

# UC San Diego

## UC San Diego Previously Published Works

### Title

A Physiological Time Series Dynamics-Based Approach to Patient Monitoring and Outcome Prediction

### Permalink

<https://escholarship.org/uc/item/4fq6r1jd>

### Journal

IEEE Journal of Biomedical and Health Informatics, 19(3)

### ISSN

2168-2194

### Authors

Lehman, Li-wei H  
Adams, Ryan P  
Mayaud, Louis  
[et al.](#)

### Publication Date

2015-05-01

### DOI

10.1109/jbhi.2014.2330827

Peer reviewed

# A Physiological Time Series Dynamics-Based Approach to Patient Monitoring and Outcome Prediction

Li-Wei H. Lehman, Ryan P. Adams, Louis Mayaud, George B. Moody, Atul Malhotra, Roger G. Mark, and Shamim Nemati

**Abstract**—Cardiovascular variables such as heart rate (HR) and blood pressure (BP) are regulated by an underlying control system, and therefore, the time series of these vital signs exhibit rich dynamical patterns of interaction in response to external perturbations (e.g., drug administration), as well as pathological states (e.g., onset of sepsis and hypotension). A question of interest is whether “similar” dynamical patterns can be identified across a heterogeneous patient cohort, and be used for prognosis of patients’ health and progress. In this paper, we used a switching vector autoregressive framework to systematically learn and identify a collection of vital sign time series dynamics, which are possibly recurrent within the same patient and may be shared across the entire cohort. We show that these dynamical behaviors can be used to characterize the physiological “state” of a patient. We validate our technique using simulated time series of the cardiovascular system, and human recordings of HR and BP time series from an orthostatic stress study with known postural states. Using the HR and BP dynamics of an intensive care unit (ICU) cohort of over 450 patients from the MIMIC II database, we demonstrate that the discovered cardiovascular dynamics are significantly associated with hospital mortality (dynamic modes 3 and 9,  $p = 0.001$ ,  $p = 0.006$  from logistic regression after adjusting for the APACHE scores). Combining the dynamics of BP time series and SAPS-I or APACHE-III provided a more accurate assessment of patient survival/mortality in the hospital than using SAPS-I and APACHE-III alone ( $p = 0.005$  and  $p = 0.045$ ). Our results suggest that the discovered dynamics of vital sign time series may contain additional prognostic value beyond that of the baseline acuity measures, and can potentially be used as an independent predictor of outcomes in the ICU.

**Index Terms**—Intensive care unit, physiological control systems, switching linear dynamical systems.

Manuscript received October 31, 2013; revised March 30, 2014; accepted June 1, 2014. This work was supported by the National Institutes of Health Grants R01-EB001659 and R01GM104987 from the National Institute of Biomedical Imaging and Bioengineering, the James S. McDonnell Foundation Postdoctoral Grant, and the DARPA Young Faculty Award N66001-12-1-4219 Grant.

L. H. Lehman, G. B. Moody, and R. G. Mark are with the Massachusetts Institute of Technology, Cambridge, MA 02142 USA (e-mail: lilehman@mit.edu; george@mit.edu; rgmark@mit.edu).

R. P. Adams and S. Nemati are with the Harvard School of Engineering and Applied Sciences, Cambridge, MA 02138 USA (e-mail: rpa@seas.harvard.edu; shamim@seas.harvard.edu).

L. Mayaud is with the Institute of Biomedical Engineering, University of Oxford, Oxford, OX1 3PJ, UK, and also with Hôpital Raymond Poincaré, Garches 92380, France (e-mail: louis.mayaud@eng.ox.ac.uk).

A. Malhotra is with the Brigham and Women’s Hospital & Harvard Medical School, Boston, MA 02115 USA, and also with the UC San Diego Division of Pulmonary & Critical Care Medicine, La Jolla, CA 92037 USA (e-mail: amalhotra1@partners.org).

Color versions of one or more of the figures in this paper are available online at <http://ieeexplore.ieee.org>.

Digital Object Identifier 10.1109/JBHI.2014.2330827

## I. INTRODUCTION

MODERN clinical data acquisition systems are capable of continuously monitoring and storing measurements of patient vital signs, such as heart rate (HR) and blood pressure (BP), over multiple days of hospitalization [1]. Despite this continuous feed of data, commonly used acuity scores, such as APACHE and SAPS [2]–[5], are based on snap-shot values of these vital signs, typically the worst values during a 24-h period. However, physiologic systems generate complex dynamics in their output signals that reflect the state of the underlying control systems [6]–[8]. The objective of this study is to consider an approach to the analysis of critical care bedside monitoring that is based on the dynamical behaviors of vital sign time series.

The time series of vital signs (e.g., HR, BP) are multidimensional, high resolution (from once a second to once a minute), highly coupled due to presence of physiological feedback loops within the body [8], and remarkably nonstationary as a result of internally and externally induced changes in the state of the underlying control systems. For instance, time series of BP can exhibit oscillations on the order of seconds (e.g., due to the variations in sympathovagal balance), to minutes (e.g., as a consequence of fever, blood loss, or behavioral factors), to hours (e.g., due to humoral variations, sleep-wake cycle, or circadian effects) [9], [10]. A growing body of the literature is pointing to the clinical utility of vital signs time series dynamics to inform prognosis [11]–[17], and to provide early predictors of potentially life-threatening conditions in the intensive care unit (ICU) [18].

Techniques for modeling and analysis of cardiovascular and respiratory time series can be broadly classified into linear mechanistic models [19], [20] and nonlinear descriptive indices [6], [7], [21]. The linear techniques commonly used (often based on variants of autoregressive modeling) have the advantage of revealing the individual relationships among the observed variables (e.g., the noninvasive measures of baroreflex gain describes the relationship between HR and BP, excluding the possible influence of respiration). On the other hand, nonlinear indices of complexity are capable of capturing a richer set of dynamical behaviors, with less emphasis on physiological interpretability in terms of specific underlying mechanisms.

In this paper, we assume that although the underlying dynamical system may be nonlinear and nonstationary, and the stochastic noise components can be non-Gaussian, the dynamics can be approximated by a mixture of linear dynamical systems. Each such linear “dynamic” (or mode) is a time-dependent rule that describes how the future state of the system evolves from its

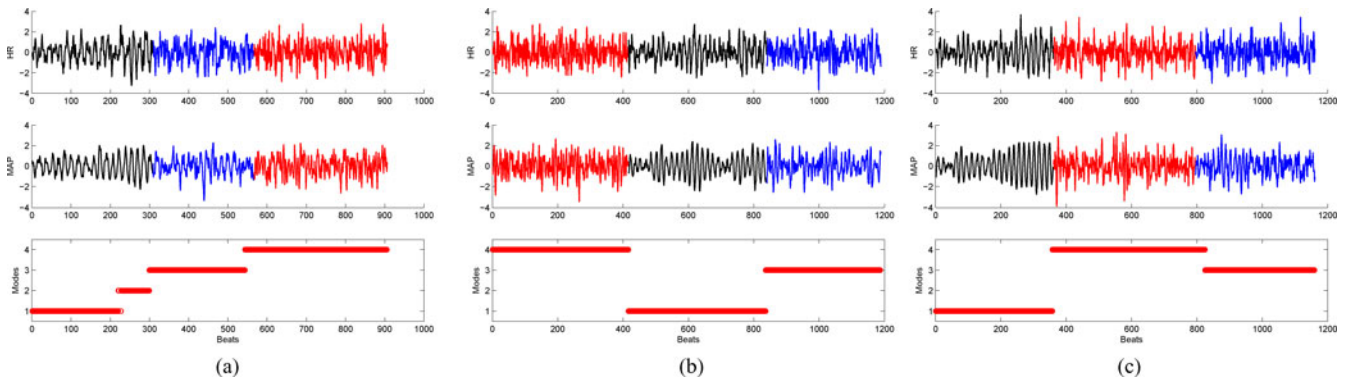


Fig. 1. Simulation study of the cardiovascular system. Three examples (out of the ten simulated time series) of HR and BP (after filtering) are shown in panels a, b, and c. In each case, the actual dynamics are color coded. The horizontal red lines show the inferred segmentation. The algorithm consistently assigned modes 4 and 3 to the dynamics color coded as red and blue, respectively, across all the simulated time series. The black dynamics are represented by modes 1 and 2. (a) Simulated subject 1, (b) Simulated subject 2, and (c) Simulated subject 3.

82 current state, centered around a given system equilibrium point.  
 83 Therefore, an ideal algorithm would be able to identify time series  
 84 segments that follow a “similar” dynamic, and would switch  
 85 to a different mode upon a change in the state of the underlying  
 86 system.

87 To formalize these objectives, we employed a switching vec-  
 88 tor autoregressive (SVAR) framework [22], [23]. Given a collec-  
 89 tion of time series from a cohort, the proposed SVAR framework  
 90 allows for simultaneous learning of the underlying dynamic be-  
 91 haviors or modes, and segmentation of the time series in terms  
 92 of the most likely dynamic describing the time series evolution  
 93 at any given point in time. The proposed framework enables  
 94 characterization of patients in terms of the dynamical modes  
 95 (e.g., the average time spent within the different modes), and  
 96 can potentially be used to capture changes in the underlying  
 97 cardiovascular control systems of human subjects in response  
 98 to internal (such as onset of infection) and external perturba-  
 99 tions (such as postural changes). Furthermore, we hypothesize  
 100 that when applied to vital sign time series of patients in a criti-  
 101 cal care setting, the proposed technique can be used to discover  
 102 dynamical modes with prognostic values for predicting clinical  
 103 outcomes of interests.

104 A preliminary version of this study was presented at the 34th  
 105 Annual International Conference of the IEEE Engineering in  
 106 Medicine and Biology Society (EMBC ’12) [14]. Here, we ex-  
 107 tend on our previous work to include a series of validation  
 108 studies, and a more comprehensive assessment of the utility of  
 109 the time series dynamics within the ICU.

110 The rest of this paper is organized as follows. We validated  
 111 the proposed technique using HR and BP time series from a  
 112 simulation dataset, and a human laboratory study of subjects  
 113 undergoing a tilt-table test, where the timing of the occurrence  
 114 of the different dynamics and the sharing of the dynamics across  
 115 multiple time series/subjects were known *a priori*. To test the  
 116 prognostic value of the discovered vital sign dynamics, we ap-  
 117 plied the proposed approach to the HR and BP dynamics of an  
 118 ICU cohort from the MIMIC II database [1] during the first 24 h  
 119 of their ICU stays, and tested whether cardiovascular dynamics  
 120 during the first 24 h of ICU admission are predictive of survival

and mortality after adjusting for the existing acuity scores, such  
 as SAPS-I and APACHE.

## II. MATERIALS AND METHODS

This section describes the utilized datasets, as well as the  
 proposed technique for discovery of shared dynamics among  
 patients, and assessment of risks and outcomes.

### A. Datasets

1) *Cardiovascular Simulation*: We simulated a cardiovas-  
 ular control system with bivariate time series of HR and BP.  
 The model is based on a delay recruitment model of HR and  
 BP regulation, as described in Fowler and McGuinness [24],  
 and McSharry *et al.* [25]. The model included a coupled system  
 of nonlinear delayed differential equations, controlling HR and  
 BP, with respiration as an exogenous input. We simulated ten  
 different multivariate time series of HR and mean arterial BP,  
 each including three different dynamics that become dominant  
 in a random order and last for a variable length of time. The  
 three dynamics (color-coded as red, blue, and black, respec-  
 tively, in Fig. 1) approximate aging-related autonomic changes;  
 a progressive reduction in parasympathetic gain (from 0.40 to  
 0.13 to 0.07 in normalized units; see [24]) and an increase in  
 sympathetic delay (from 3 to 5 s). To be consistent, we used the  
 same preprocessing step as the tilt-table experiment to remove  
 the steady-state baseline and any oscillation in the time series  
 slower than 100 beats/cycle (see below for details).

