	4.96	1.13	-12.88	22.13	451	1	2	J	7	3	O	Weighted avg. hazards - weighted avg. benefits
Wtd. avg. hazard credulity	12.67	5.21	12.25	1.88	34.63	461							Weighted by centrists' perceived hazardousness
Wtd. avg. benefit credulity	11.46	4.59	11.00	1.63	33.00	466							Weighted by centrists' perceived beneficialness
Cred. difference (unweighted)	02	.82	.00	-2.25	3.88	465							Avg. hazards - avg. benefits
Political Summary	.00	1.73	16	-3.57	3.97	476							Principal component of the four politics measures
Pol. summary (Non-PCA)	.02	3.48	24	-7.02	7.81	450							Summary of the four politics measures
"Society works best"	-2.00	5.86	-3.00	-13.00	13.00	467							Positive values more conservative
Political Likert	4.98	2.14	5.00	1.00	9.00	476							1 = extremely liberal 9 = extremely conservative
Political category	NA	NA	.00	NA	NA	470	.36	.36	.29				1=liberal party 2=libertarian or unaffiliated 3=conservative party
Wilson- Patterson index	-2.93	10.47	-4.00	-25.00	22.00	463							Positive values more conservative
Social conservatism	.00	2.14	92	-2.28	5.45	476							Principal component of sub-measure of Wilson- Patterson index
Fiscal conservatism	.00	1.60	.01	-3.52	3.35	476							Principal component of sub-measure of Wilson-Patterson index
Military conservatism	.00	1.82	02	-4.05	3.38	476							Principal component of sub-measure of Wilson-

													Patterson index
Raven's test	.02	.96	.16	-2.13	1.31	469							Correct - incorrect
Wonderlic test	.01	.98	.08	-2.00	1.50	464							Correct - incorrect
Age	34.32	12.47	31.00	18.00	73.00	476							Given in years
Income	4.60	4.07	4.00	.00	17.00	471							
Sex	NA	NA	2.00	NA	NA	476	.40	.60					1=female, 2=male
Ethnicity	NA	NA	2.00	NA	NA	476	.21	.79					1=other, 2=white
Education	NA	NA	4.00	NA	NA	475	.10	.28	.08	.35	.04	.14	1=high school, 2=some college, 3=associate's, 4=bachelor's, 5=at least some advanced degree. 6 = advanced degree
Parenthood	NA	NA	1.00	NA	NA	380	.59	.41					1=no, 2=yes. 3=no reply

 Table S2a. Study 1: descriptive statistics for participants excluded from analyses.

Variable	Mean	SD	Median	Min	Max	N	%	%	%	%	%	Notes

							level	level 2	level 3	level 4	level 5	
Credulity difference	46	4.23	-1.38	-9.00	8.25	37	1	2	5	•	J	Weighted avg. hazards - weighted avg. benefits
Wtd. avg. hazard credulity	11.71	4.55	12.44	3.75	21.00	38						Weighted by centrists' perceived hazardousness
Wtd. avg. benefit credulity	12.01	4.03	12.00	1.50	19.25	42						Weighted by centrists' perceived beneficialness
Cred. difference (unweighted)	28	.71	25	-1.88	1.00	37						Avg. hazards - avg. benefits
Political PCA	.02	1.01	.00	-2.84	3.27	65						Principal component of the following four measures
Pol. summary (Non-PCA)	04	2.59	.35	-5.45	5.47	30						Summary of the four politics measures
"Society works best"	-3.87	4.75	-3.00	-13.00	7.00	30						Positive values more conservative
Political Likert	3.94	1.85	4.00	1.00	8.00	32						1 = extremely liberal9 = extremely conservative
Political category	NA	NA	.00	NA	NA	32	.28	.50	.22			1=liberal party 0=libertarian or unaffiliated 3=conservative party
Wilson- Patterson index	-6.20	7.95	-7.50	-19.00	16.78	36						Positive values more conservative
Social conservatism	01	1.49	.00	-2.26	5.18	65						Principal component of sub- measure of Wilson- Patterson index
Fiscal conservatism	05	1.19	.00	-2.90	4.09	65						Principal component of sub- measure of Wilson- Patterson index
Military conservatism	18	1.15	.00	-2.31	2.87	65						Principal component of sub- measure of Wilson- Patterson index

