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Genome-wide association studies of alcohol dependence, DSM-IV criterion count, and individual criteria

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Abstract

Genome-wide association studies (GWAS) of alcohol dependence (AD) have reliably identified variation within alcohol metabolizing genes (e.g., *ADH1B*) but have inconsistently located other signals, which may be partially attributable to symptom heterogeneity underlying the disorder. We conducted GWASs of DSM-IV AD (primary analysis), DSM-IV AD criterion count (secondary analysis), and individual dependence criteria (tertiary analysis) among 7,418 (1,121 families) European American (EA) individuals from the Collaborative Study on the Genetics of Alcoholism (COGA). Trans-ancestral meta-analyses combined these results with data from 3,175 (585 families) African American (AA) individuals from COGA. In the EA GWAS, three loci were genome-wide significant: rs1229984 in *ADH1B* for AD criterion count ($p=4.16E-11$) and *Desire to cut drinking* ($p=1.21E-11$); rs188227250 (chromosome 8, *Drinking more than intended*, $p=6.72E-09$); rs1912461 (chromosome 15, *Time spent drinking*, $p=1.77E-08$). In the trans-ancestral meta-analysis, rs1229984 was associated with multiple phenotypes and two additional loci were genome-wide significant: rs61826952 (chromosome 1, DSM-IV AD, $p=8.42E-11$); rs7597960 (chromosome 2, *Time spent drinking*, $p=1.22E-08$). Associations with rs1229984 and rs18822750 were replicated in independent datasets. Polygenic risk scores derived from the EA GWAS of AD predicted AD in two EA datasets ($p<0.01$; 0.61-1.82% of variance). Identified novel variants (i.e., rs1912461, rs61826952) were associated with differential central evoked theta power (loss minus gain; $p=0.0037$) and reward-related ventral striatum reactivity ($p=0.008$), respectively. This study suggests that studying individual criteria may unveil new insights into the genetic etiology of AD liability.

Keywords

alcohol dependence; DSM-IV alcohol dependence criterion; DSM-IV criterion count; DSM-IV individual criteria; item response analysis; genome-wide association study; meta-analysis; polygenic risk score; Event-Related Theta Oscillations (ERO); functional Magnetic Resonance Imaging (fMRI)

INTRODUCTION

Alcohol dependence (AD), characterized by excessive drinking and diagnosed using features such as loss of control over drinking and excessive consumption despite negative consequences, is one of the most common and costly public health problems worldwide ¹. In the United States (U.S.), 12.5% of the population meets criteria for DSM-IV AD^{1,2}. AD is a complex disease with both genetic and environmental underpinnings and an estimated heritability around 50% ³. Identification of loci associated with AD liability could provide

new insights into the biological mechanisms underlying this serious disorder and lead to new therapeutic pathways.

Individual genome-wide association studies (GWAS) of AD have been relatively modest in size (but see a recent large publication using International Classification of Disease codes⁴) and have failed to identify consistently replicable loci⁵, with the exception of variants within the alcohol metabolizing genes, notably *ADH1B*, and to a lesser degree, *ADH1C*. A recent large GWAS meta-analysis of 14,904 AD cases and 37,944 controls, which includes some of the samples used in this study, also only detected genome-wide significant (GWS) association with rs1229984 (Europeans) and rs2066702 (African-Americans); both SNPs are in *ADH1B*⁶. However, when examining a broader definition of alcohol use disorders from medical records, loci in additional genes have recently been identified⁴. We have previously conducted GWAS of AD-related phenotypes in smaller subsets of the data used in the present study, but results have eluded replication and power to detect rs1229984 has been low (e.g., for AD in a subset of 1884 unrelateds⁷, for AD, criterion count and criteria in 2010-2,322 individuals from 118 families^{8,9}).

One possible challenge to identification of novel loci contributing to AD susceptibility may be the heterogeneity underlying the diagnosis of AD. Meeting criteria for DSM-IV AD requires that an individual endorse any three (or more) of the seven DSM-IV criteria (*Tolerance*; *Withdrawal*; *Drinking more than intended*; *Desire to cut drinking*; *Giving up activities*; *Time spent drinking*; *Drinking despite problems*) during the same 12-month period. However, psychometric literature points to the differential severity and contribution of individual criteria¹⁰. An approach to reduce diagnostic heterogeneity may be the analysis of individual DSM-IV criteria in addition to the overall AD diagnosis. Twin studies have suggested that the individual criteria that comprise the AD diagnosis are heritable¹¹⁻¹³. For instance, Kendler and colleagues showed the heritability of individual criteria ranged from 36% (*Desire to cut drinking*) to 59% (*Time spent drinking*)¹⁴. Another study found that heritability of individual criteria (in a subset of the data used here) were between 29% (*Tolerance*) and 59% (*Drinking more than intended*)⁹. Genomic data also support this variability with Palmer et al reporting a SNP-based heritability ranging from 13% (*Time spent drinking*) to 34% (*Tolerance*)¹⁵. The variability across these estimates likely arises from ascertainment (e.g., ascertained for addiction vs. twin epidemiologic sample) and the analytic approach (e.g., using SNPs vs. family relatedness). In addition, in one study, the observed associations with *ADH1B* loci were also differentially attributable to *Tolerance*, *Withdrawal*, *Drinking more than intended*, and *Time spent drinking*, relative to other criteria¹⁶.

Another strategy to improve the ability to detect variants contributing to DSM-IV AD is to consider the severity of the AD. One approach is to analyze a quantitative variable representing the total number of criteria that a person endorses. Although multiple combinations of criteria and study characteristics may result in a similar criterion count¹⁷, especially when fewer criteria are endorsed¹⁸, this proxy for AD severity has been successfully employed in previous studies^{19,20} as it makes no assumptions about the cut-off of three or more criteria as an index of “affection status” nor does it equate individuals with 1-2 criteria with those who endorse no criteria during their lifetime.