2) *Tilt-Table Experiment*: Time series of HR and BP were  
 acquired from ten healthy subjects (five males, five females)  
 undergoing a tilt-table test of orthostatic tolerance [26], [27].  
 The mean age was  $28.7 \pm 1.2$  years. The details of the pro-  
 tocol are described by Heldt in [27]. Briefly, subjects were  
 placed in a supine position. Tilting was performed from the  
 horizontal position to the vertical position and back to supine.  
 The study was approved by MIT’s Committee on the Use of  
 Humans as Experimental Subjects and the Advisory Board  
 of the MIT-MGH General Clinical Research Center [27].  
 Volunteers gave written, informed consent prior to participation

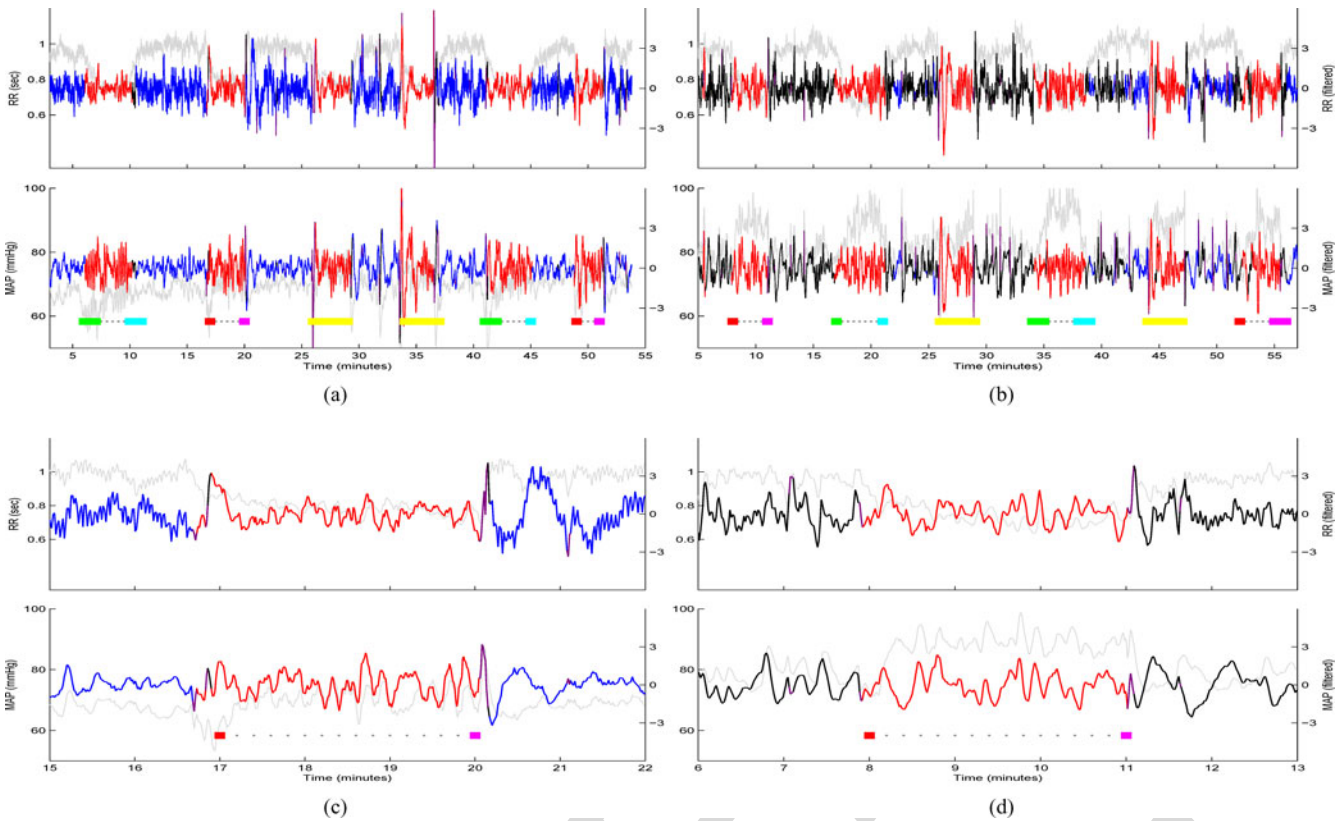


Fig. 2. Tilt-table study modeled using four dynamic modes—1 (Blue), 2 (Red), 3 (Black), and 4 (Purple). Two examples out of the ten recordings of HR and BP from the tilt-table experiment are shown in Panels a and b. Panels c and d show a zoomed in 7-min recording of HR and BP, while the subjects transition to/from supine to nonsupine positions after a fast tilt procedure. Actual values are in gray (*Y*-axis on left) and filtered values (*Y*-axis on right) are color coded based on the inferred dynamical modes. Note that Subjects 1 and 2 shared the same inferred nonsupine dynamics (in red); the algorithm consistently assigns the red mode to the nonsupine position for both subjects. The supine position for Subjects 1 and 2 are captured by the modes in blue and black, respectively. The purple mode seems to capture the high-frequency noise components of the time series. In each case, annotations for the actual tilt procedures performed are plotted as horizontal bars on the bottom of each figure and are color coded (green to cyan: slow tilt up and down to supine; red to pink: rapid tilt up and down to supine; yellow: stand up and back to supine). (a) Tilt Subject 1, (b) Tilt Subject 2, (c) Tilt Subject 1 (zoomed in), and (d) Tilt Subject 2 (zoomed in).

157 in the study. Since we were interested in the dynamics of inter-  
 158 action between HR and BP in the frequency range pertinent  
 159 to sympathetic and parasympathetic regulation [28], time series  
 160 of HR and BP were high-pass filtered to remove the steady-  
 161 state baseline and any oscillation in the time series slower than  
 162 100 beats/cycle. This filtering was done using a seventh-order  
 163 Butterworth digital filter with a cutoff frequency of 0.01 cycles/beat.  
 164 Example time series from before and after filtering  
 165 are shown in Fig. 2.

166 3) *MIMIC II Dataset*: The MIMIC II database [1], publicly  
 167 available via PhysioNet [29], includes clinical (laboratory  
 168 values, IV medications, etc.) and physiological data (HR,  
 169 BP, oxygen saturation, etc.) collected from the bedside moni-  
 170 tors (Component Monitoring System Intellivue MP-70; Philips  
 171 Healthcare, Andover, MA, USA) in ICUs of the Beth Israel  
 172 Deaconess Medical Center (BIDMC) in Boston. The MIMIC  
 173 II waveform database (version 2) includes approximately 4000  
 174 records of high-resolution physiological waveforms of adult  
 175 ICU patients with associated minute-by-minute (averages of  
 176 the calculated numerics during the previous minute) vital sign  
 177 trends. Data collection for the MIMIC II database was approved  
 178 by the Institutional Review Boards of BIDMC and the Mas-  
 179 sachusetts Institute of Technology (Cambridge, MA, USA). In-

dividual patient consent was waived because the study did not  
 impact clinical care and protected health information was de-  
 identified.

This study includes adult patients from the MIMIC II wave-  
 form database with at least 8 h of continuous minute-by-minute  
 HR and invasive arterial BP trends during the first 24 h in  
 the ICU. Patients with more than 15% of missing or invalid  
 samples (i.e., outside physiologically plausible bounds of 20 to  
 200 mmHg for mean pressures) were excluded from this study,  
 as were patients with missing SAPS I and APACHE scores.  
 The dataset contains over 9000 h of minute-by-minute HR and  
 invasive mean arterial BP measurements (over 20 h per patient  
 on average) from 453 adult patients collected during the first  
 24 h in the ICU. HR and BP time series were detrended. Gaus-  
 sian noise was used to fill in the missing or invalid values. The  
 median age of this cohort was 69 with an interquartile range of  
 (57, 79). 59% of the patients were male. Approximately 15%  
 (67 out of 453) of patients in this cohort died in the hospital;  
 28-day mortality of this cohort was approximately 19% (85 out  
 of 453). Distributions of the 453 patients in care units are 21%  
 coronary care unit (CCU), 42% cardiac surgery recovery unit  
 (CSRU), 26% medical intensive care unit (MICU), and 12%  
 surgical intensive care unit.

## 203 B. SVAR Modeling of Cohort Time Series

204 Our approach to discovery of shared dynamics among pa-  
 205 tients is based on the SVAR model [22]. For the  $n$ th patient  
 206 ( $n = 1 \dots N$ ), let  $y_t^{(n)}$  be a  $M \times 1$  vector of observed values of  
 207 the vital signs at time  $t$  ( $t = 1 \dots T^{(n)}$ ). We assume that there  
 208 exists a library of  $K$  possible dynamics or *modes*; a set of multi-  
 209 variate autoregressive model coefficient matrices  $\{A_p^{(k)}\}_{k=1}^K$  of  
 210 size  $M \times M$ , with maximal time lag  $p = 1 \dots P$ , and the cor-  
 211 responding noise covariances  $\{Q^{(k)}\}_{k=1}^K$ . Let  $s_t$  be a switching  
 212 variable, indicating the active dynamic mode at time  $t$ , and  
 213 evolving according to a Markovian dynamic with initial distri-  
 214 bution  $\pi^{(n)}$  and a  $K \times K$  transition matrix  $Z$ . Following these  
 215 definitions, an SVAR model for the  $n$ th patient is defined as

$$y_t^{(n)} = \sum_{p=1}^P A_p^{(s_t^{(n)})} y_{t-p}^{(n)} + w^{(s_t^{(n)})} \quad (1)$$

216 where the fluctuation term  $w^{(s_t^{(n)})}$  is assumed Gaussian dis-  
 217 tributed with covariance  $Q^{(s_t^{(n)})}$ . A collection of related time se-  
 218 ries can be modeled as switching between these dynamic behav-  
 219 iors which describe a locally coherent linear model that persists  
 220 over a segment of time. However, in practice, we neither know  
 221 the set of switching variables (i.e., segmentation of the time  
 222 series) nor the modes. In this study, we perform expectation-  
 223 maximization (EM) to find the maximum-likelihood set of  
 224 model parameters, as well as a factored estimate of the posterior  
 225 distribution over the latent switching variables. A comprehen-  
 226 sive treatment of the EM algorithm for SVAR is presented by  
 227 Murphy (1998) [22]. Briefly, EM is a two-pass iterative algo-  
 228 rithm: 1) in the expectation (E) step, we obtain the expected val-  
 229 ues of the latent switching variables  $\{s_t^{(n)}\}_{t=1}^T$  using a forward-  
 230 backward algorithm [22], and 2) in the maximization (M) step,  
 231 we update all the model parameters  $\{A_p^{(k)}\}, \{Q^{(k)}\}$ , the Markov  
 232 dynamics  $Z$ , and the initial conditions  $\pi^{(n)}$  that maximize the  
 233 expected complete data log likelihood. In our implementation  
 234 of the EM algorithm, we achieve *shared* dynamics by pooling  
 235 together all subjects' inferred variables in the M step. Iteration  
 236 through several steps of the EM algorithm results in learning a  
 237 set of  $K$  shared modes and a global transition matrix  $Z$  for all  
 238 the patients.

239 For the simulated and the tilt datasets, we modeled the beat-  
 240 by-beat HR/BP time series as a switching AR(5) process to  
 241 model most of the parasympathetic responses and at least some  
 242 of the sympathetic effects, without introducing an unduly com-  
 243 plex model. Minute-by-minute BP time series from MIMIC II  
 244 were modeled as a switching AR(3) process to capture a real os-  
 245 cillation and a possible trend per mode. The number of dynamic  
 246 modes ( $K = 20$ ) was determined using the Bayesian information  
 247 criterion (BIC) [30]. Briefly, we computed the BIC scores from  
 248 switching-VAR models using 5 to 30 modes. Results presented  
 249 were based on the model with the minimum BIC scores (20  
 250 modes).