Age	32.32	10.42	12.00	19.00	64.00	37						Given in years
Income	3.86	4.10	2.00	.00	14.00	29						
Sex	NA	NA	2.00	NA	NA	37	.43	.57				1=female, 2=male
Ethnicity	NA	NA	2.00	NA	NA	32	.44	.56				1=other, 2=white
Education	NA	NA	4.00	NA	NA	30	.13	.37	.03	.37	.10	1=high school, 2=some college, 3=associate's, 4=bachelor's, 5=at least some advanced degree
Parenthood	NA	NA	1.00	NA	NA	33	.70	.30				1=no 2=yes 3=no reply

Table S2b. Study 2: descriptive statistics for participants excluded from analysis (prior to exclusion of liberals or centrists).

Variable	Mean	SD	Median	Min	Max	N	% level	% level 2	% level 3	% level 4	% level 5	% level	Notes
Credulity difference	.75	5.36	0.63	-11.13	15.38	53	1	2	3	4	3	6	Weighted avg. hazards - weighted avg. benefits
Wtd. avg. hazard credulity	12.84	4.94	12.63	5.00	30.00	53							Weighted by centrists' perceived hazardousness
Wtd. avg. benefit credulity	11.92	4.69	11.81	3.63	23.63	58							Weighted by centrists' perceived beneficialness
Cred. difference (unweighted)	02	.91	13	-2.00	2.25	56							Avg. hazards - avg. benefits
Political PCA	05	1.61	41	-3.19	3.69	58							Principal component of the following four measures
Pol. summary (Non-PCA)	15	3.26	86	-6.25	7.27	47							Summary of the four politics measures
"Society works best"	-1.69	5.54	-1.00	-13.00	13.00	52							Positive values more conservative
Political Likert	4.86	2.26	5.00	1.00	9.00	58							1 = extremely liberal9 = extremely conservative
Political category	NA	NA	.00	NA	NA	57	.40	.28	.32				1=liberal party 0=libertarian or unaffiliated 3=conservative party
Wilson- Patterson index	-3.06	9.01	-4.00	-24.00	17.00	52							Positive values more conservative
Social conservatism	.13	2.02	69	-2.29	4.22	58							Principal component of sub-measure of Wilson- Patterson index
Fiscal conservatism	13	1.43	06	-3.50	3.30	58							Principal component of sub-measure of Wilson-Patterson index
Military conservatism	06	1.78	04	-3.44	3.33	58							Principal component of sub-measure of Wilson-

SOM: Conservatism and Credulity

													Patterson index
Raven's test	45	1.01	12	-2.27	1.31	53							Correct - incorrect
Wonderlic test	47	.98	35	-2.19	1.49	47							Correct - incorrect
Age	32.17	10.97	12.00	18.00	56.00	58							Given in years
Income	4.71	4.04	4.00	65.00	15.00	55							
Sex	NA	NA	2.00	NA	NA	58	.38	.62					1=female, 2=male
Ethnicity	NA	NA	2.00	NA	NA	57	.26	.74					1=other, 2=white
Education	NA	NA	4.00	NA	NA	55	.18	.35	.04	.29	.04	.11	1=high school, 2=some college, 3=associate's, 4=bachelor's, 5=at least some advanced degree. 6 = advanced degree
Parenthood	NA	NA	1.00	NA	NA	63	.52	.48					1=no 2=yes 3=no reply

Table S3. Parameter estimates, 95% confidence intervals, and *p* values for logistic model exploring predictors of being excluded for incomplete responses, not speaking English as a first language, repeat participation, and not answering catch questions.