In this study, we sought to harness the phenotypic richness of the high density alcohol dependent families recruited as part of the Collaborative Study on the Genetics of Alcoholism (COGA) to perform a series of complementary analyses designed to identify variation contributing to the risk of AD. Our primary GWAS focused on DSM-IV AD diagnosis, a clinically validated measure of pathological drinking that is commonly used in GWAS⁶. We also conducted secondary GWAS of AD severity defined as the count of these criteria (range 0-7), as this quantitative phenotype has been shown to facilitate identification of GWS loci over the binary diagnostic measure of DSM-IV AD (e.g.,²¹). In tertiary analyses, we conducted exploratory GWASs of the seven individual DSM-IV AD criteria, in order to assess which criteria were the most significant contributors to the overall findings observed for DSM-IV AD diagnosis and criterion count, and further, examine whether novel loci emerged for individual criteria. To identify common variants associated with these phenotypes, a GWAS was performed in the European American (EA, n=1,114 families; “EA GWAS”) subsample of COGA, followed by a trans-ancestral genome-wide meta-analysis of the EA and African American (AA; N=585 families) subsamples. GWS ($p < 5 \times 10^{-8}$) findings were tested for replication in three independent datasets (Study of Addiction: Genetics and Environment (SAGE)²², Alcohol Dependence GWAS in European and African Americans (Yale-Penn)²¹, and the Australian Twin-family Study of Alcohol Use Disorder (OZALC)²³, which included EA (OZ-ALC, SAGE) and AA (SAGE, Yale-Penn) individuals. Polygenic risk scores (PRS) were created from the COGA EA GWAS and used to predict AD in EAs from SAGE and OZ-ALC. We also performed gene based analyses using COGA EA GWAS. Lastly, to probe the potential neural correlates of the GWS variants associated with aspects of AD, we tested whether GWS variants identified in the primary (DSM-IV AD), secondary (AD criterion count) or tertiary (individual criteria) analyses were associated with two reward-related neural phenotypes, one within a subset of young individuals from COGA²⁴ and another within the independent Duke Neurogenetics Study²⁵. The overall design of this study is shown in Figure 1.

MATERIALS AND METHODS

Collaborative Study on the Genetics of Alcoholism

Sample: COGA recruited AD probands from inpatient and outpatient AD treatment facilities in seven sites. Community-based families were also recruited from a variety of sources²⁶. Institutional review boards from all seven sites approved the study and all participants provided informed consent. COGA participants were administered the Semi-Structured Assessment for the Genetics of Alcoholism (SSAGA), a poly-diagnostic interview^{27,28}. Individuals below age 18 were administered the child version of the SSAGA, the C-SSAGA. If an individual was interviewed more than once, data from the interview with the maximum total number of endorsed DSM-IV AD criteria were utilized.

Measures: To avoid the inclusion of individuals with high genetic risk who do not drink for personal, social or cultural reasons, only individuals who reported ever drinking at least one full drink of alcohol in their lifetime were included in analyses (EA: N=7,418; AA: N=3,175).

The *primary phenotype* in this study was diagnosis of DSM-IV AD²⁹. Individuals meeting criteria for DSM-IV AD at age 15 or older were coded as affected. Individuals were coded as unaffected if they met all of the following criteria: 1) ≥ 21 years; 2) endorsed < 2 criteria for DSM-IV dependence or abuse for alcohol, and 3) endorsed < 2 criteria for DSM-IV dependence or abuse for cocaine, opioids, marijuana, sedatives, and stimulants. Affected individuals < 15 years of age and unaffected individuals < 21 years of age were excluded. Exclusions for age removed affected individuals with early onset AD who might be etiologically distinct, due to the potentially stronger role of environmental than genetic influences³⁰. For unaffected individuals, exclusion of those < 21 years of age removed those who may not have passed through the peak period of risk for the onset of AD^{31,32}. Due to the strong evidence for shared genetic influences on alcohol and other forms of substance use disorders, individuals who did not meet criteria for AD but endorsed multiple abuse or dependence criteria for other substances were also excluded from the analysis.

The *secondary phenotype* in this study was the sum of endorsed criteria out of the seven DSM-IV AD criteria.

Tertiary phenotypes included each of the seven individual DSM-IV AD criteria. Individuals who drank alcohol but did not endorse that specific criterion were coded as unaffected.

Phenotypic analysis: Tetrachoric correlations (for binary phenotypes) and polychoric correlations (for binary and count phenotypes) were calculated using SAS9.4 (SAS Institute Inc. Cary, NC, USA). We conducted an item response analysis in Mplusv8³³, using a two-parameter logistic model, to confirm the uni-dimensionality underlying the seven criteria and to examine the discrimination and difficulty associated with each criterion (see Supplemental Text).

Genotyping, Quality Review, Ancestry and Imputation: Four different genome-wide genotyping arrays were used in COGA: 1. COGA case/control data were genotyped on the Illumina Human1M array (Illumina, San Diego, CA, USA) at the Center for Inherited Disease Research (CIDR), Johns Hopkins University⁷; 2. COGA European American family data were genotyped on the Illumina Human OmniExpress 12V1 array (Illumina, San Diego, CA, USA) at the Genome Technology Access Center, Washington University School of Medicine^{9,34}; 3. COGA AA family data were genotyped on the Illumina 2.5M array (Illumina, San Diego, CA, USA) at CIDR³⁵; 4. The remaining samples were genotyped on the Smokescreen genotyping array (Biorealm LLC, Walnut, CA, USA) at Rutgers University. Among these arrays, two to 127 samples were genotyped on at least two different arrays with pairwise concordance rates all $> 99.18\%$.

A set of 47,000 variants genotyped on all arrays and meeting the following four criteria: common (defined as MAF $> 10\%$ in the combined sample), independent (defined as $R^2 < 0.5$), high quality (missing rate $< 2\%$ and Hardy-Weinberg Equilibrium (HWE) P-values > 0.001), were used to assess duplicate samples included on multiple arrays and also to confirm the reported pedigree structure. Family structures were altered as needed, and genotypes were checked for Mendelian inconsistencies using Pedcheck³⁶ with the revised family structure. Genotype inconsistencies were set to missing. The same set of 47,000

variants was also employed to calculate principal components (PCs) using Eigenstrat³⁷ and 1000 Genomes (Phase 3, version 5). Based on the first two PCs, each individual was then assigned a race classification (AA, EA, and Other). To maximize the value of the multiplex family recruitment strategy of COGA, family-based analyses were performed. Families were assigned a family-based race, according to the majority of individual-based race in that family.