251 1) *Parallel Computation for Scalable Learning*: One of the  
 252 advantages of the proposed technique is its scalability to hun-  
 253 dreds or thousands of patients, due to the parallel implementa-  
 254 tion of the inference step of the SVAR learning algorithm via

EM [22]. This parallelization strategy is effective since the ma-  
 255 jority of the computational cost of the SVAR training is in run-  
 256 ning the forward-backward algorithm, which can be done in  
 257 parallel for each patient time series. We used MATLAB's par-  
 258 allel computation toolbox in association with 120 nodes on our  
 259 computer cluster to perform a tenfold cross-validated study (12  
 260 cores per fold). Ten SVAR models were learned on the training  
 261 set of each of the folds, followed by mapping the corresponding  
 262 mode proportions to outcomes (e.g., hospital mortality) using  
 263 logistic regression. Next, mode assignments of time series in the  
 264 test set of each fold were inferred based on the modes learned  
 265 from the corresponding training set (by running only the infer-  
 266 ence), and the regression weights from the training fold were  
 267 used to predict outcomes. 268

## 269 C. Evaluation Methods and Statistical Analysis

Let us define a *mode proportion*  $MP_k^{(n)}$  as the proportion  
 270 of time the  $n$ th patient spends within the  $k$ th mode. Given the  
 271 maximum expected log-likelihood estimates of the switching  
 272 variables  $s_t$  from the EM algorithm, we have 273

$$MP_k^{(n)} = \frac{1}{T^{(n)}} \sum_{t=1}^{T^{(n)}} \text{Prob}(s_t^{(n)} = k). \quad (2)$$

For classification and prediction purposes, we characterize each  
 274 time series with its corresponding mode proportion (a  $1 \times K$   
 275 feature-vector), and use a logistic regression classifier to make  
 276 predictions about the outcome variables of interest. For illus-  
 277 tration of the algorithm's segmentation performance, each time  
 278 series sample is assigned to the dynamic mode with the maxi-  
 279 mum posterior probability. 280

281 1) *Time Series Classification and Outcome Prediction*: For  
 282 the simulated and the tilt-table experiment, we used the mode  
 283 proportions within each segment (e.g., supine versus nonsupine)  
 284 as inputs to a logistic regression classifier, and report the clas-  
 285 sification performance in discriminating between 1) the three  
 286 different dynamics (corresponding to different aging-related au-  
 287 tonomic changes) in the simulated dataset, and 2) two different  
 288 postural positions (supine versus nonsupine) in the tilt dataset.

289 To assess the predictive power of the dynamical modes, we  
 290 performed a tenfold cross-validation study. Ten SVAR models  
 291 were learned on the training set of each of the folds, followed  
 292 by mapping the corresponding mode proportions to outcomes  
 293 (e.g., hospital mortality) using logistic regression. Next, mode  
 294 assignments of time series in the test set of each fold was in-  
 295 ferred based on the modes learned from the corresponding train-  
 296 ing set (by running only the inference or the E-step), and the  
 297 regression weights from the training fold was used to predict  
 298 outcomes. We compared the mortality prediction performance  
 299 of our approach using the mode proportion from the top ten  
 300 most common dynamic modes with the existing acuity metrics,  
 301 SAPS I [2], APACHE III [4], and APACHE IV [5]. Comparison  
 302 of AUCs was based on the method described in [31].

303 2) *MIMIC Association Analysis*: We used univariate and  
 304 multivariate logistic regressions to examine the associations be-  
 305 tween dynamic mode proportions and hospital mortality. We

TABLE I  
PERFORMANCE OF MORTALITY PREDICTORS

	Hosp. Mortality (AUC)	28-Days Mortality (AUC)
HR <sub>dyn</sub>	0.59 (0.54, 0.68)	0.61 (0.51, 0.67)
BP <sub>dyn</sub> /HR <sub>dyn</sub>	0.64 (0.61, 0.71)	0.65 (0.64, 0.68)
BP <sub>dyn</sub>	<b>0.70 (0.67, 0.77)</b>	<b>0.66 (0.61, 0.73)</b>
SAPS I	0.65 (0.59, 0.71)	0.64 (0.56, 0.70)
BP <sub>dyn</sub> + SAPS I	<b>0.77 (0.69, 0.82)</b>	<b>0.71 (0.69, 0.79)</b>
APACHE III	0.80 (0.70, 0.84)	0.79 (0.65, 0.84)
BP <sub>dyn</sub> + APACHE III	<b>0.84 (0.79, 0.88)</b>	<b>0.79 (0.76, 0.86)</b>
APACHE IV	0.82 (0.77, 0.85)	0.83 (0.74, 0.86)
BP <sub>dyn</sub> + APACHE IV	<b>0.85 (0.80, 0.87)</b>	<b>0.82 (0.81, 0.88)</b>

built a separate multivariate logistic regression model for each of the discovered dynamic modes, with the mode proportion as the primary predictive variable, and APACHE IV as a covariate. For each mode, we reported its  $p$  value, odds ratio (OR, with 95% confidence interval), and adjusted OR (after including APACHE IV as a covariate). The Hosmer–Lemeshow  $p$  values (HL  $p$  values) were reported to assess the model fit. The odds ratios were per 10% increase in the mode proportion. Two-sided  $p$  values less than 0.05 were considered statistically significant. The analysis was performed to quantify the mortality risk associated with each dynamic mode; modes with significant ( $p < 0.05$ ) associations with mortality were established as either *low-risk* (OR  $< 1$ ), or *high-risk* (OR  $> 1$ ) dynamics depending on their odds ratios. Dynamic modes without statistically significant associations with mortality were *neutral* modes. The test of statistical significance was based on  $p$ -values after correcting for the false discovery rate (FDR) using the technique described in [32].

### III. RESULTS

#### A. Simulated Study

Fig. 1 shows two examples of simulated time series with the inferred segmentation. In all ten simulated cases, the algorithm was able to divide each time series into distinct segments corresponding to different underlying actual dynamics. The sharing of the dynamics is consistent across the different time series. Using the mode proportion from each segment for multilabel classification, the algorithm achieved classification accuracy of 100%.

#### B. Tilt-Table Experiment

Fig. 2 shows the segmentation results for two subjects. Note that the two subjects shared the same inferred nonsupine dynamics (in red); the algorithm consistently assigns the red mode to the nonsupine position for both subjects. The application of logistic regression with tenfold cross-validation yielded a median AUC of 1.00 with an interquartile range of (0.98, 1.00).

#### C. MIMIC II Database

1) *Mortality Prediction*: Table I evaluates the prognostic power of HR and BP dynamic features (HR<sub>dyn</sub> and BP<sub>dyn</sub>).

SAPS I, APACHE III, and APACHE IV are used as the baselines. Median AUCs (from tenfold cross validation) and the interquartile range are shown. Note that the BP dynamics outperformed both the HR and HR and BP combined dynamic features. Subsequent analyses focus on the predictive power of the BP dynamics in comparison to the baseline. For each baseline, we show the performance from the baseline alone, and the combined approach (combining BP dynamics and the baseline).

The application of tenfold cross-validation demonstrated that dynamic features from BP alone achieved a median AUC of 0.70, comparable to 0.65 from SAPS I. In comparison, using standard deviation of the mean arterial BP resulted in a median AUC (IQR) of 0.55 (0.43, 0.63).

Combining dynamic BP features with SAPS I resulted in an improved prediction power both in hospital mortality prediction ( $p = 0.005$ ) and 28-day mortality prediction ( $p = 0.002$ ). Combining dynamic features with APACHE III significantly outperformed APACHE III alone ( $p = 0.045$ ) with an improvement in median AUC from 0.80 to 0.84 in hospital mortality prediction. These results indicate that the dynamic features from vital signs contain complementary information to the SAPS I and APACHE III scores.

State-of-the-art risk score APACHE IV achieved better prediction performance than the BP dynamic features alone ( $p = 0.008$ ). Adding BP dynamics to APACHE IV yielded a slight performance improvement from a median AUC of 0.82 to 0.85, however, the performance gain was not statistically significant.

2) *Association Analysis*: Table II presents logistical regression analyses to test the associations between the proportion of time patients spent in each of the top ten most common BP dynamics and hospital mortality. See Fig. 3 for illustrations of these dynamic modes. Dynamic modes were numbered based on their prevalence across the entire cohort (i.e., mode 1 is the most common dynamic mode). Our results indicate that six of the modes had significant associations (after FDR correction) with hospital mortality. Specifically, two dynamic modes (modes 3 and 5) were significant “high-risk” modes ( $p < 0.001$ ,  $p < 0.001$ ) in which increased proportions of time in these modes were associated with higher hospital mortality with odds ratios 1.81 (1.41, 2.32), 1.36 (1.15, 1.61) respectively.

Dynamic modes 1, 9, 7, and 2 were “low-risk” modes in which increasing proportions of time in these modes were significantly associated with a decreased risk of hospital mortality, with odds ratios less than one. Table II lists the AR coefficients and covariances of the two high-risk and four low-risk dynamic modes, as well as their respective associations with hospital mortality. Note that the high-risk modes appear to correspond to less variability in their dynamics.

For the multivariate analysis (see the right panel in Table II), each row is a separate multivariate model, in which the mode proportion for a given target mode is the primary predictive variable, and APACHE IV is added as a control variable in the multivariate model. Results from multivariate logistic regression indicate that two of the modes (modes 3 and 9) remain significant predictors of patients’ outcome even after adjustment for APACHE IV scores ( $p = 0.001$ ,  $p = 0.006$ ), indicating that the proportion of time patients spent in these two dynamic modes

TABLE II  
ASSOCIATIONS OF BP DYNAMIC MODES AND HOSPITAL MORTALITY

Mode	AR Coef	Cov.	P-Val	OR(95%CI)	Adjusted P-Val	Adjusted OR(95%CI)	HL PVAL
3	(0.66, 0.22, 0.12)	0.58	< <b>0.001</b>	1.81 (1.41, 2.32)	<b>0.001</b>	1.60 (1.21 2.11)	0.64
5	(1.00, 0.00, -0.00)	0.22	< <b>0.001</b>	1.36 (1.15, 1.61)	0.426	1.08 (0.89 1.32)	0.16
1	(0.66, 0.16, 0.17)	2.69	<b>0.002</b>	0.59 (0.42, 0.82)	0.489	0.88 (0.62 1.26)	0.54
9	(1.50, -0.65, 0.06)	7.26	<b>0.002</b>	0.25 (0.10, 0.62)	<b>0.006</b>	0.26 (0.10 0.68)	0.80
7	(1.00, -0.01, -0.00)	3.46	<b>0.003</b>	0.30 (0.13, 0.67)	0.124	0.54 (0.25 1.18)	0.58
2	(0.79, 0.05, 0.12)	8.81	<b>0.005</b>	0.65 (0.48, 0.88)	0.265	0.84 (0.62 1.14)	0.69
10	(1.05, -0.01, -0.02)	0.71	0.032	2.95 (1.10, 7.94)	0.791	1.18 (0.36 3.88)	0.07
8	(0.44, 0.30, 0.24)	1.27	0.373	1.18 (0.82, 1.69)	0.318	1.22 (0.82 1.82)	0.22
4	(0.96, -0.01, 0.04)	1.31	0.417	0.81 (0.48, 1.36)	0.887	0.96 (0.53 1.72)	0.02
6	(0.92, -0.10, 0.07)	46.70	0.419	0.83 (0.53, 1.30)	0.658	0.90 (0.57 1.43)	0.08

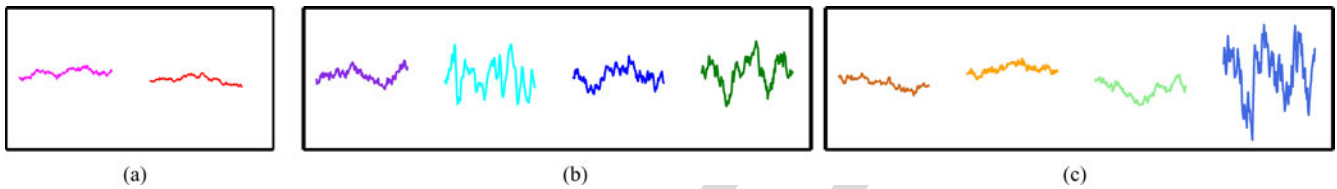


Fig. 3. Discovered dynamic modes of mean arterial BP of 453 patients during the first 24 h in the ICU. Figure shows the top ten most common dynamic modes, simulated using the AR coefficients from each dynamic mode. High-risk dynamic modes (from left to right): 3 (Magenta), 5 (Red). Low-risk dynamic modes: 1 (Violet), 9 (Cyan), 7 (Blue), and 2 (Green). Neutral dynamic modes: 10 (Brown), 8 (Orange), 4 (Light Green), 6 (Royal Blue). All modes were simulated and plotted with the same time duration (150 min) and amplitude scale. (a) High-risk modes, (b) Low-risk modes, and (c) Neutral modes.