			Study 1				Stud	ly 2
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	-1.78	-3.19	37	.01	-2.72	-4.03	-1.41	.00
Political summary	.10	14	.34	.42	09	31	.14	.45
Sex: Male	.27	53	1.08	.51	.61	14	1.37	.11
Age	04	09	.00	.08	.02	02	.05	.33
Ethnicity: White ¹	99	-1.80	18	.02	76	-1.52	01	.05
Income	.01	10	.13	.81	.03	06	.13	.49
Educ: Associate's	-1.01	-3.27	1.26	.38	-1.56	-3.82	.70	.18
Educ: Bachelor's	.02	-1.23	1.28	.97	22	-1.58	1.14	.75
Educ: Some associate's	.31	91	1.52	.62	05	-1.12	1.02	.93
Educ: Some adv. grad	04	-1.67	1.59	.96	.15	90	1.21	.77
Parenthood	.15	78	1.09	.75	42	-2.69	1.85	.72
Raven's test					38	76	.01	.06
Wonderlic test					46	87	04	.03

Study 1: N = 428. Study 2: N = 487. Effect not robust across iterations.

Table S4. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor, weighted hazard credulity minus weighted benefit credulity as the outcome, full model.

		Sti	udy 1			Sta	udy 2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	27	-2.09	1.54	.77	.63	-1.25	2.52	.51
Political summary	.36	.08	.65	.01	.54	.28	.81	.00
Sex: Male	57	-1.53	.39	.24	.52	43	1.46	.29
Age	.00	05	.04	.98	.03	01	.07	.16
Ethnicity: White	05	-1.28	1.17	.93	32	-1.44	.79	.57
Income	.04	11	.18	.63	01	12	.10	.87
Educ: Advanced degree					49	-2.37	1.39	.61
Educ: Associate's	.38	-1.60	2.36	.71	.14	-1.98	2.26	.90
Educ: Bachelor's	.18	-1.29	1.65	.81	.38	-1.22	1.98	.64
Educ: Some associate's	1.23	31	2.77	.12	.28	-1.34	1.89	.74
Educ: Some adv. grad	.99	85	2.82	.29	.12	-2.42	1.89	.74
Raven's test					05	59	.49	.86
Wonderlic test					15	70	.41	.61
Parenthood	.67	46	1.80	.25				

Study 1: N = 472. Adjusted $R^2 = .01$, F(10, 461) = 1.66, p = .09. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 476. Adjusted $R^2 = .03$, F(12, 463) = 2.09, p = .02. Women, "other" ethnicity, some high school/high school diploma. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes.

Table S5. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor, weighted hazard credulity minus weighted benefit credulity as the outcome, full model with no imputation.

		Sta	udy 1			Sti	udy 2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	63	-2.57	1.31	.53	1.05	99	3.08	.31
Political summary	.37	.07	.67	.01	.58	.30	.86	.00
Sex: Male	52	-1.53	.50	.32	.52	49	1.52	.32
Age	01	05	.04	.83	.03	01	.07	.17
Ethnicity: White	.22	-1.09	1.52	.75	23	-1.44	.98	.71
Income	.04	11	.19	.64	01	14	.11	.82
Educ: Advanced degree					81	-2.81	1.18	.42
Educ: Associate's	.59	-1.51	2.69	.58	06	-2.36	2.25	.96
Educ: Bachelor's	.35	-1.23	1.94	.66	17	-1.87	1.52	.84
Educ: Some associate's	1.33	31	2.97	.11	21	-1.91	1.49	.81
Educ: Some adv. grad	1.03	94	2.99	.31	64	-3.28	2.01	.64
Raven's test					20	78	.38	.51
Wonderlic test					18	77	.41	.54
Parenthood	.83	35	2.02	.17				

Study 1: N = 441. Adjusted $R^2 = .01$, F(10, 430) = 1.63, p = .09. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 432. Adjusted $R^2 = .03$, F(12, 419) = 2.19, p = .011. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes.

Table S6a. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor, weighted hazard credulity as the outcome.