All samples were imputed to 1000 Genomes using the cosmopolitan reference panel (Phase 3, version 5, NCBI GRCh37) using SHAPEIT2³⁸ then Minimac3³⁹ within each array. Only variants with non A/T or C/G alleles, missing rates < 5%, MAF > 3%, and HWE p values > 0.0001 were used for imputation. Imputed variants with $R^2 < 0.30$ were excluded, and genotype probabilities were converted to genotypes if probabilities ≥ 0.90 . Pedcheck³⁶ was used again to detect and clean Mendelian inconsistencies for imputed variants. All genotyped and imputed variants with missing rates < 25%, MAF $\geq 1\%$ and HWE p values > $1E-6$ were included in analyses. 8,021,023 and 6,832,792 genotyped and imputed variants passed QC and were included in COGA EA and trans-ancestral (EA+AA) meta-analysis respectively.

Genome-wide association studies and meta-analysis: Discovery GWAS were focused on the EA subsample and a trans-ancestral meta-analysis of GWAS summary statistics from the COGA AA and EA subsamples (EA+AA; see Figure 1). Even though a GWAS was conducted in the AA subsample, results were only used in the trans-ancestral meta-analysis. Due to the strict definition of AD controls, the individual AA subsample was too small for use as a discovery sample (both cases and controls had a sample size < 1000; full results available upon request). For binary traits, association analysis was performed using a generalized estimating equation (GEE) framework (with a binomial probability distribution) to control for relatedness with each family treated as a cluster. For the criterion count measure, a linear mixed effects model was fit to continuously distributed data with family relationship adjusted through a kinship matrix. The R package GWAF⁴⁰ was used to test both models. Birth cohort (birth year: 1890-1929; 1930-1949; 1950-1969; ≥ 1970) was a stronger predictor of alcohol dependence than was age (see also: Grucza et al., 2008⁴¹), and hence was selected along with sex, GWAS array indicator, and the first four ancestral principal components (as in a prior study by³⁴) as covariates in the model. In GWS regions, conditional analyses were performed by including the most significant variant in the region as a covariate to evaluate whether a single locus explained the association signal. The trans-ancestral (EA+AA) meta-analysis was performed using inverse-variance weighting in METAL⁴². As implemented in METAL, genomic control, which was estimated by comparing the median test statistics to those expected by chance alone, was applied to the GWAS of COGA AA and COGA EA. For the trans-ancestral meta-analysis (EA+AA), genomic control was applied to the standard errors of the effect sizes. All genomic control estimations were implemented in METAL. Only GWS variants ($p < 5E-8$) were evaluated in replication samples. As we tested seven individual criteria for the tertiary analyses, a matrix of the phenotypic correlations between these criteria in the EA participants (Supplemental Table 1B) was spectrally decomposed using matSpD^{43,44} resulting in 3 effectively

independent tests and thus a revised GWS p value threshold of $1.67E-8$ was used for the tertiary analyses.

Replication Samples

Three independent datasets from the database of Genotypes and Phenotypes (dbGaP) were used to replicate significant findings from primary, secondary and tertiary analyses: Study of Addiction: Genetics and Environment (non-overlapping individuals from SAGE, phs000092.v1.p1, https://www.ncbi.nlm.nih.gov/projects/gap/cgi-bin/study.cgi?study_id=phs000092.v1.p1), Alcohol Dependence GWAS in European and African Americans (Yale-Penn, phs000425.v1.p1, https://www.ncbi.nlm.nih.gov/projects/gap/cgi-bin/study.cgi?study_id=phs000425.v1.p1), and the Australian Twin-family Study of Alcohol Use Disorder (OZALC, phs000181.v1.p1, https://www.ncbi.nlm.nih.gov/projects/gap/cgi-bin/study.cgi?study_id=phs000181.v1.p1). Genotypic data from these samples were combined with genotypic data from the COGA samples to identify identical individuals across all datasets; overlapping subjects were retained in the discovery GWAS in COGA but excluded from the replication samples. Ancestry in the combined replication sample was determined in a manner similar to COGA. A similar definition of AD was employed where unaffected individuals with alcohol abuse, or other substance dependence were excluded. The secondary (DSM-IV AD criterion count) and tertiary (individual criteria) phenotypes were also coded in an identical manner. In each replication attempt, only the identical phenotype was tested in the replication cohort (e.g., for a variant that was GWS for one criterion but not others, only association with that criterion was tested in the replication samples). Due to the small sizes of the individual AA and EA subsets of the replication datasets, only the AA subsample of SAGE (SAGE-AA), EA subsample of SAGE (SAGE-EA), AA subsample of Yale-Penn (Yale-Penn-AA), and EA subsample of OZALC (OZALC-EA) were included as replication samples. Empirical kinships were estimated from genome-wide genotypic data using the “vcf2kinship” tool as implemented in RVTESTS, then mixed models adjusting for empirical kinships were fitted to the data using RVTESTS⁴⁵. For both SAGE-AA and SAGE-EA, sex and birth cohort (as defined in COGA) were used as covariates, while for OZALC-EA and Yale-Penn-AA, sex and age were used, as in publications of the parent studies. In addition, the first three PCs were included in all replication analyses.

Polygenic risk scores analyses

PRS analyses were performed using PRSice-2⁴⁶. EA summary statistics for the primary phenotype, DSM-IV AD, were used to score individuals in SAGE-EA and OZALC-EA datasets. Due to their well-known roles in AD, the alcohol dehydrogenase (ADH) gene cluster on chromosome 4 (99,985,095bp to 100,430,930bp) and *ALDH2* on chromosome 12 (112,196,532bp to 112,276,464bp) were excluded from PRS analyses to allow for estimation of polygenicity attributable to loci with smaller effects. A set of unrelated individuals was randomly selected from each replication sample (SAGE-EA: N=1,373; OZALC-EA: N=1,441) as required by PRSice-2. Variants located within 500kb of the index variant and having $r^2 \geq 0.25$ with the index variant were clumped. PRS were derived by multiplying effect sizes from the EA GWAS of the primary phenotype, DSM-IV AD, with the number of effect alleles in each individual in the target dataset. These product terms were then averaged

across the total number of included variants. We only used the p-value threshold of $p = 0.05$ (i.e., SNPs associated with DSM-IV AD in the discovery EA GWAS at $p = 0.05$) in order to reduce the burden of multiple testing and included the same covariates as those used in replication analyses in each dataset.