401 during the first 24 h in the ICU are independent risk predictors  
402 of hospital mortality.

403 3) *Example Time Series of Patients With Estimated Mortality*  
404 *Risks Over Time:* Fig. 3 shows examples of low-risk and high-  
405 risk dynamical modes learned using the SVAR technique (see  
406 Table II for the odds-ratio associated with each mode). BP time  
407 series of four patients are presented in Fig. 4 panels (a) and  
408 (b). Hourly risk scores (dark green lines) were computed as the  
409 probability of death from the logistic function using a sliding  
410 window of 6 h to illustrate that these risk scores could be updated  
411 on a continuous basis for real-time monitoring purposes.

412 Panel (a) shows two of the patients with the highest risk scores  
413 (within the test set) at the end of the 24-h period; both patients  
414 died in the hospital. Panel (b) shows two patients with a decreas-  
415 ing trend in their risk scores during their first day in the ICU;  
416 both patients survived the hospital stay. All four patients were  
417 from the same test set, with mode assignment inferred based  
418 on dynamic modes learned from the corresponding training set.  
419 Note that as time progresses, patients in panel (a) tend to spend  
420 more time in the high-risk dynamic modes (mode 3 in magenta,  
421 mode 5 in red); their estimated mortality risks rise accordingly  
422 over time. In contrast, panel (b) patients show a decreasing trend  
423 in mortality risks as they transit to lower-risk dynamic modes  
424 over time.

#### 425 IV. DISCUSSION AND CONCLUSION

426 We presented a SVAR framework to systematically learn and  
427 identify dynamic behaviors from vital sign time series within a  
428 patient cohort. We demonstrated that the discovered dynamics  
429 may contain prognostic values and can be used for prediction

and tracking of a patient's propensity to survive a hospital stay, 430  
as well as their 28-days survival. Interestingly, the BP time 431  
series dynamics alone had a comparable performance to that of 432  
the SAPS I score which uses age and the most extreme values of 433  
13 variables, including systolic BP, HR, temperature, respiratory 434  
rate, urinary output, blood nitrogen, hematocrit, white blood cell 435  
count, serum glucose, serum potassium, serum sodium, serum 436  
bicarbonate, and Glasgow coma score. 437

438 Additionally, our results indicate that the BP dynamics may 438  
contain complimentary information to existing acuity metrics, 439  
which assess the health of multiple organ systems based on a vari- 440  
ety of physiological and lab variables. Specifically, combining 441  
the dynamics of BP time series and SAPS I or APACHE III pro- 442  
vided a more accurate assessment of patient survival/mortality 443  
in the hospital ( $p = 0.005$  and  $p = 0.045$ ) than using SAPS I 444  
and APACHE III alone. 445

446 Association analysis of individual dynamic mode and hospi- 446  
tal mortality revealed that two of the dynamic modes (modes 447  
3 and 9) remained significant predictors of patients' outcome 448  
even after adjusting for APACHE IV scores, indicating that the 449  
proportion of time patients spent in these two dynamic modes 450  
during the first 24 h in the ICU may contain additional, inde- 451  
pendent prognostic value beyond that in the APACHE IV acuity 452  
score. Future work remains to investigate the prognostic power 453  
of these discovered dynamic modes using a larger cohort. 454

455 The dynamic features can be calculated in an online man- 455  
ner without delay, and well before the end of the first 24 h of 456  
the ICU stay as is required for the standard risk scores. One 457  
possible online deployment strategy is to construct a library of 458  
dynamic modes on archived patient data, and assign each in- 459  
coming time series sample (or a sliding window of samples) 460

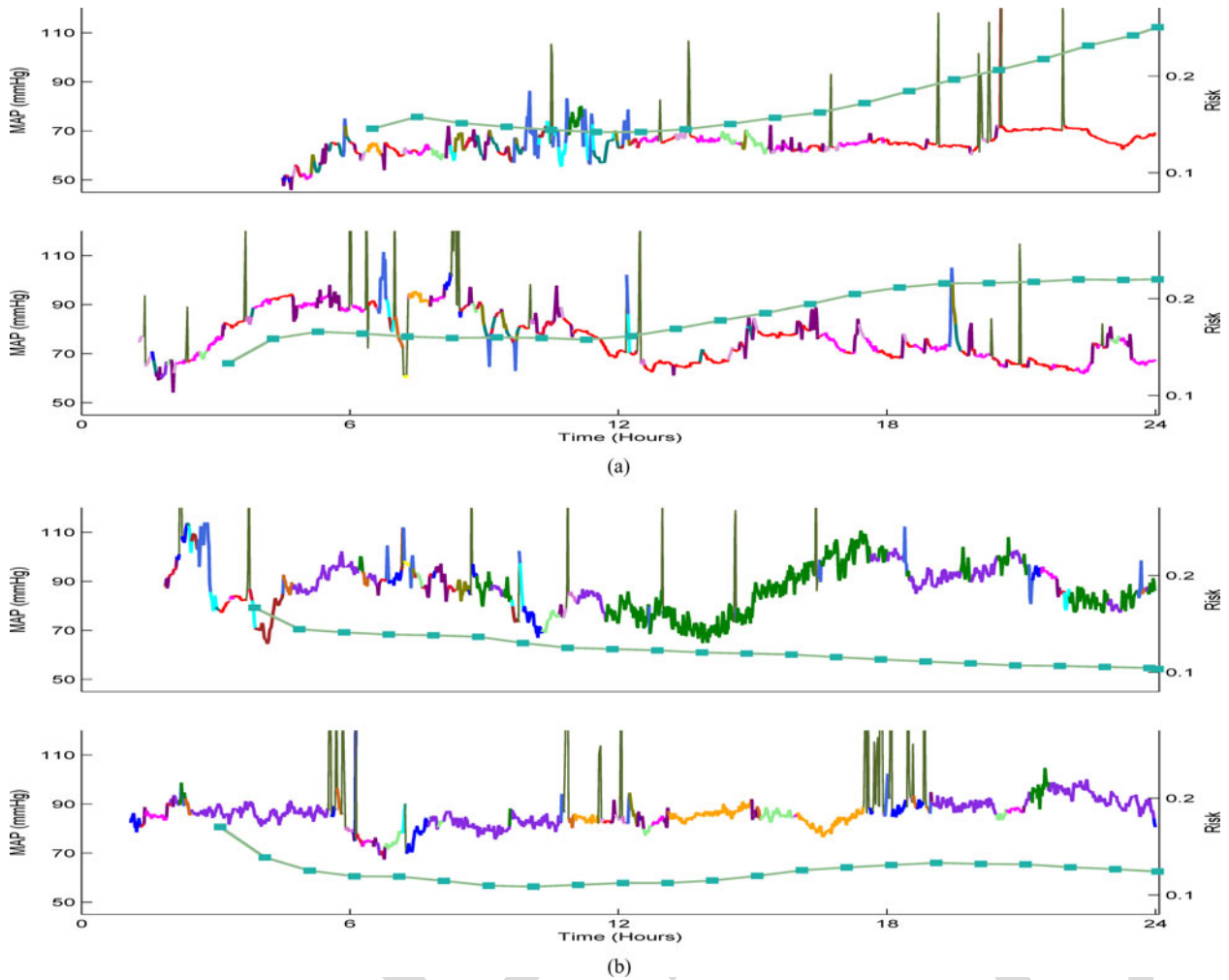


Fig. 4. Mortality risk scores and mean arterial BP of four patients during the first 24 h in the ICU. Samples are color coded by their mode assignment. Mortality risk scores, computed as the probability of death from the logistic regression, were based on mode proportions from a 6-h sliding window by stride of 1 h; estimated risks were plotted as dark green lines with scale indicated by  $y$ -axes on the right side of each graph. BP measurements plotted in original units (before detrending). All four patients were from the same test set, with dynamic modes and logistic regression parameters learned from the corresponding training set. (a) Patients with the highest ending risk scores at the end of the first day ICU stay. Patients were from MICU (top) and CCU (bottom). Both patients died in the hospital and (b) Patients with decreasing risk scores during their first day ICU stays. Patients were from CSRU (top) and CCU (bottom). Both patients survived the hospital stay.

461 to the most likely mode in the library (for instance, by  
 462 using the Viterbi algorithm [16], [22]). Recent studies suggest  
 463 that therapeutic interventions not only should aim at maintaining  
 464 the mean BP within an acceptable range, but also should direct  
 465 the patient's trajectory toward healthy dynamical regimes with  
 466 enhanced variability [10]. Thus, a real-time implementation of  
 467 the technique presented here may provide clinicians with a tool  
 468 for quantification of the effectiveness of such interventions in  
 469 the ICU.

470 We showed that changes in the dynamics of HR and BP, either  
 471 as a result of an altered underlying control system (aging-related  
 472 changes in the simulated data) or due to external perturbations  
 473 (positional changes in the tilt-table experiment), can be captured  
 474 in an automated fashion. Since the proposed framework is built  
 475 on the dynamical systems framework (which includes the class  
 476 of vector autoregressive models), the discovered modes can be  
 477 used to reveal the oscillations that are present within the indi-  
 478 vidual time series, and therefore can be used to extract useful

479 indices of HR and BP variability (assuming beat-to-beat time  
 480 series). Moreover, given beat-to-beat multivariate time-series of  
 481 vital-signs, one may use the learned dynamics to derive the di-  
 482 rectional transfer functions of the system [8] (e.g., baroreflex  
 483 control of HR and BP).

484 Association analysis using the minute-by-minute MIMIC-  
 485 II BP time series revealed that the high-risk modes often  
 486 correspond to less variable dynamical patterns. It is interest-  
 487 ing to note that such low-frequency variability, observed at  
 488 the minute-to-minute scale, is associated with an enhanced  
 489 chance of survival, corresponding well to the existing HR/BP  
 490 variability literature using beat-by-beat vital sign time series  
 491 [10], [12], [13], [33]. The working hypothesis of our ongo-  
 492 ing research is that the observed dynamical patterns are due  
 493 to patients' underlying physiology, patient-specific response  
 494 to clinical interventions, and measurement artifacts. Future  
 495 developments of machine-learning techniques should aim at  
 496 combining time series dynamics with contextual information



497 pertaining to clinical intervention (administration of fluids,  
498 pressors, and titration of medications) to further investigate  
499 the clinical and physiological interpretation of the discovered  
500 modes.

501 The SVAR framework allows for defining a notion of “sim-  
502 ilarity” among multivariate physiological time series based on  
503 their underlying shared dynamics. Therefore, one may consider  
504 two subjects to be similar if their underlying vital signs time se-  
505 ries exhibit similar dynamics in response to external (e.g., tilting  
506 of body) or internal perturbations (e.g., onset of blood infection).  
507 This approach provides an improvement over time series sim-  
508 ilarity measures based on trend-detection [34], wavelet-based  
509 symbolic representations [35], or Gaussian Mixture modeling  
510 [36] due to its compact representation and sharing of the model  
511 parameters within and across time series. Prior work using a fac-  
512 torial switching linear dynamical systems for patient monitoring  
513 [37] focused on detection of events associated with artifactual  
514 measurements and pathological states. Our study, in contrast,  
515 jointly models multiple time series across a large patient cohort  
516 to identify phenotypic dynamical patterns for patient outcome  
517 prediction.