		Study	y 1			Stud	y 2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	13.72	12.03	15.41	.00	12.74	10.87	14.60	.00
Political summary	.48	.22	.75	.00	.59	.33	.85	.00
Sex: Male	-1.74	-2.63	85	.00	18	-1.12	.75	.70
Age	.02	02	.06	.40	.07	.03	.11	.00
Ethnicity: White	88	-2.02	.26	.13	59	-1.70	.52	.30
Income	13	26	.00	.05	06	17	.05	.27
Educ: Advanced degree					48	-2.35	1.39	.61
Educ: Associate's	13	-1.97	1.71	.89	.25	-1.86	2.35	.82
Educ: Bachelor's	16	-1.52	1.21	.82	36	-1.94	1.23	.66
Educ: Some associate's	.73	71	2.16	.32	38	-1.98	1.22	.64
Educ: Some adv. grad	39	-2.09	1.31	.66	.13	-2.39	2.66	.92
Parenthood	.77	28	1.82	.15				
Raven's test					17	71	.36	.53
Wonderlic test					97	-1.52	42	.00

Study 1: N = 472. Adjusted $R^2 = .03$, F(10, 461) = 2.52, p = .006. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 476. Adjusted $R^2 = .12$, F(12, 463) = 6.61, p < .001. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes.

Table S6b. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor, weighted benefit credulity as the outcome.

		Study	y 1			Stud	v 2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	13.99	12.41	15.57	.00	12.10	10.37	13.83	.00
Political summary	.12	13	.37	.34	.05	19	.29	.69
Sex: Male	-1.17	-2.00	34	.01	70	-1.57	.17	.12
Age	.02	02	.06	.35	.04	.01	.08	.02
Ethnicity: White	83	-1.89	.24	.13	27	-1.29	.76	.61
Income	17	29	04	.01	05	16	.05	.32
Educ: Advanced degree					.00	-1.73	1.74	1.00
Educ: Associate's	51	-2.23	1.21	.56	.10	-1.85	2.05	.92
Educ: Bachelor's	34	-1.62	.94	.61	74	-2.21	.73	.33
Educ: Some associate's	50	-1.84	.84	.46	66	-2.14	.82	.38
Educ: Some adv. grad.	-1.37	-2.96	.22	.09	.01	-2.33	2.35	.99
Parenthood	.10	89	1.08	.84				
Raven's test					12	62	.38	.63
Wonderlic test					82	-1.33	32	.00

Study 1: N = 472. Adjusted $R^2 = .03$, F(10, 461) = 2.52, p = .006. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 476. Adjusted $R^2 = .06$, F(12, 463) = 3.47, p < .001. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes.

Table S7. Parameter estimates, 95% confidence intervals, and *p* values for models with the non-principal components analysis political summary measure as a predictor, weighted hazard credulity minus weighted benefit credulity as the outcome.

		Study	1			Study	2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	27	-2.09	1.54	.77	.63	-1.25	2.51	.51
Pol. summary (non-PCA)	.18	.04	.32	.01	.27	.14	.40	.00
Sex: Male	57	-1.53	.39	.24	.51	43	1.46	.29
Age	.00	05	.04	.98	.03	01	.07	.16
Ethnicity: White	06	-1.28	1.17	.93	33	-1.44	.79	.57
Income	.04	11	.18	.63	01	12	.10	.87
Educ: Advanced degree					49	-2.37	1.40	.61
Educ: Associate's	.38	-1.60	2.36	.71	.15	-1.97	2.27	.89
Educ: Bachelor's	.18	-1.29	1.65	.81	.38	-1.21	1.98	.64
Educ: Some associate's	1.23	31	2.77	.12	.28	-1.34	1.89	.74
Educ: Some adv. grad.	.98	85	2.82	.29	.13	-2.42	2.67	.92
Parenthood	.67	46	1.81	.24				
Raven's test					05	60	.49	.85
Wonderlic test					15	70	.40	.60

Study 1: N = 472. Adjusted $R^2 = .01$, F(10, 461) = 1.66, p = .09. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 476. Adjusted $R^2 = .03$, F(12, 463) = 2.08, p = .017. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes.

Table S8. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor, with the unweighted difference between a participant's hazard and benefit credulity as the outcome.