Gene based analysis

MAGMA (De Leeuw et al., 2015), which is implemented in FUMA, a web based functional mapping and annotation tool (Watanabe et al., 2017) was used to perform gene based analysis. LD was estimated using the European samples from 1000 Genomes projects.

Neural extension I: Event-Related Theta Oscillations (ERO) analysis of GWS loci in COGA Prospective Sample

The COGA Prospective Sample includes offspring aged 12-34 years from COGA families, and was designed to assess multiple domains (e.g., clinical, neurophysiological), at 2-year intervals,²⁴. Neurophysiological analyses of reward-related theta ERO data from the most recent assessments were carried out in a subsample of 825 COGA AA (49.9% male, 22.12 ± 5.21 years of age) and 1,726 COGA EA (48.8% male, 22.26 ± 5.21 years of age) young adults. Further details are in Supplemental Text.

A monetary gambling task was implemented as detailed elsewhere⁴⁷. Briefly, individuals bet 50¢ or 10¢ in each of 172 trials, with one of four possible outcomes: lose 50¢, lose 10¢, gain 50¢, or gain 10¢, with equal number of loss and gain trials (Supplemental Figure 1). Evoked theta ERO power (3.5–7.5 Hz) during monetary loss and gain feedback were measured and differential reward processing ('loss – gain') was derived at frontal, central, and parietal regions (Supplemental Figure 2). Linear regression was applied to test the associations between the top variants and theta ERO power after adjusting for sex, age, and first three PCs. We did not examine rs1229984 in *ADH1B* in either the COGA Prospective Sample or the Duke Neurogenetics Study (below) due to its well-known role in the alcohol metabolizing process. For the remaining four GWS loci (rs61826952 and rs7597960 from EA+AA meta-analysis, as well as rs188227250 and rs1912461 from the EA GWAS), three brain regions were tested; therefore, after multiple testing correction, the significance threshold was $p = 0.0042$ (i.e., 12 tests). Further details on data acquisition and processing are in Supplemental Text.

Neural extension II: Reward-related functional magnetic resonance imaging analyses of GWS loci in the Duke Neurogenetics Study

We examined whether GWS loci identified in analyses of alcohol-related phenotypes were associated with reward-related brain function among non-Hispanic AA ($n=118$; 72% female, 19.6 ± 1.2 years of age) and EA ($n=481$; 54.5% female, 19.8 ± 1.2 years of age) undergraduate students who completed the Duke Neurogenetics Study (DNS;²⁵; see Supplemental Text). For rs7597960, which was unavailable in DNS imputed data, we used a proxy SNP, rs2418646, which is in complete LD (i.e., $r^2=1.0$, $D'=1.0$) within those of African and European ancestries. The chromosome 8 and 15 loci were unavailable in DNS imputed data and no proxies were available; due to their low MAFs, they were difficult to impute in this smaller sample. A number guessing paradigm was used to elicit ventral

striatum (VS) reactivity associated with positive and negative feedback linked to monetary gains and losses while blood-oxygen-level dependence (BOLD) functional magnetic resonance imaging (fMRI) data were acquired⁴⁸. Statistical Parametric Mapping version 8 (SPM8) software was used to extract parameter estimates for the contrast of Positive Feedback > Negative Feedback from maximal voxels within left and right VS regions of interest (ROIs). Imaging acquisition protocol, task, ROIs, and preprocessing details are described in the Supplemental Text. Extracted parameter estimates from VS activity in each hemisphere were regressed on genotype (rs61826952 coded as 1 or more copies of the minor allele due to sample size; rs2418646 coded using an additive model for the number of C alleles) while co-varying for sex, and three (AA) or two (EA) ancestral principal components using Full Information Maximum Likelihood in MPlus v7.3⁴⁹. Trans-ancestral meta-analysis was conducted using METAL⁴². To adjust for multiple comparisons, we used a Bonferroni-corrected p-value threshold ($p < 0.0125$), to account for our hypothesized 4 tests (i.e., rs61826952 and rs2418646 in both brain hemispheres in a trans-ancestral meta-analysis).

RESULTS

Phenotypic analyses:

Tables 1 (primary and secondary phenotypes of DSM-IV AD and criterion count) and 2 (tertiary analysis of seven individual criteria) summarize the samples used in discovery and replication analyses. There were 7,418 (1,114 families) EA and 3,175 (585 families) AA individuals, respectively. In total, there were 18,586 individuals evaluated for DSM-IV AD in both discovery and replication samples, with 7,482 AD cases and 6,169 controls. As shown in Supplemental Table 1, the primary, secondary and tertiary phenotypes were highly correlated with each other in both EAs and AAs, with DSM-IV AD and DSM-IV AD criterion count having the highest correlations with each individual criterion in both AA and EA subsamples ($r > 0.87$). As shown in Supplemental Table 2, the item response analysis demonstrated that all criteria loaded well on a single underlying AD factor. Some criteria discriminated liability at the lower end of the liability distribution (e.g., *Drinking more than intended*) while others (e.g., *Withdrawal*, *Time spent drinking*, *Giving up activities*) contributed at the higher end of the severity continuum (Supplemental Text).

GWAS findings:

Regions on chromosomes 1, 2, 4, 8 and 15 reached GWS ($p \leq 5E-8$) for primary, secondary and tertiary phenotypes in EA and EA+AA GWAS, respectively (Table 3; Manhattan, quantile-quantile and regional association plots for GWS findings are in Supplemental Figures 3 and 4 respectively; effect sizes, standard errors and p-values for EA and AA subsamples and the EA+AA analysis in Supplemental Table 3). All genomic controls (λ) are listed in Supplemental Table 4.

Primary phenotype (DSM-IV AD diagnosis): In EA, no GWS findings were identified. In the trans-ancestral meta-analysis (EA+AA), consistent with prior GWAS, rs1229984 in *ADH1B* was significantly associated with AD ($p = 1.72E-8$). In addition, a novel GWS locus was also identified on chromosome 1 (rs61826952, $p = 8.42E-11$) in the EA+AA analysis.