518 Although we used mortality as our target outcome, there are  
519 many other physiological events of significant interest, includ-  
520 ing timely and successful discontinuation of procedures such  
521 as hemodialysis [38] or mechanical ventilation [39], as well as  
522 prediction of potentially life-threatening clinical events such as  
523 onset of severe sepsis and hypotension [13]. Other short- and  
524 long-term outcomes such as probability of readmission to hospi-  
525 tal and long-term cognitive impairment beyond ICU [40] also  
526 play an important role in closing the gap between the critical care  
527 medicine, primary care doctors, and other healthcare providers.

528 Current and ongoing work involve combining the switching  
529 linear dynamical system framework with all available clinical  
530 data, including lab tests, medication records, and nursing notes  
531 [41] to devise a comprehensive risk score, capable of integrating  
532 clinical data of diverse modality over long temporal stretches  
533 (order of hours to days). This will allow us to investigate whether  
534 continuous patient monitoring based on vital signs dynamics,  
535 and other types of sequential data, can alert clinicians to deter-  
536 iorating patient conditions at an earlier stage than the existing  
537 acuity scores, and result in improved patient care and outcome  
538 both within ICU and after hospital discharge. Such analysis is  
539 likely to provide some insight into the promise of large-scale  
540 critical care databases for the future of medicine.

#### 541 ACKNOWLEDGMENT

542 The authors would like to thank Dr. T. Heldt (Harvard-MIT  
543 Division of Health Sciences and Technology) for kindly provid-  
544 ing the tilt-table data analyzed in this study.

#### 545 REFERENCES

546 [1] M. Saeed, M. Villarroel, A. T. Reisner, G. Clifford, L. H. Lehman,  
547 G. Moody, T. Heldt, T. H. Kyaw, B. Moody, and R. G. Mark, “Multi-  
548 parameter intelligent monitoring in intensive care (MIMIC II): A public-  
549 access intensive care unit database,” *Crit. Care Med.*, vol. 39, no. 5,  
550 pp. 952–960, May 2011.

[2] J. R. Le Gall, P. Loirat, A. Alperovitch, P. Glaser, C. Granthil, D. Mathieu, 551  
P. Mercier, R. Thomas, and D. Villers, “A simplified acute physiology 552  
score for ICU patients,” *Crit. Care Med.*, vol. 12, no. 11, pp. 975–977, 553  
Nov. 1984.

[3] J. R. Le Gall, S. Lemeshow, and F. Saulnier, “A new simplified acute 554  
physiology score (SAPS II) based on a European/North American 555  
multicenter study,” *J. Amer. Med. Assoc.*, vol. 270, no. 24, pp. 2957–2963, 556  
1993.

[4] W. A. Knaus, D. P. Wagner, E. A. Draper, J. E. Zimmerman, M. Bergner, 557  
P. G. Bastos, C. A. Sirio, D. J. Murphy, T. Lotring, A. Damiano, 558  
and F. Harrell, “The APACHE III prognostic system,” *Chest*, vol. 100, 559  
pp. 1619–1636, 1991.

[5] J. E. Zimmerman, A. A. Kramer, D. S. McNair, and F. M. Malila, “Acute 560  
physiology and chronic health evaluation (APACHE) IV: Hospital mor- 561  
tality assessment for today’s critically ill patients,” *Crit. Care Med.*, 562  
vol. 34, no. 5, pp. 1297–1310, May 2006.

[6] P. C. Ivanov, L. A. Amaral, A. L. Goldberger, S. Havlin, M. G. Rosenblum, 563  
Z. R. Struzik, and H. E. Stanley, “Multifractality in human heartbeat 564  
dynamics,” *Nature*, vol. 399, pp. 461–465, 1999.

[7] M. Costa, A. L. Goldberger, and C. K. Peng, “Multiscale entropy analysis 565  
of complex physiologic time series,” *Phys. Rev. Lett.*, vol. 89, no. 6, 566  
p. 068102, 2002.

[8] S. Nemati, B. A. Edwards, S. A. Sands, P. J. Berger, A. Wellman, G. C. 567  
Verghese, A. Malhotra, and J. P. Butler, “Model-based characterization 568  
of ventilatory stability using spontaneous breathing,” *J. Appl. Physiol.*, 569  
vol. 111, no. 1, pp. 55–67, 2011.

[9] G. Mancia, “Short-and long-term blood pressure variability present and 570  
future,” *Hypertension*, vol. 60, no. 2, pp. 512–517, 2012.

[10] G. Parati, J. E. Ochoa, C. Lombardi, and G. Bilo, “Assessment and 571  
management of blood-pressure variability,” *Nat. Rev. Cardiol.*, vol. 10, 572  
pp. 143–155, 2013.

[11] T. G. Buchman, “Nonlinear dynamics, complex systems, and the patho- 573  
biology of critical illness,” *Current Opinion Crit. Care*, vol. 10, no. 5, 574  
pp. 378–382, 2004.

[12] S. Saria, A. K. Rajani, J. Gould, D. Koller, and A. Penn, “Integration 575  
of early physiological responses predicts later illness severity in preterm 576  
infants,” *Sci. Translational Med.*, vol. 2, no. 48, pp. 48–65, 2010.

[13] J. R. Moorman, J. B. Delos, A. A. Flower, H. Cao, B. P. Kovatchev, 577  
J. S. Richman, and D. E. Lake, “Cardiovascular oscillations at the bed- 578  
side: Early diagnosis of neonatal sepsis using heart rate characteris- 579  
tics monitoring,” *Physiol. Meas.*, vol. 32, no. 11, pp. 1821–1832, Nov. 580  
2011.

[14] L. H. Lehman, S. Nemati, R. P. Adams, and R. G. Mark, “Discovering 581  
shared dynamics in physiological signals: Application to patient monitor- 582  
ing in ICU,” in *Proc. IEEE Eng. Med. Biol. Soc.*, 2012, pp. 5939–5942. 583

[15] J. Wiens, E. Horvitz, and J. V. Guttag, “Patient risk stratification for 584  
hospital-associated C. diff as a time-series classification task,” in *Adv. 585  
Neural Inf. Process. Syst.*, 2012, pp. 476–484.

[16] L. H. Lehman, S. Nemati, R. P. Adams, G. Moody, A. Malhotra, and R. G. 586  
Mark, “Tracking progression of patient state of health in critical care using 587  
inferred shared dynamics in physiological time series,” in *Proc. IEEE Eng. 588  
Med. Biol. Soc.*, 2013, pp. 7072–7075.

[17] L. Mayaud, P. S. Lai, G. D. Clifford, L. Tarassenko, L. A. Celi, and 589  
D. Annane, “Dynamic data during hypotensive episode improves mortality 590  
predictions among patients with sepsis and hypotension,” *Crit. Care Med.*, 591  
vol. 41, no. 4, pp. 954–962, 2013.

[18] M. Blount, M. R. Ebling, J. M. Eklund, A. G. James, C. McGregor, 592  
N. Percival, K. P. Smith, and D. Sow, “Real-time analysis for intensive 593  
care: Development and deployment of the artemis analytic system,” *IEEE 594  
Eng. Med. Biol. Mag.*, vol. 29, no. 2, pp. 110–118, Mar–Apr 2010.

[19] A. Porta, T. Bassani, V. Bari, E. Tobaldini, A. C. M. Takahashi, A. M. 595  
Catai, and N. Montano, “Model-based assessment of baroreflex and cardiopulmonary 596  
couplings during graded head-up tilt,” *Comput. Biol. Med.*, 597  
vol. 42, no. 3, pp. 298–305, Mar. 2012.

[20] R. Mukkamala, J. M. Mathias, T. J. Mullen, R. J. Cohen, and R. Freeman, 598  
“System identification of closed-loop cardiovascular control mechanisms: 599  
Diabetic autonomic neuropathy,” *Amer. J. Physiol.-Regulatory, Integrative 600  
Comparative Physiol.*, vol. 276, no. 3, pp. R905–R912, 1999.

[21] I. Korhonen, “Multivariate closed-loop model for analysis of cardiovas- 601  
cular dynamics,” *Methods Inf. Med.*, vol. 36, pp. 264–267, 1997.

[22] K. P. Murphy, “Switching Kalman filter,” Compaq Cambridge Res. Lab., 602  
Cambridge, CA, USA, Tech. Rep. 98-10, 1998.

[23] S. Nemati, L. H. Lehman, R. P. Adams, and A. Malhotra, “Discovering 603  
shared cardiovascular dynamics within a patient cohort,” in *Proc. IEEE 604  
Eng. Med. Biol. Soc.*, 2012, pp. 6526–6529. 605  
606  
607  
608  
609  
610  
611  
612  
613  
614  
615  
616  
617  
618  
619  
620  
621  
622  
623  
624  
625

- 626 [24] A. Fowler and M. McGuinness, "A delay recruitment model of the cardiovascular control system," *J. Math. Biol.*, vol. 51, no. 5, pp. 508–526, 2005. 664
- 627
- 628
- 629 [25] P. McSharry, M. McGuinness, and A. Fowler, "Confronting a cardiovascular system model with heart rate and blood pressure data," in *Proc. Comput. Cardiol.*, 2005, pp. 587–590. 665
- 630
- 631
- 632 [26] T. Heldt, M. B. Oefinger, M. Hoshiyama, and R. G. Mark, "Circulatory response to passive and active changes in posture," in *Proc. Comput. Cardiol.*, 2003, vol. 30, pp. 263–266. 666
- 633
- 634
- 635 [27] T. Heldt, "Computational models of cardiovascular response to orthostatic stress," Ph.D. dissertation, Massachusetts Inst. Technol., Cambridge, MA, USA, Sep. 2004. 667
- 636
- 637
- Q3 638 [28] A. Camm, M. Malik, J. Bigger, G. Breithardt, S. Cerutti, R. Cohen, P. Coumel, E. Fallen, H. Kennedy, R. Kleiger *et al.*, "Heart rate variability: Standards of measurement, physiological interpretation and clinical use. Task force of the european society of cardiology and the north american society of pacing and electrophysiology," *Circulation*, vol. 93, no. 5, pp. 1043–1065, 1996. 668
- 639
- 640
- 641
- 642
- 643
- Q4 644 [29] A. L. Goldberger, L. A. N. Amaral, L. Glass, J. M. Hausdorff, P. C. Ivanov, R. G. Mark, J. E. Mietus, G. B. Moody, C.-K. Peng, and H. E. Stanley. (2000 Jun. 13) PhysioBank, PhysioToolkit, and PhysioNet: Components of a new research resource for complex physiologic signals. *Circulation*. [Online]. 101(23), pp. e215–e220. Available: <http://circ.ahajournals.org/cgi/content/full/101/23/e215> PMID:1085218; doi: 10.1161/01.CIR.101.23.e215 669
- 645
- 646
- 647
- 648
- 649
- 650
- 651 [30] G. Schwarz, "Estimating the dimension of a model," *Ann. Statist.*, vol. 6, no. 2, pp. 461–464, 1978. 670
- 652
- 653 [31] E. R. DeLong, D. M. DeLong, and D. L. Clarke-Pearson, "Comparing the areas under two or more correlated receiver operating characteristic curves: A nonparametric approach," *Biometrics*, vol. 44, pp. 837–845, 1988. 671
- 654
- 655
- 656
- 657 [32] Y. Benjamini and Y. Hochberg, "Controlling the false discovery rate: A practical and powerful approach to multiple testing," *J. Roy. Statist. Soc.*, vol. 57, no. 1, pp. 289–300, 1995. 672
- 658
- 659
- 660 [33] W. P. Riordan Jr., P. R. Norris, J. M. Jenkins, and J. A. Morris Jr., "Early loss of heart rate complexity predicts mortality regardless of mechanism, anatomic location, or severity of injury in 2178 trauma patients," *J. Surg. Res.*, vol. 156, no. 2, pp. 283–289, 2009. 673
- 661
- 662
- 663
- 664 [34] R. K. Arent and J. D. Charlton, "A critical review of trend-detection methodologies for biomedical monitoring systems," *Crit. Rev. Biomed. Eng.*, vol. 17, no. 6, pp. 621–659, 1990. 665
- 666
- 667 [35] M. Saeed and R. G. Mark, "A novel method for the efficient retrieval of similar multiparameter physiologic time series using wavelet-based symbolic representations," in *Proc. AMIA Annu. Symp.*, 2006, pp. 679–683. 668
- 669
- 670
- 671 [36] L. H. Lehman, M. Saeed, G. Moody, and R. G. Mark, "Similarity-based searching in multi-parameter time series databases," in *Proc. Comput. Cardiol.*, 2008, pp. 653–656. 672
- 673
- 674 [37] J. A. Quinn, C. K. Williams, and N. McIntosh, "Factorial switching linear dynamical systems applied to physiological condition monitoring," *IEEE Trans. Pattern Anal. Mach. Intell.*, vol. 31, no. 9, pp. 1537–1551, Sep. 2009. 675
- 676
- 677
- 678 [38] N. Gibney, E. Hoste, E. A. Burdman, T. Bunchman, V. Kher, R. Viswanathan, R. L. Mehta, and C. Ronco, "Timing of initiation and discontinuation of renal replacement therapy in AKI: Unanswered key questions," *Clin. J. Amer. Soc. Nephrol.*, vol. 3, no. 3, pp. 876–880, 2008. 679
- 680
- 681
- 682 [39] A. Mikhno and C. M. Ennett, "Prediction of extubation failure for neonates with respiratory distress syndrome using the MIMIC-II clinical database," in *Proc. IEEE Eng. Med. Biol. Soc.*, 2012, pp. 5094–5097. 683
- 684
- 685 [40] A. E. Wolters, A. J. Slooter, A. W. van der Kooi, and D. van Dijk, "Cognitive impairment after intensive care unit admission: A systematic review," *Intensive Care Med.*, vol. 39, no. 3, pp. 376–386, 2013. 686
- 687
- 688 [41] L. H. Lehman, M. Saeed, W. Long, J. Lee, and R. G. Mark, "Risk stratification of ICU patients using topic models inferred from unstructured progress notes," in *Proc. AMIA Annu. Symp Amer. Med. Informat. Assoc.*, 2012, pp. 505–511. 689
- 690
- 691
- Authors', photographs and biographies not available at the time of publication. 692
- 693