		Study	y 1			Stud	y 2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	20	49	.10	.19	15	46	.16	.34
Political summary	.05	.01	.10	.03	.06	.02	.11	.00
Sex: Male	07	23	.08	.34	.14	01	.30	.07
Age	.00	01	.01	.87	.01	.00	.01	.06
Ethnicity: White	13	33	.06	.19	14	32	.05	.14
Income	.01	01	.04	.30	.01	01	.03	.47
Educ: Advanced degree					07	38	.24	.66
Educ: Associate's	.07	25	.39	.67	01	36	.34	.97
Educ: Bachelor's	04	28	.20	.74	.08	18	.35	.54
Educ: Some associate's	.14	11	.38	.28	.03	24	.30	.82
Educ: Some adv. grad.	.09	20	.39	.54	09	51	.33	.68
Parenthood	.14	04	.32	.14				
Raven's test					.02	07	.11	.72
Wonderlic test					10	19	01	.03

Study 1: N = 472. Adjusted $R^2 = .01$, F(10, 461) = 1.66, p = .09. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 476. Adjusted $R^2 = .03$, F(12, 463) = 2.08, p = .017. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes.

Table S9a. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor, with a participant's credulity for each hazard item (i.e., not their mean credulity) as the outcome.

		Study	, 1			Study	2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	2.19	1.63	2.76	.00	2.11	1.50	2.72	.00
Political summary	.05	.01	.09	.01	.06	.03	.10	.00
Sex: Male	14	28	01	.04	.00	13	.13	1.00
Age	.00	01	.01	.72	.01	.00	.01	.00
Ethnicity: White	14	31	.03	.12	12	28	.04	.14
Income	01	03	.01	.28	.00	01	.02	.69
Educ: Advanced degree					09	36	.18	.52
Educ: Associate's	.06	22	.33	.67	02	31	.27	.89
Educ: Bachelor's	03	24	.17	.76	04	26	.19	.76
Educ: Some associate's	.09	13	.30	.43	15	38	.08	.21
Educ: Some adv. grad.	.09	17	.34	.49	03	40	.33	.87
Parenthood	.06	05	.26	.19				
Raven's test					01	09	.06	.75
Wonderlic test					13	20	05	.00
Gravity	.09	.06	.13	.00	.10	.06	.13	.00

Study 1: N = 472. Variance explained by random intercepts for participant: .20, and for question: .48; residual variance: 2.57. Log likelihood = -7294.42.

Study 2: N = 476. Variance explained by random intercepts for participant: .13, and for question: .59; residual variance: 2.66. Log likelihood = -7357.97.

Table S9b. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor, with a participant's credulity for each benefit item (i.e., not their mean credulity) as the outcome.

		Stud	y 1			Stud	y 2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	1.49	1.10	1.88	.00	1.39	1.04	1.75	.00
Political summary	.00	04	.04	.90	.02	02	.06	.26
Sex: Male	01	14	.12	.93	07	20	.06	.30
Age	.00	01	.00	.43	.00	.00	.01	.50
Ethnicity: White	.06	11	.23	.48	.05	10	.21	.50
Income	02	04	.00	.09	01	02	.01	.43
Educ: Advanced degree					.07	19	.33	.60
Educ: Associate's	.02	25	.29	.88	01	29	.28	.97
Educ: Bachelor's	.02	18	.22	.86	06	28	.16	.61
Educ: Some associate's	.00	20	.21	.97	15	37	.07	.19
Educ: Some adv. grad.	.09	15	.34	.46	.17	18	.53	.34
Parenthood	06	22	.09	.42				
Raven's test					03	.97	1.12	.41
Wonderlic test					.03	05	.10	.45
Gravity	.30	.27	.33	.00	.30	.27	.33	.00

Study 1: N = 472. Variance explained by random intercepts for participant: .19, and for question: .15; residual variance: 2.42. Log likelihood = -7177.94.

Study 2: N = 476. Variance explained by random intercepts for participant: .16, and for question: .09; residual variance: 2.25. Log likelihood = -7041.79.