Both the EA ($p=7.73E-6$) and AA ($p=1.50E-07$; results available upon request) subsamples contributed to the finding, with the same direction of effect. Conditional analyses confirmed that there were independent associations in the *ADH1B* region but not in the chromosome 1 region (Supplemental Figure 5A, 5C).

Secondary phenotype (DSM-IV AD criterion count): Rs1229984 in *ADH1B* was associated at GWS levels in the EA and the EA+AA analysis.

Tertiary phenotypes (individual criteria): In EA, rs1229984 was associated with *Desire to cut drinking* ($p=1.21E-11$). Two novel regions were GWS for two individual DSM4 criteria: rs188227250 on chromosome 8 for *Drinking more than intended* ($p=6.72E-09$); rs1912461 on chromosome 15 for *Time spent drinking* ($p=1.77E-08$). For the trans-ancestral (EA+AA) analysis, rs1229984 was significantly associated with *Desire to cut drinking* ($p=6.01E-14$) and *Tolerance* ($p=8.06E-9$). An additional GWS region on chromosome 2 (rs7597960, $p=1.22E-8$) was noted for *Time spent drinking*. The regions on chromosome 2, 4 and 8 survived the more stringent correction for the seven criteria ($p=1.67E-8$) while the chromosome 15 variant was GWS but did not survive the additional correction for multiple testing of individual criteria (i.e., $p=1.77E-8$). Conditional analyses demonstrated that there was only one association signal in the chromosome 15 region; however, the possibility of a second independent signal in the chromosome 8 region could not be ruled out ($p<0.001$) (Supplemental Figures 5D and 5E). Conditional analyses also suggested independent associations in the chromosome 2 region (Supplemental Figures 5B).

Replication:

Rs1229984 in *ADH1B* was replicated in OZALC-EA for the primary AD phenotype (Table 3); in SAGE-AA for the secondary DSM-IV AD criterion count as well as for tertiary phenotypes of *Desire to cut drinking* in SAGE-AA, SAGE-EA, and OZALC-EA, and in SAGE-AA, for *Tolerance*. Meta-analysis of all available datasets enhanced significance across primary and tertiary phenotypes (Table 3). The association between rs188227250 and *Drinking more than intended* was replicated in OZALC-EA and a meta-analysis of EA, SAGE-EA, and OZALC-EA strengthened the association ($p=3.71E-09$, Table 3). Although rs1912461 on chromosome 15 was not significantly associated with *Time spent drinking* in either the SAGE-EA or OZALC-EA samples ($p>0.12$), the direction of the effect was the same and meta-analysis across COGA and the replication samples retained significance for this variant ($p=2.31E-08$, Table 3). Variants on chromosomes 1 and 2 did not replicate in any dataset (all $p>0.07$ or opposite direction of effects; Table 3).

Polygenic risk score analyses:

PRS derived using the EA discovery GWAS of the primary phenotype (i.e., DSM-IV AD) predicted 1.82% and 0.61% of the variance in AD in SAGE-EA ($p=1.32E-05$) and OZALC-EA ($p=7.73E-03$), respectively.

Gene based analyses:

Supplemental Table 5 lists the results of gene based analyses. Two genes, *OTOP1* ($P=8.73E-7$) for DSM-IV criterion count, and *BRINP1* ($P=7.85E-8$) for *Drinking despite problem*, were genome-wide significant.

Neural extension I: COGA Prospective Sample: Theta ERO

Rs1912461 on chromosome 15 for *Time spent drinking* was significantly associated with differential evoked theta power (loss-gain) in the Central ($F_{1,1370}=8.4346$; $p=0.0037$) region (Supplemental Table 6). The minor allele carriers of rs1912461 manifested higher differentiation of gambling outcomes (loss-gain) at the anterior region of the brain (Supplemental Figure 6). Other variants did not survive the multiple testing correction.

Neural Extension II: Duke Neurogenetics Study: fMRI

Carriers of the minor (G) allele of rs61826952 had lower left, but not right, reward-related (positive feedback – negative feedback) VS activity when compared to non-carrier individuals in the combined and AA and EA samples (*Left*: trans-ancestral meta-analysis: $\beta=-0.041$, $p=0.008$; AA: $\beta=-0.124$, $p=0.018$; EA: $\beta=-0.033$, $p=0.041$; *Right*: trans-ancestral meta-analysis: $\beta=-0.01$, $p=0.570$). Reward-related VS activity was not significantly associated with rs2418646 genotype (*Left*: trans-ancestral meta-analysis: $\beta=-0.007$, $p=0.560$; *Right*: trans-ancestral meta-analysis: $\beta=0.0003$, $p=0.97$).

DISCUSSION

This large, family study of AA and EA individuals utilized a multi-pronged approach (Figure 1) to dissect the genetic underpinnings of alcohol dependence (DSM-IV AD). In addition to the primary phenotype of DSM-IV diagnosis of AD, and severity as captured by the AD criterion count, it is, to our knowledge, the largest GWAS of each DSM-IV AD criterion. We detected five regions with variants meeting traditional GWS criteria, of which four were novel (chromosomes 1, 2, 8, and 15). Notably, the chromosome 8 signal was replicated in an independent dataset, as was the well-known association with rs1229984 in *ADH1B*. Even when excluding the larger effect size associated with rs1229984, PRS derived from the EA GWAS predicted 0.61-1.82% of the variation in AD in independent datasets, underscoring significant polygenicity underlying liability to the disorder. Analyses of two reward-related neural phenotypes also showed associations with two GWS variants.