- 695 Q1. Author: Please verify whether the affiliations of all the authors are OK as set.  
696 Q2. Author: Please provide complete page range in Ref. [7].  
697 Q3. Author: Please provide the department name in Ref. [27].  
698 Q4. Author: Please provide names of all the authors in Ref. [28].

IEEE  
Proof

# A Physiological Time Series Dynamics-Based Approach to Patient Monitoring and Outcome Prediction

Li-Wei H. Lehman, Ryan P. Adams, Louis Mayaud, George B. Moody, Atul Malhotra, Roger G. Mark, and Shamim Nemati

**Abstract**—Cardiovascular variables such as heart rate (HR) and blood pressure (BP) are regulated by an underlying control system, and therefore, the time series of these vital signs exhibit rich dynamical patterns of interaction in response to external perturbations (e.g., drug administration), as well as pathological states (e.g., onset of sepsis and hypotension). A question of interest is whether “similar” dynamical patterns can be identified across a heterogeneous patient cohort, and be used for prognosis of patients’ health and progress. In this paper, we used a switching vector autoregressive framework to systematically learn and identify a collection of vital sign time series dynamics, which are possibly recurrent within the same patient and may be shared across the entire cohort. We show that these dynamical behaviors can be used to characterize the physiological “state” of a patient. We validate our technique using simulated time series of the cardiovascular system, and human recordings of HR and BP time series from an orthostatic stress study with known postural states. Using the HR and BP dynamics of an intensive care unit (ICU) cohort of over 450 patients from the MIMIC II database, we demonstrate that the discovered cardiovascular dynamics are significantly associated with hospital mortality (dynamic modes 3 and 9,  $p = 0.001$ ,  $p = 0.006$  from logistic regression after adjusting for the APACHE scores). Combining the dynamics of BP time series and SAPS-I or APACHE-III provided a more accurate assessment of patient survival/mortality in the hospital than using SAPS-I and APACHE-III alone ( $p = 0.005$  and  $p = 0.045$ ). Our results suggest that the discovered dynamics of vital sign time series may contain additional prognostic value beyond that of the baseline acuity measures, and can potentially be used as an independent predictor of outcomes in the ICU.

**Index Terms**—Intensive care unit, physiological control systems, switching linear dynamical systems.

Manuscript received October 31, 2013; revised March 30, 2014; accepted June 1, 2014. This work was supported by the National Institutes of Health Grants R01-EB001659 and R01GM104987 from the National Institute of Biomedical Imaging and Bioengineering, the James S. McDonnell Foundation Postdoctoral Grant, and the DARPA Young Faculty Award N66001-12-1-4219 Grant.

L. H. Lehman, G. B. Moody, and R. G. Mark are with the Massachusetts Institute of Technology, Cambridge, MA 02142 USA (e-mail: lilehman@mit.edu; george@mit.edu; rgmark@mit.edu).

R. P. Adams and S. Nemati are with the Harvard School of Engineering and Applied Sciences, Cambridge, MA 02138 USA (e-mail: rpa@seas.harvard.edu; shamim@seas.harvard.edu).

L. Mayaud is with the Institute of Biomedical Engineering, University of Oxford, Oxford, OX1 3PJ, UK, and also with Hôpital Raymond Poincaré, Garches 92380, France (e-mail: louis.mayaud@eng.ox.ac.uk).

A. Malhotra is with the Brigham and Women’s Hospital & Harvard Medical School, Boston, MA 02115 USA, and also with the UC San Diego Division of Pulmonary & Critical Care Medicine, La Jolla, CA 92037 USA (e-mail: amalhotra1@partners.org).

Color versions of one or more of the figures in this paper are available online at <http://ieeexplore.ieee.org>.

Digital Object Identifier 10.1109/JBHI.2014.2330827

## I. INTRODUCTION

MODERN clinical data acquisition systems are capable of continuously monitoring and storing measurements of patient vital signs, such as heart rate (HR) and blood pressure (BP), over multiple days of hospitalization [1]. Despite this continuous feed of data, commonly used acuity scores, such as APACHE and SAPS [2]–[5], are based on snap-shot values of these vital signs, typically the worst values during a 24-h period. However, physiologic systems generate complex dynamics in their output signals that reflect the state of the underlying control systems [6]–[8]. The objective of this study is to consider an approach to the analysis of critical care bedside monitoring that is based on the dynamical behaviors of vital sign time series.

The time series of vital signs (e.g., HR, BP) are multidimensional, high resolution (from once a second to once a minute), highly coupled due to presence of physiological feedback loops within the body [8], and remarkably nonstationary as a result of internally and externally induced changes in the state of the underlying control systems. For instance, time series of BP can exhibit oscillations on the order of seconds (e.g., due to the variations in sympathovagal balance), to minutes (e.g., as a consequence of fever, blood loss, or behavioral factors), to hours (e.g., due to humoral variations, sleep-wake cycle, or circadian effects) [9], [10]. A growing body of the literature is pointing to the clinical utility of vital signs time series dynamics to inform prognosis [11]–[17], and to provide early predictors of potentially life-threatening conditions in the intensive care unit (ICU) [18].

Techniques for modeling and analysis of cardiovascular and respiratory time series can be broadly classified into linear mechanistic models [19], [20] and nonlinear descriptive indices [6], [7], [21]. The linear techniques commonly used (often based on variants of autoregressive modeling) have the advantage of revealing the individual relationships among the observed variables (e.g., the noninvasive measures of baroreflex gain describes the relationship between HR and BP, excluding the possible influence of respiration). On the other hand, nonlinear indices of complexity are capable of capturing a richer set of dynamical behaviors, with less emphasis on physiological interpretability in terms of specific underlying mechanisms.

In this paper, we assume that although the underlying dynamical system may be nonlinear and nonstationary, and the stochastic noise components can be non-Gaussian, the dynamics can be approximated by a mixture of linear dynamical systems. Each such linear “dynamic” (or mode) is a time-dependent rule that describes how the future state of the system evolves from its

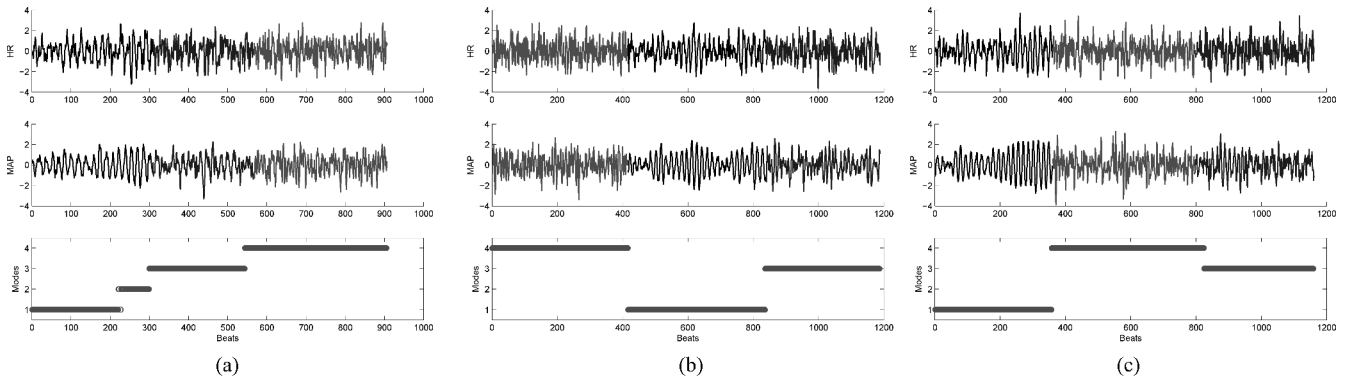


Fig. 1. Simulation study of the cardiovascular system. Three examples (out of the ten simulated time series) of HR and BP (after filtering) are shown in panels a, b, and c. In each case, the actual dynamics are color coded. The horizontal red lines show the inferred segmentation. The algorithm consistently assigned modes 4 and 3 to the dynamics color coded as red and blue, respectively, across all the simulated time series. The black dynamics are represented by modes 1 and 2. (a) Simulated subject 1, (b) Simulated subject 2, and (c) Simulated subject 3.

82 current state, centered around a given system equilibrium point.  
 83 Therefore, an ideal algorithm would be able to identify time series  
 84 segments that follow a “similar” dynamic, and would switch  
 85 to a different mode upon a change in the state of the underlying  
 86 system.

87 To formalize these objectives, we employed a switching vec-  
 88 tor autoregressive (SVAR) framework [22], [23]. Given a collec-  
 89 tion of time series from a cohort, the proposed SVAR framework  
 90 allows for simultaneous learning of the underlying dynamic be-  
 91 haviors or modes, and segmentation of the time series in terms  
 92 of the most likely dynamic describing the time series evolution  
 93 at any given point in time. The proposed framework enables  
 94 characterization of patients in terms of the dynamical modes  
 95 (e.g., the average time spent within the different modes), and  
 96 can potentially be used to capture changes in the underlying  
 97 cardiovascular control systems of human subjects in response  
 98 to internal (such as onset of infection) and external perturba-  
 99 tions (such as postural changes). Furthermore, we hypothesize  
 100 that when applied to vital sign time series of patients in a criti-  
 101 cal care setting, the proposed technique can be used to discover  
 102 dynamical modes with prognostic values for predicting clinical  
 103 outcomes of interests.

104 A preliminary version of this study was presented at the 34th  
 105 Annual International Conference of the IEEE Engineering in  
 106 Medicine and Biology Society (EMBC ’12) [14]. Here, we ex-  
 107 tend on our previous work to include a series of validation  
 108 studies, and a more comprehensive assessment of the utility of  
 109 the time series dynamics within the ICU.

110 The rest of this paper is organized as follows. We validated  
 111 the proposed technique using HR and BP time series from a  
 112 simulation dataset, and a human laboratory study of subjects  
 113 undergoing a tilt-table test, where the timing of the occurrence  
 114 of the different dynamics and the sharing of the dynamics across  
 115 multiple time series/subjects were known *a priori*. To test the  
 116 prognostic value of the discovered vital sign dynamics, we ap-  
 117 plied the proposed approach to the HR and BP dynamics of an  
 118 ICU cohort from the MIMIC II database [1] during the first 24 h  
 119 of their ICU stays, and tested whether cardiovascular dynamics  
 120 during the first 24 h of ICU admission are predictive of survival

and mortality after adjusting for the existing acuity scores, such  
 as SAPS-I and APACHE.