Table S10. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor and parenthood status as a control, weighted hazard credulity minus weighted benefit credulity as the outcome, Study 2.

Variable	Parm. Est.	5% CI	95% CI	p
(Intercept)	.83	-1.16	2.83	.41
Political summary	.57	.29	.86	.00
Sex: Male	.42	60	1.43	.42
Age	.03	02	.07	.23
Ethnicity: White	72	-1.92	.49	.25
Income	.04	08	.17	.52
Educ: Advanced degree	60	-2.60	1.41	.56
Educ: Associate's	.03	-2.27	2.32	.98
Educ: Bachelor's	.33	-1.41	2.06	.71
Educ: Some associate's	.32	-1.41	2.06	.71
Educ: Some adv. grad.	43	-3.26	2.40	.77
Raven's test	28	86	.30	.34
Wonderlic test	08	67	.50	.78
Parenthood	.10	96	1.17	.85

 \overline{N} = 418. Adjusted R^2 = .03, F(13, 404) = 2.03, p = .017. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Table S11. Parameter estimates, 95% confidence intervals, and *p* values for models with political summary measure as a predictor, weighted hazard credulity minus weighted benefit credulity as the outcome, including all liberals excluded for Study 2 analyses.

Variable	Parm. Est.	5% CI	95% CI	р
(Intercept)	.68	98	2.34	.42
Political summary	.56	.33	.79	.00
Sex: Male	.47	35	1.28	.26
Age	.02	02	.05	.32
Ethnicity: White	46	-1.43	.52	.36
Income	02	11	.08	.77
Educ: Advanced degree	62	-2.29	1.04	.46
Educ: Associate's	.63	-1.20	2.45	.50
Educ: Bachelor's	.51	89	1.91	.48
Educ: Some associate's	.23	-1.18	1.65	.75
Educ: Some adv. grad.	.44	-1.74	2.62	.69
Raven's test	15	63	.33	.55
Wonderlic test	21	70	.27	.39

 \overline{N} = 607. Adjusted R^2 = .04, F(12, 594) = 3.09, p < .001. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19. Parenthood status excluded for Study 2 due to large number of incompletes.

Table S12. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor interacting with sex of the participant, weighted hazard credulity minus weighted benefit credulity as the outcome.

		Stud	y 1			Stud	v 2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	27	-2.09	1.55	.77	.61	-1.27	2.49	.53
Political summary	.41	.01	.81	.04	.31	08	.69	.12
Sex: Male	57	-1.53	.39	.24	.53	41	1.47	.27
Age	.00	05	.05	.99	.03	01	.07	.13
Ethnicity: White	04	-1.27	1.19	.94	31	-1.42	.81	.59
Income	.04	11	.18	.62	.00	11	.11	.99
Educ: Advanced degree					58	-2.46	1.30	.55
Educ: Associate's	.37	-1.60	2.35	.71	.00	-2.13	2.12	.99
Educ: Bachelor's	.16	-1.31	1.64	.83	.31	-1.29	1.90	.71
Educ: Some associate's	1.21	34	2.75	.13	.21	-1.40	1.82	.80
Educ: Some adv. grad.	.96	88	2.80	.31	03	-2.58	2.51	.98
Parenthood	.66	47	1.80	.25				
Raven's test					05	60	.49	.85
Wonderlic test					18	73	.37	.53
Political Summary * Sex	10	65	.46	.74	.43	08	.95	.10

Study 1: N = 472. Adjusted $R^2 = .01$, F(11, 460) = 1.52, p = .12. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 476. Adjusted $R^2 = .03$, F(13, 462) = 2.15, p = .011. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes.

Table S13. Parameter estimates, 95% confidence intervals, and p values for models with political summary measure as a predictor interacting with the age of the participant, weighted hazard credulity minus weighted benefit credulity as the outcome.