Consistent with several prior studies⁶, rs1229984 in *ADH1B* was associated with DSM-IV AD. Although GWS was only noted in the trans-ancestral (EA+AA) analysis, as shown in Supplemental Table 7, rs1229984 was associated with the AD criterion count and criteria indexing physiological dependence and *Desire to cut drinking* at GWS levels, and with other AD criteria at nominal levels of significance. Despite the robust relationship between this functional variant and AD, its relatively low minor allele frequency necessitates fairly large samples to detect a GWS effect for a binary trait, as was shown in a recent meta-analysis of DSM-IV AD⁶. However, for DSM-IV AD criterion count, rs1229984 was GWS in both the EA and EA+AA analyses. Similar to another study¹⁶, we found that while rs1229984 was associated with each individual criterion (EA all $p<3.61E-04$; EA+AA all $p<4.54E-05$), the

association was stronger with certain DSM-IV AD criteria. Consistent with Hart et al., *Tolerance* was strongly associated with rs1229984 ($p=8.06E-09$ in EA+AA). However, the additional GWS associations with *Desire to cut drinking* in our study differs from the prior study which used a sequential regression approach to identify *Withdrawal* and *Drinking more than intended* as additional criteria related to rs1229984 in EA, and *Time spent drinking* in AA. However, another study of 1,130 individuals of Jewish descent reported associations between rs1229984 and both *Tolerance* and *Desire to cut drinking*⁵⁰. Across these studies, the most robust association signal for rs1229984 appears to arise from *Tolerance*, which is notably an index of excessive consumption and consistent with the role of *ADH1B* in other studies of non-problem alcohol intake⁵¹. Plausibly, the strong findings with *Desire to cut drinking* might also support this as epidemiological studies have shown this criterion to index liability to less severe AD (Supplemental Table 2; Supplemental Figure 7), and therefore, serve as a marker of excessive drinking, rather than severe pathology and impairment^{10,52-54}. Differences in associations with other criteria could stem from the relative severity of individual criteria in each dataset or their relationship with excessive drinking.

The GWS findings for the other loci are novel and have not been previously reported for AD or related phenotypes, although these regions have been linked to some neuropsychiatric diseases/traits. The region on chromosome 1 was previously linked to cerebrospinal fluid biomarker level⁵⁵, migraine⁵⁶, illegal substance dependence⁵⁷, and neuroticism⁵⁸. This region encompasses gene *RABGAP1L*, with many other genes nearby (Supplemental Figure 4A). *RABGAP1L* is broadly expressed in brain regions and showed association with cerebrospinal fluid biomarker levels⁵⁵, and migraine⁵⁶. Other genes near this region seem interesting too, e.g. *KIAA0040*, which is downstream of this region, was associated with alcohol dependence⁵⁹. The chromosome 2 region is in a gene desert (Supplemental Figure 4B) and has been linked to cognitive test scores⁶⁰, ADHD symptom count⁶¹, ADHD⁶², current smoking⁶³, and juvenile myoclonic epilepsy⁶⁴. The region on chromosome 8 has been linked to bipolar disorder⁶⁵. The only gene near the chromosome 8 region is *FAM84B* (Supplemental Figure 4D), however, this gene doesn't seem to be related any neuropsychiatric diseases. The chromosome 15 region harbors some non-coding RNAs (Supplemental Figure 4E) and was previously linked to the rate of cognitive decline⁵⁵, ADHD⁶⁶, and major depression⁶⁷. Thus, despite our discovery of novel loci, much further study is needed to investigate the role of these variants in the etiology of alcohol dependence and related traits.

In our data, the chromosome 1 variant showed nominal association with multiple AD criteria and the criterion count, but none at GWS levels. However, a highly correlated variant (rs1890881) was associated at GWS with a phenotype representing dependence on alcohol or illicit drugs (cannabis, cocaine, sedatives, stimulants, opioids) in the same sample (see accompanying paper by Wetherill et al). It is possible that this variant is associated with overall liability to AD and dependence on other drugs but to a lesser extent with AD severity as indexed by a single continuous criterion count. Research has noted that mere summation does not capture the heterogeneity underlying AD severity, where constellations of criteria could result in meaningful individual differences¹⁰. Prior latent class analyses aimed to parse out such groups of individuals with unique sets of criteria including in a subset of

these data ⁹. However, assessment of the genomic underpinnings of such heterogeneous groups of individuals would require extremely large sample sizes. The chromosome 8 variant, rs188227250, was uniquely associated with *Drinking more than intended* (Supplemental Table 7). In epidemiological studies and in COGA (Supplemental Table 2), this criterion is endorsed quite frequently by individuals with AD, and also by those who do not meet criteria for DSM-IV AD and thus, might index lower severity. Indeed, in IRT analyses, this criterion had the lowest difficulty as indicated by the item characteristic curves in Supplemental Figure 7. In contrast, the finding on chromosomes 2 and 15, while GWS for *Time spent drinking* were also associated with *Giving up activities* (at nearly GWS for chromosome 15), both highly correlated criteria indicative of high difficulty, and thus, risk for DSM-IV AD ⁹. In addition to *Withdrawal*, we previously found these criteria to distinguish a highly heritable high-risk group of individuals at risk for AD from those in both low and moderate-risk groups. Thus, as shown in Supplemental Figure 7, while the chromosome 8 finding potential maps to lower AD severity, the chromosome 2 and 15 findings potentially indicate greater severity. However none of these loci were GWS for our AD criterion count measure, which is commonly used as an index of severity. These results are consistent with the argument that the validity of an individual criterion, and its impact on impairment may rely heavily on the other criteria that are endorsed alongside it ¹⁰. Importantly, these results underscore that novel information can be gained from studying individual criteria that index differing levels of AD severity that may operate discontinuously.

Gene based analysis identified two genome-wide significant genes for two different phenotypes. *OTOP1* was associated with DSM-IV criterion count. This gene is related to maintaining metabolic homeostasis but it is not well-studied. *BRINP1* showed association with *Drinking despite problems*. This gene is mostly expressed in brain regions and has been linked to schizophrenia^{68,69}; cognition disorders⁵⁵, and Parkinson's disease⁷⁰. Further studies are needed to test its role in AD.

Previous studies indicate that AD may be related to variations in the brain's reward system ⁷¹, including decreased reward-network volume ⁷² and differential neural activity in reward circuitry ⁷³⁻⁷⁵. In the COGA Prospective Sample, minor allele carriers of rs1912461 showed greater differentiation in frontal evoked theta power between loss and gain feedback trials in an EEG-based Monetary Gambling Task. Prior studies have found lower reward-related theta power in alcoholics and in high-risk offspring of alcoholics than controls performing the same task ^{47,76}. Frontal theta response underlies a variety of cognitive processes ^{77,78} including reward processing ⁷⁹⁻⁸¹. Moreover, it has recently been proposed that frontal theta reflects a promising mechanism through which cognitive control may be enacted by invoking a shift from habitual-based striatum responses to deliberative prefrontal-based control of behavior ⁸². Furthermore, the frontal-central theta power difference between loss and gain conditions may reflect the need for cognitive control to process goal-relevant information, such as decision making and action selection, based on choice-relevant information (approach-avoidance, reward-punishment, success-failure, etc.) for optimal functioning in the environment ⁸². In this study, the COGA Prospective participants were included in the COGA discovery GWAS. We, therefore, examined the sensitivity of our discovery findings to exclusion of these overlapping individuals from the Prospective

sample. The resulting GWAS found that while statistical significance decreased in some instances due to the decrease in sample size, the overall results remained highly consistent (e.g., for the EA-only finding of *Drinking more than intended*, the p-value decreased from 6.72E-09 to 3.61E-08; data not shown), indicating that the overlapping subjects were not solely responsible for the GWS findings from the discovery GWAS.