## II. MATERIALS AND METHODS

This section describes the utilized datasets, as well as the  
 proposed technique for discovery of shared dynamics among  
 patients, and assessment of risks and outcomes.

### A. Datasets

1) *Cardiovascular Simulation*: We simulated a cardiovas-  
 cular control system with bivariate time series of HR and BP.  
 The model is based on a delay recruitment model of HR and  
 BP regulation, as described in Fowler and McGuinness [24],  
 and McSharry *et al.* [25]. The model included a coupled system  
 of nonlinear delayed differential equations, controlling HR and  
 BP, with respiration as an exogenous input. We simulated ten  
 different multivariate time series of HR and mean arterial BP,  
 each including three different dynamics that become dominant  
 in a random order and last for a variable length of time. The  
 three dynamics (color-coded as red, blue, and black, respec-  
 tively, in Fig. 1) approximate aging-related autonomic changes;  
 a progressive reduction in parasympathetic gain (from 0.40 to  
 0.13 to 0.07 in normalized units; see [24]) and an increase in  
 sympathetic delay (from 3 to 5 s). To be consistent, we used the  
 same preprocessing step as the tilt-table experiment to remove  
 the steady-state baseline and any oscillation in the time series  
 slower than 100 beats/cycle (see below for details).

2) *Tilt-Table Experiment*: Time series of HR and BP were  
 acquired from ten healthy subjects (five males, five females)  
 undergoing a tilt-table test of orthostatic tolerance [26], [27].  
 The mean age was  $28.7 \pm 1.2$  years. The details of the pro-  
 tocol are described by Heldt in [27]. Briefly, subjects were  
 placed in a supine position. Tilting was performed from the  
 horizontal position to the vertical position and back to supine.  
 The study was approved by MIT’s Committee on the Use of  
 Humans as Experimental Subjects and the Advisory Board  
 of the MIT-MGH General Clinical Research Center [27].  
 Volunteers gave written, informed consent prior to participation

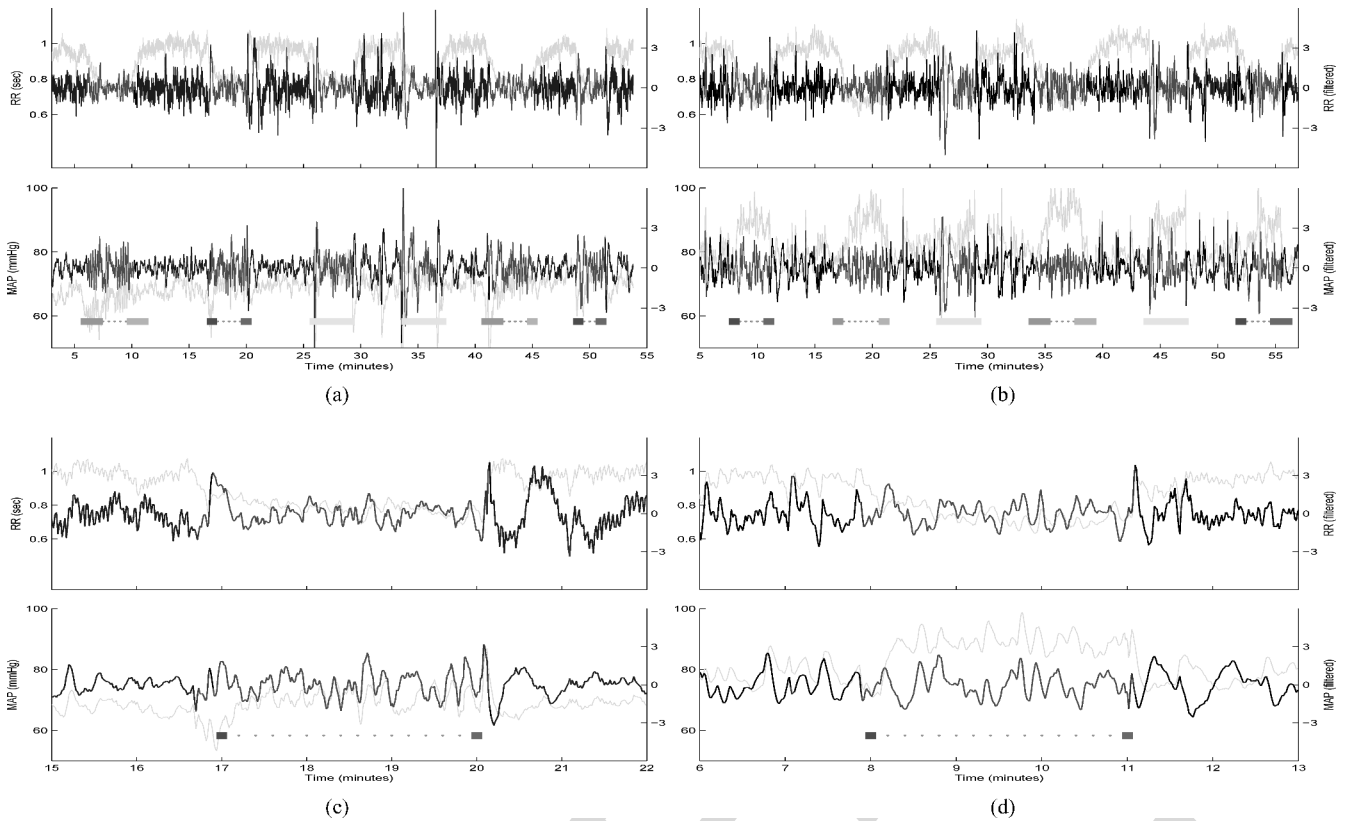


Fig. 2. Tilt-table study modeled using four dynamic modes—1 (Blue), 2 (Red), 3 (Black), and 4 (Purple). Two examples out of the ten recordings of HR and BP from the tilt-table experiment are shown in Panels a and b. Panels c and d show a zoomed in 7-min recording of HR and BP, while the subjects transition to/from supine to nonsupine positions after a fast tilt procedure. Actual values are in gray (*Y*-axis on left) and filtered values (*Y*-axis on right) are color coded based on the inferred dynamical modes. Note that Subjects 1 and 2 shared the same inferred nonsupine dynamics (in red); the algorithm consistently assigns the red mode to the nonsupine position for both subjects. The supine position for Subjects 1 and 2 are captured by the modes in blue and black, respectively. The purple mode seems to capture the high-frequency noise components of the time series. In each case, annotations for the actual tilt procedures performed are plotted as horizontal bars on the bottom of each figure and are color coded (green to cyan: slow tilt up and down to supine; red to pink: rapid tilt up and down to supine; yellow: stand up and back to supine). (a) Tilt Subject 1, (b) Tilt Subject 2, (c) Tilt Subject 1 (zoomed in), and (d) Tilt Subject 2 (zoomed in).

157 in the study. Since we were interested in the dynamics of inter-  
 158 action between HR and BP in the frequency range pertinent  
 159 to sympathetic and parasympathetic regulation [28], time series  
 160 of HR and BP were high-pass filtered to remove the steady-  
 161 state baseline and any oscillation in the time series slower than  
 162 100 beats/cycle. This filtering was done using a seventh-order  
 163 Butterworth digital filter with a cutoff frequency of 0.01 cycles/beat.  
 164 Example time series from before and after filtering  
 165 are shown in Fig. 2.

166 3) *MIMIC II Dataset*: The MIMIC II database [1], publicly  
 167 available via PhysioNet [29], includes clinical (laboratory  
 168 values, IV medications, etc.) and physiological data (HR,  
 169 BP, oxygen saturation, etc.) collected from the bedside moni-  
 170 tors (Component Monitoring System Intellivue MP-70; Philips  
 171 Healthcare, Andover, MA, USA) in ICUs of the Beth Israel  
 172 Deaconess Medical Center (BIDMC) in Boston. The MIMIC  
 173 II waveform database (version 2) includes approximately 4000  
 174 records of high-resolution physiological waveforms of adult  
 175 ICU patients with associated minute-by-minute (averages of  
 176 the calculated numerics during the previous minute) vital sign  
 177 trends. Data collection for the MIMIC II database was approved  
 178 by the Institutional Review Boards of BIDMC and the Mas-  
 179 sachusetts Institute of Technology (Cambridge, MA, USA). In-

dividual patient consent was waived because the study did not  
 impact clinical care and protected health information was de-  
 identified.

This study includes adult patients from the MIMIC II wave-  
 form database with at least 8 h of continuous minute-by-minute  
 HR and invasive arterial BP trends during the first 24 h in  
 the ICU. Patients with more than 15% of missing or invalid  
 samples (i.e., outside physiologically plausible bounds of 20 to  
 200 mmHg for mean pressures) were excluded from this study,  
 as were patients with missing SAPS I and APACHE scores.  
 The dataset contains over 9000 h of minute-by-minute HR and  
 invasive mean arterial BP measurements (over 20 h per patient  
 on average) from 453 adult patients collected during the first  
 24 h in the ICU. HR and BP time series were detrended. Gaus-  
 sian noise was used to fill in the missing or invalid values. The  
 median age of this cohort was 69 with an interquartile range of  
 (57, 79). 59% of the patients were male. Approximately 15%  
 (67 out of 453) of patients in this cohort died in the hospital;  
 28-day mortality of this cohort was approximately 19% (85 out  
 of 453). Distributions of the 453 patients in care units are 21%  
 coronary care unit (CCU), 42% cardiac surgery recovery unit  
 (CSRU), 26% medical intensive care unit (MICU), and 12%  
 surgical intensive care unit.

## 203 B. SVAR Modeling of Cohort Time Series

204 Our approach to discovery of shared dynamics among pa-  
 205 tients is based on the SVAR model [22]. For the  $n$ th patient  
 206 ( $n = 1 \dots N$ ), let  $y_t^{(n)}$  be a  $M \times 1$  vector of observed values of  
 207 the vital signs at time  $t$  ( $t = 1 \dots T^{(n)}$ ). We assume that there  
 208 exists a library of  $K$  possible dynamics or *modes*; a set of multi-  
 209 variate autoregressive model coefficient matrices  $\{A_p^{(k)}\}_{k=1}^K$  of  
 210 size  $M \times M$ , with maximal time lag  $p = 1 \dots P$ , and the cor-  
 211 responding noise covariances  $\{Q^{(k)}\}_{k=1}^K$ . Let  $s_t$  be a switching  
 212 variable, indicating the active dynamic mode at time  $t$ , and  
 213 evolving according to a Markovian dynamic with initial distri-  
 214 bution  $\pi^{(n)}$  and a  $K \times K$  transition matrix  $Z$ . Following these  
 215 definitions, an SVAR model for the  $n$ th patient is defined as

$$y_t^{(n)} = \sum_{p=1}^P A_p^{(s_t^{(n)})} y_{t-p}^{(n)} + w^{(s_t^{(n)})} \quad (1)$$

216 where the fluctuation term  $w^{(s_t^{(n)})}$  is assumed Gaussian dis-  
 217 tributed with covariance  $Q^{(s_t^{(n)})}$ . A collection of related time se-  
 218 ries can be modeled as switching between these dynamic behav-  
 219 iors which describe a locally coherent linear model that persists  
 220 over a segment of time. However, in practice, we neither know  
 221 the set of switching variables (i.e., segmentation of the time  
 222 series) nor the modes. In this study, we perform expectation-  
 223 maximization (EM) to find the maximum-likelihood set of  
 224 model parameters, as well as a factored estimate of the posterior  
 225 distribution over the latent switching variables. A comprehen-  
 226 sive treatment of the EM algorithm for SVAR is presented by  
 227 Murphy (1998) [22]. Briefly, EM is a two-pass iterative algo-  
 228 rithm: 1) in the expectation (E) step, we obtain the expected val-  
 229 ues of the latent switching variables  $\{s_t^{(n)}\}_{t=1}^T$  using a forward-  
 230 backward algorithm [22], and 2) in the maximization (M) step,  
 231 we update all the model parameters  $\{A_p^{(k)}\}, \{Q^{(k)}\}$ , the Markov  
 232 dynamics  $Z$ , and the initial conditions  $\pi^{(n)}$  that maximize the  
 233 expected complete data log likelihood. In our implementation  
 234 of the EM algorithm, we achieve *shared* dynamics by pooling  
 235 together all subjects' inferred variables in the M step. Iteration  
 236 through several steps of the EM algorithm results in learning a  
 237 set of  $K$  shared modes and a global transition matrix  $Z$  for all  
 238 the patients.