		Study	v 1			Stud	y 2	
Variable	Parm. Est.	5% CI	95% CI	p	Parm. Est.	5% CI	95% CI	p
(Intercept)	28	-2.10	1.53	.76	.63	-1.25	2.52	.51
Political summary	.15	36	.66	.57	.53	.07	1.00	.02
Sex: Male	.00	05	.04	.91	.03	01	.07	.18
Age	59	-1.55	.38	.23	.52	43	1.46	.29
Ethnicity: White	04	-1.26	1.19	.95	32	-1.44	.80	.57
Income	.03	11	.17	.69	01	12	.10	.87
Educ: Advanced degree					48	-2.37	1.40	.62
Educ: Associate's	.45	-1.53	2.43	.66	.14	-1.98	2.27	.90
Educ: Bachelor's	.20	-1.27	1.68	.79	.38	-1.22	1.98	.64
Educ: Some associate's	1.23	31	2.77	.12	.27	-1.34	1.89	.74
Educ: Some adv. grad.	1.07	77	2.91	.25	.12	-2.43	2.67	.93
Parenthood	.68	45	1.82	.24				
Raven's test					05	59	.49	.86
Wonderlic test					15	70	.41	.60
Political Summary * Age	.01	01	.03	.33	.00	02	.02	.97

Study 1: N = 472. Adjusted $R^2 = .01$, F(11, 460) = 1.59, p = .10. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 19.

Study 2: N = 476. Adjusted $R^2 = .03$, F(13, 462) = 1.93, p = .025. Women, "other" ethnicity, some high school/high school diploma, and non-parents are held at zero. Age is centered such that the intercept represents age 18. Parenthood status excluded for Study 2 due to large number of incompletes.

Figure S1. The estimated effect of political orientation on the difference between hazard and benefit credulity with the terrorism item excluded, across five seeds for imputation (with 95% confidence intervals) for (A) Study 1 and (B) Study 2. The effect of political orientation on credulity was robust across the exclusion of any of the other 15 items.

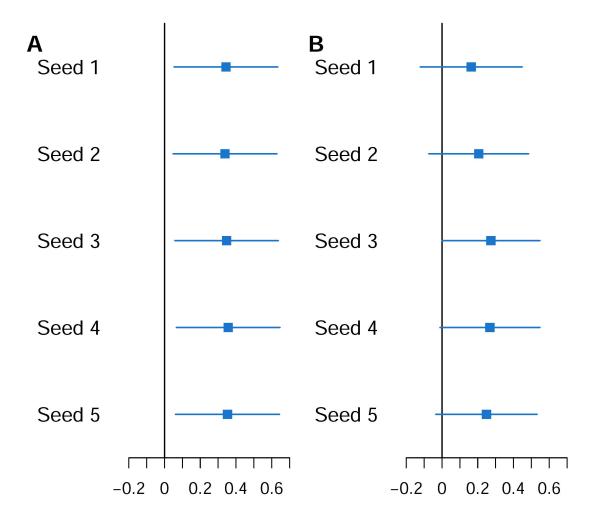
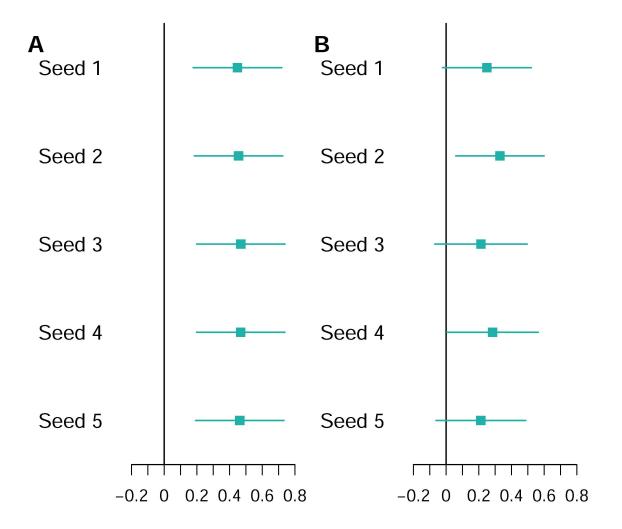
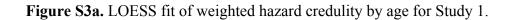
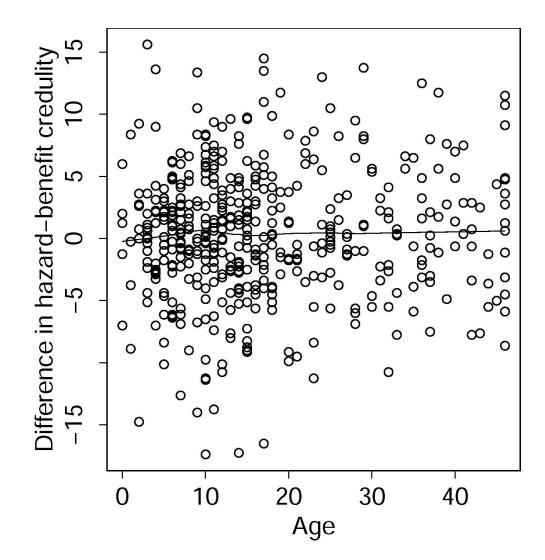
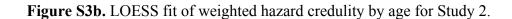