In the Duke Neurogenetics Study, rs61826952 minor allele carriers had decreased VS activity to positive versus negative feedback in a number-guessing fMRI task. Increased VS activity and dopamine release to non-alcohol reward have been associated with substance use initiation and problematic drinking^{25,83-85}. In contrast, studies of AD reported relatively reduced VS activity to non-alcohol reward^{86,87} and heightened activity to alcohol cues⁸⁸. These apparently disparate findings can be integrated with stage-based theories of addiction, which hypothesize that initial problematic use is associated with the positively reinforcing aspects of a substance, while later compulsive use is driven by negative reinforcement and diminished cognitive control, resulting from changes in neural plasticity induced by chronic alcohol use⁸⁹ (see also Wetherill et al accompanying paper). Thus, results from the college-based Duke Neurogenetics Study suggested that the minor allele of rs61826952 may protect from AD by reducing VS-related reward drive, thereby diminishing the likelihood of initiating problematic drinking behavior.

Replication of individual variants/genes other than those involved in alcohol metabolism can be challenging and notably influenced by heterogeneity across samples, ascertainment approach, definitions of affected and unaffected, and even nuanced differences in interview instruments¹⁷. For instance, although families ascertained for AD were included in the replication samples, OZALC had samples ascertained for heavy smoking and drinking (as well as sibships ascertained merely for large pedigree size), and SAGE included two subsamples recruited for nicotine and cocaine dependence. In addition, unlike the prior large AD GWAS by Gelernter and colleagues¹⁹, we excluded individuals with 2 abuse or dependence criteria for alcohol or any illicit drug from our unaffected group¹⁹. This may have led to a greater degree of genetic separation between affecteds and unaffecteds in the current analysis and contributed to the lack of replication. Despite these potential differences, for 2 of the 5 loci (rs1229984 and rs188227250), meta-analyses across samples yielded more significant associations. In addition, the PRS analyses found that the aggregated effect of variants in regions other than the ADH cluster and the *ALDH2* locus significantly contributed to AD liability in these diversely ascertained samples. While the proportion of explained variance is modest, it is consistent with other PRS analyses⁹⁰ and supports the generalizability of our findings at a polygenic level.

We also examined whether our analyses supported recent findings from Kranzler et al., who conducted a GWAS of alcohol use disorders defined using International Classification of Disease (ICD) codes derived from the electronic health records of individuals participating in the Million Veterans Project⁴. In this multi-ancestral sample of 274,424 predominantly male veterans, Kranzler et al identified 18 genome wide significant loci for AUD as well as for the consumption subscale of the Alcohol Use Disorders Identification Testkit (AUDIT-C). Their signal for rs1229984 was also noted in our COGA GWAS. In addition, modest evidence for directional and statistical support was also noted for rs12639940 on

chromosome 4 ($p=0.03$; COGA-EA), and rs2961816 on chromosome 5 ($p=0.04$; COGA EA +AA).

Our findings should be considered within the context of a few key limitations. First, despite being large, it is evident that our sample is underpowered to detect loci of modest effect. However, our sample was considerably larger than in our prior efforts in a subset of these data (e.g., ⁷⁻⁹) and one GWS SNP from those prior studies, previously linked to a latent class representing high-risk for AD⁹, continued to be nominally associated with DSM-IV AD in the current analysis (rs17484734, prior $p=4.1E-8$, current $p=8.77E-5$) but two other borderline significant variants were not as strongly associated in the current larger sample (rs11035102, for *Desire to cut back*⁹: prior $p=7.3E-8$, current $p=0.002$; rs12903120, for AD criterion count⁸: prior $p=5.45E-8$, current $p=0.03$). Second, some of our GWS loci had low minor allele frequencies which may also have limited replication efforts. Third, our AA subsample, while utilized in the EA+AA analysis, was too small to report on individually, due to the strict definition of AD affecteds. Larger discovery GWAS of non-EA samples is much needed.

In summary, our study highlights the importance of utilizing a variety of phenotypes, including individual dependence criteria in locus discovery for AD. The heterogeneity that underlies the diagnosis of AD due to the various combinations of individual criteria that can be endorsed to meet diagnostic criteria, is also true for major depression disorder (MDD), and has been shown to hinder GWAS⁹¹. While significant increases in sample size can potentially overcome this heterogeneity (as has been shown in the GWAS of MDD⁹²), the study of individual criteria, alongside diagnosis and severity, can provide a more detailed characterization of common and specific genetic influences on aspects of AD, especially when viewing individual criteria as psychometric indices of various cut-points of AD liability, and may eventually shape individualized treatment based on criterion profiles and other related features, over and above a mere diagnosis of AD.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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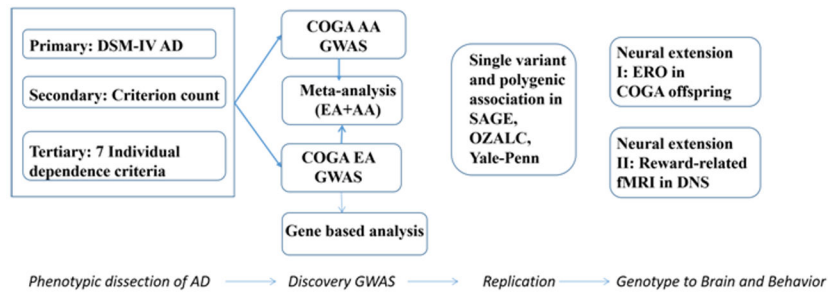


Figure 1:
flow chart of analyses.