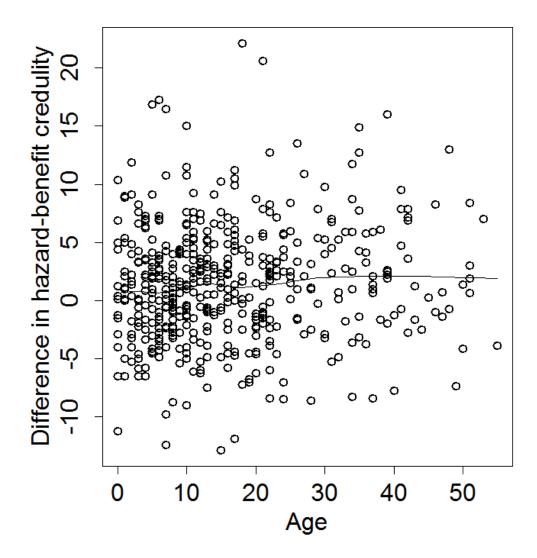
Figure S2. The estimated effect of political orientation on hazard credulity with the terrorism item excluded, across five seeds for imputation (with 95% confidence intervals) for (A) Study 1 and (B) Study 2. The effect of political orientation on credulity was robust across the exclusion of any of the other 15 items.











References

- Dodd, M. D., Balzer, A., Jacobs, C. M., Gruszczynski, M. W., Smith, K. B., & Hibbing, J. R. (2012). The political left rolls with the good and the political right confronts the bad: connecting physiology and cognition to preferences. *Philosophical Transactions of the Royal Society of London B: Biological Sciences*, 367(1589), 640-649. (doi: 10.1098/rstb.2011.0268)
- Fessler, D. M. T., Pisor, A. C., & Navarrete, C. D. (2014). Negatively-biased credulity and the cultural evolution of beliefs. *PLoS ONE*, *9*(4), e95167. (doi: 10.1371/journal.pone.0095167)
- Jones, J., & Saad, L. (2016). Gallup Poll Social Series: Values and Beliefs. Retrieved October 20, 2016 from http://www.gallup.com/poll/191741/democrats-liberal-social-issues-economic-ones.aspx?g_source=liberal&g_medium=search&g_campaign=tiles
- Little, R. J. A. (1988). Missing-data adjustments in large surveys. *Journal of Business & Economic Statistics* 6(3):287–296.
- Raven, J., Raven, J. C., & Court, J. H. (1998). *Manual for Raven's Progressive Matrices and Vocabulary Scales, Section 1: General Overview*. San Antonio, TX: Harcourt Assessment.
- Van Buuren S., and Groothuis-Oudshoorn K. (2011) MICE: Multivariate Imputation by Chained Equations. *Journal of Statistical Software 45*(3):1–67.
- Wonderlic (1992). Wonderlic Personnel Test: User's manual for the WPT and SLE. Liberty, IL: Wonderlic Personnel Test, Inc.