Table 1:

Summary of characteristics of COGA and replication datasets.

Sample		AA			EA		
		# AD case (%Male)	# AD control (%Male)	# Individuals with DSM-IV criterion (%Male)	# AD case (%Male)	# AD control (%Male)	# Individuals with DSM-IV criterion (%Male)
Discovery	COGA	880 (61.70)	951 (25.45)	3,175 (46.58)	2,411 (62.01)	2,438 (28.47)	7,418 (47.53)
Replication	Yale-Penn	1,524 (60.50)	485 (29.69)	2,010 (53.08)	-	-	-
	SAGE	387 (59.17)	330 (39.09)	930 (46.24)	630 (52.70)	758 (34.17)	1,708 (38.82)
	OZALC	-	-	-	1,650 (62.24)	1,206 (46.10)	3,345 (53.69)
Total		2,791	1,767	6,115	4,691	4,402	12,471

Table 2:

Summary of samples with individual DSM-IV AD criteria in all datasets.

DSM-IV AD criterion number	Criterion description	Sample	AA		EA	
			# Case	# Control	# Case	# Control
1	<i>Tolerance</i>	COGA	1,110	2,024	3,348	3,958
		Yale-Penn	1,192	818	-	-
		SAGE	353	577	777	930
		OZALC	-	-	2,274	1,071
2	<i>Withdrawal</i>	COGA	514	2,616	1,259	6,046
		Yale-Penn	694	1,316	-	-
		SAGE	200	730	257	1,451
		OZALC	-	-	478	2,867
3	<i>Drinking more than intended</i>	COGA	1,317	1,817	3,826	3,480
		Yale-Penn	1,525	485	-	-
		SAGE	507	421	1,074	631
		OZALC	-	-	2,055	1,290
4	<i>Desire to cut drinking</i>	COGA	1,436	1,701	2,896	4,413
		Yale-Penn	1,411	599	-	-
		SAGE	425	505	601	1,107
		OZALC	-	-	1,420	1,925
5	<i>Giving up activities</i>	COGA	578	2,558	1,437	5,871
		Yale-Penn	1,201	809	-	-
		SAGE	215	715	274	1,434
		OZALC	-	-	246	3,099
6	<i>Time spent drinking</i>	COGA	546	2,590	1,533	5,776
		Yale-Penn	1,004	1,006	-	-
		SAGE	251	679	354	1,354
		OZALC	-	-	668	2,677
7	<i>Drinking despite problems</i>	COGA	784	2,351	2,163	5,144
		Yale-Penn	989	1,021	-	-
		SAGE	310	619	741	966
		OZALC	-	-	1,180	2,165

1: *Tolerance*. Need for markedly increased amounts of alcohol to achieve intoxication or desired effect; or markedly diminished effect with continued use of the same amount of alcohol.

2: *Withdrawal*. The characteristic withdrawal syndrome for alcohol; or drinking (or using a closely related substance) to relieve or avoid withdrawal symptoms.

3: *Drinking more than intended*. Drinking in larger amounts or over a longer period than intended.

4: *Desire to cut drinking*. Persistent desire or one or more unsuccessful efforts to cut down or control drinking.

5: *Giving up activities*. Important social, occupational, or recreational activities given up or reduced because of drinking.

6: *Time spent drinking*. A great deal of time spent in activities necessary to obtain, to use, or to recover from the effects of drinking.

7: *Drinking despite problems*. Continued drinking despite knowledge of having a persistent or recurrent physical or psychological problem that is likely to be caused or exacerbated by drinking.

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Table 3:

Top variants in each region that met genome-wide significance.

Phenotype	CHR	BP	variant	Gene	Minor Allele	Major Allele	AA MAF	EA MAF	EA P value	COGA AA+EA P value	Yale-Penn-AA P value	SAGE-AA P value	SAGE-EA P value	OZALC EA P value	Meta P value	Direction
DSM -IV AD	1	174,637,937	rs61826952	<i>RABGAP1L</i>	G	A	0.10	0.07	7.73E-06	8.42E-11	0.01	0.13	0.46	0.24	1.66E-04	---++
<i>Time spent drinking</i>	2	123,424,651	rs7597960	<i>TSN.LINC01826</i>	T	A	0.65	0.21	4.43E-06	1.22E-08	0.23	8.94E-03	0.83	0.33	8.86E-05	---++
DSM -IV AD	4	100,239,319	rs1229984	<i>ADH1B</i>	T	C	0.01	0.03	2.09E-07	1.72E-08	0.82	0.11	0.54	0.01	5.30E-09	---++
DSM -IV AD criterion count	4	100,239,319	rs1229984	<i>ADH1B</i>	T	C	0.01	0.03	4.16E-11	2.61E-13	0.86	0.01	0.10	0.06	4.09E-10	---++
<i>Desire to cut drinking</i>	4	100,239,319	rs1229984	<i>ADH1B</i>	T	C	0.01	0.03	1.21E-11	6.01E-14	0.82	0.03	4.89 E-03	5.52E-03	9.79E-18	---++
<i>Tolerance</i>	4	100,239,319	rs1229984	<i>ADH1B</i>	T	C	0.01	0.03	1.89E-07	8.06E-09	0.95	7.31E-03	0.32	0.14	1.20E-09	---++
<i>Drinking more than intended</i>	8	127,213,398	rs188227250	<i>LINC00861, LOC101927657</i>	A	G	0.02	0.03	6.72E-09	1.03E-06	-	-	0.91	1.96E-03	3.71E-09	++
<i>Time spent drinking</i>	15	36,344,539	rs1912461	<i>MIR4510, C15orf41</i>	C	T	0.02	0.01	1.77E-08	1.60E-05	-	-	0.18	0.12	2.31E-08	+++

Minor allele is defined as minor allele in EA samples. Directions of effect are in terms of minor alleles with + and - meaning increasing or decreasing the chance of having AD or DSM-IV AD criterion, or the number of DSM-IV AD criterion count. For significant findings in COGA AA+EA, the order of directions are: COGA AA, EA, Yale-Penn-AA, SAGE-AA, SAGE-EA; for significant findings in EA, the order of directions are: EA, SAGE-EA, OZALC-EA.