239 For the simulated and the tilt datasets, we modeled the beat-  
 240 by-beat HR/BP time series as a switching AR(5) process to  
 241 model most of the parasympathetic responses and at least some  
 242 of the sympathetic effects, without introducing an unduly com-  
 243 plex model. Minute-by-minute BP time series from MIMIC II  
 244 were modeled as a switching AR(3) process to capture a real os-  
 245 cillation and a possible trend per mode. The number of dynamic  
 246 modes ( $K = 20$ ) was determined using the Bayesian information  
 247 criterion (BIC) [30]. Briefly, we computed the BIC scores from  
 248 switching-VAR models using 5 to 30 modes. Results presented  
 249 were based on the model with the minimum BIC scores (20  
 250 modes).

251 1) *Parallel Computation for Scalable Learning*: One of the  
 252 advantages of the proposed technique is its scalability to hun-  
 253 dreds or thousands of patients, due to the parallel implementa-  
 254 tion of the inference step of the SVAR learning algorithm via

EM [22]. This parallelization strategy is effective since the ma-  
 255 jority of the computational cost of the SVAR training is in run-  
 256 ning the forward-backward algorithm, which can be done in  
 257 parallel for each patient time series. We used MATLAB's par-  
 258 allel computation toolbox in association with 120 nodes on our  
 259 computer cluster to perform a tenfold cross-validated study (12  
 260 cores per fold). Ten SVAR models were learned on the training  
 261 set of each of the folds, followed by mapping the corresponding  
 262 mode proportions to outcomes (e.g., hospital mortality) using  
 263 logistic regression. Next, mode assignments of time series in the  
 264 test set of each fold were inferred based on the modes learned  
 265 from the corresponding training set (by running only the infer-  
 266 ence), and the regression weights from the training fold were  
 267 used to predict outcomes. 268

## 269 C. Evaluation Methods and Statistical Analysis

Let us define a *mode proportion*  $MP_k^{(n)}$  as the proportion  
 270 of time the  $n$ th patient spends within the  $k$ th mode. Given the  
 271 maximum expected log-likelihood estimates of the switching  
 272 variables  $s_t$  from the EM algorithm, we have 273

$$MP_k^{(n)} = \frac{1}{T^{(n)}} \sum_{t=1}^{T^{(n)}} \text{Prob}(s_t^{(n)} = k). \quad (2)$$

For classification and prediction purposes, we characterize each  
 274 time series with its corresponding mode proportion (a  $1 \times K$   
 275 feature-vector), and use a logistic regression classifier to make  
 276 predictions about the outcome variables of interest. For illus-  
 277 tration of the algorithm's segmentation performance, each time  
 278 series sample is assigned to the dynamic mode with the maxi-  
 279 mum posterior probability. 280

281 1) *Time Series Classification and Outcome Prediction*: For  
 282 the simulated and the tilt-table experiment, we used the mode  
 283 proportions within each segment (e.g., supine versus nonsupine)  
 284 as inputs to a logistic regression classifier, and report the clas-  
 285 sification performance in discriminating between 1) the three  
 286 different dynamics (corresponding to different aging-related au-  
 287 tonomic changes) in the simulated dataset, and 2) two different  
 288 postural positions (supine versus nonsupine) in the tilt dataset.

289 To assess the predictive power of the dynamical modes, we  
 290 performed a tenfold cross-validation study. Ten SVAR models  
 291 were learned on the training set of each of the folds, followed  
 292 by mapping the corresponding mode proportions to outcomes  
 293 (e.g., hospital mortality) using logistic regression. Next, mode  
 294 assignments of time series in the test set of each fold was in-  
 295 ferred based on the modes learned from the corresponding train-  
 296 ing set (by running only the inference or the E-step), and the  
 297 regression weights from the training fold was used to predict  
 298 outcomes. We compared the mortality prediction performance  
 299 of our approach using the mode proportion from the top ten  
 300 most common dynamic modes with the existing acuity metrics,  
 301 SAPS I [2], APACHE III [4], and APACHE IV [5]. Comparison  
 302 of AUCs was based on the method described in [31].

303 2) *MIMIC Association Analysis*: We used univariate and  
 304 multivariate logistic regressions to examine the associations be-  
 305 tween dynamic mode proportions and hospital mortality. We

TABLE I  
PERFORMANCE OF MORTALITY PREDICTORS

	Hosp. Mortality (AUC)	28-Days Mortality (AUC)
HR <sub>dyn</sub>	0.59 (0.54, 0.68)	0.61 (0.51, 0.67)
BP <sub>dyn</sub> /HR <sub>dyn</sub>	0.64 (0.61, 0.71)	0.65 (0.64, 0.68)
BP <sub>dyn</sub>	<b>0.70 (0.67, 0.77)</b>	<b>0.66 (0.61, 0.73)</b>
SAPS I	0.65 (0.59, 0.71)	0.64 (0.56, 0.70)
BP <sub>dyn</sub> + SAPS I	<b>0.77 (0.69, 0.82)</b>	<b>0.71 (0.69, 0.79)</b>
APACHE III	0.80 (0.70, 0.84)	0.79 (0.65, 0.84)
BP <sub>dyn</sub> + APACHE III	<b>0.84 (0.79, 0.88)</b>	<b>0.79 (0.76, 0.86)</b>
APACHE IV	0.82 (0.77, 0.85)	0.83 (0.74, 0.86)
BP <sub>dyn</sub> + APACHE IV	<b>0.85 (0.80, 0.87)</b>	<b>0.82 (0.81, 0.88)</b>

built a separate multivariate logistic regression model for each of the discovered dynamic modes, with the mode proportion as the primary predictive variable, and APACHE IV as a covariate. For each mode, we reported its  $p$  value, odds ratio (OR, with 95% confidence interval), and adjusted OR (after including APACHE IV as a covariate). The Hosmer–Lemeshow  $p$  values (HL  $p$  values) were reported to assess the model fit. The odds ratios were per 10% increase in the mode proportion. Two-sided  $p$  values less than 0.05 were considered statistically significant. The analysis was performed to quantify the mortality risk associated with each dynamic mode; modes with significant ( $p < 0.05$ ) associations with mortality were established as either *low-risk* (OR  $< 1$ ), or *high-risk* (OR  $> 1$ ) dynamics depending on their odds ratios. Dynamic modes without statistically significant associations with mortality were *neutral* modes. The test of statistical significance was based on  $p$ -values after correcting for the false discovery rate (FDR) using the technique described in [32].

### III. RESULTS

#### A. Simulated Study

Fig. 1 shows two examples of simulated time series with the inferred segmentation. In all ten simulated cases, the algorithm was able to divide each time series into distinct segments corresponding to different underlying actual dynamics. The sharing of the dynamics is consistent across the different time series. Using the mode proportion from each segment for multilabel classification, the algorithm achieved classification accuracy of 100%.

#### B. Tilt-Table Experiment

Fig. 2 shows the segmentation results for two subjects. Note that the two subjects shared the same inferred nonsupine dynamics (in red); the algorithm consistently assigns the red mode to the nonsupine position for both subjects. The application of logistic regression with tenfold cross-validation yielded a median AUC of 1.00 with an interquartile range of (0.98, 1.00).

#### C. MIMIC II Database

1) *Mortality Prediction*: Table I evaluates the prognostic power of HR and BP dynamic features (HR<sub>dyn</sub> and BP<sub>dyn</sub>).

SAPS I, APACHE III, and APACHE IV are used as the baselines. Median AUCs (from tenfold cross validation) and the interquartile range are shown. Note that the BP dynamics outperformed both the HR and HR and BP combined dynamic features. Subsequent analyses focus on the predictive power of the BP dynamics in comparison to the baseline. For each baseline, we show the performance from the baseline alone, and the combined approach (combining BP dynamics and the baseline).

The application of tenfold cross-validation demonstrated that dynamic features from BP alone achieved a median AUC of 0.70, comparable to 0.65 from SAPS I. In comparison, using standard deviation of the mean arterial BP resulted in a median AUC (IQR) of 0.55 (0.43, 0.63).

Combining dynamic BP features with SAPS I resulted in an improved prediction power both in hospital mortality prediction ( $p = 0.005$ ) and 28-day mortality prediction ( $p = 0.002$ ). Combining dynamic features with APACHE III significantly outperformed APACHE III alone ( $p = 0.045$ ) with an improvement in median AUC from 0.80 to 0.84 in hospital mortality prediction. These results indicate that the dynamic features from vital signs contain complementary information to the SAPS I and APACHE III scores.

State-of-the-art risk score APACHE IV achieved better prediction performance than the BP dynamic features alone ( $p = 0.008$ ). Adding BP dynamics to APACHE IV yielded a slight performance improvement from a median AUC of 0.82 to 0.85, however, the performance gain was not statistically significant.

2) *Association Analysis*: Table II presents logistical regression analyses to test the associations between the proportion of time patients spent in each of the top ten most common BP dynamics and hospital mortality. See Fig. 3 for illustrations of these dynamic modes. Dynamic modes were numbered based on their prevalence across the entire cohort (i.e., mode 1 is the most common dynamic mode). Our results indicate that six of the modes had significant associations (after FDR correction) with hospital mortality. Specifically, two dynamic modes (modes 3 and 5) were significant “high-risk” modes ( $p < 0.001$ ,  $p < 0.001$ ) in which increased proportions of time in these modes were associated with higher hospital mortality with odds ratios 1.81 (1.41, 2.32), 1.36 (1.15, 1.61) respectively.

Dynamic modes 1, 9, 7, and 2 were “low-risk” modes in which increasing proportions of time in these modes were significantly associated with a decreased risk of hospital mortality, with odds ratios less than one. Table II lists the AR coefficients and covariances of the two high-risk and four low-risk dynamic modes, as well as their respective associations with hospital mortality. Note that the high-risk modes appear to correspond to less variability in their dynamics.

For the multivariate analysis (see the right panel in Table II), each row is a separate multivariate model, in which the mode proportion for a given target mode is the primary predictive variable, and APACHE IV is added as a control variable in the multivariate model. Results from multivariate logistic regression indicate that two of the modes (modes 3 and 9) remain significant predictors of patients’ outcome even after adjustment for APACHE IV scores ( $p = 0.001$ ,  $p = 0.006$ ), indicating that the proportion of time patients spent in these two dynamic modes



TABLE II  
ASSOCIATIONS OF BP DYNAMIC MODES AND HOSPITAL MORTALITY

Mode	AR Coef	Cov.	P-Val	OR(95%CI)	Adjusted P-Val	Adjusted OR(95%CI)	HL PVAL
3	(0.66, 0.22, 0.12)	0.58	< <b>0.001</b>	1.81 (1.41, 2.32)	<b>0.001</b>	1.60 (1.21 2.11)	0.64
5	(1.00, 0.00, -0.00)	0.22	< <b>0.001</b>	1.36 (1.15, 1.61)	0.426	1.08 (0.89 1.32)	0.16
1	(0.66, 0.16, 0.17)	2.69	<b>0.002</b>	0.59 (0.42, 0.82)	0.489	0.88 (0.62 1.26)	0.54
9	(1.50, -0.65, 0.06)	7.26	<b>0.002</b>	0.25 (0.10, 0.62)	<b>0.006</b>	0.26 (0.10 0.68)	0.80
7	(1.00, -0.01, -0.00)	3.46	<b>0.003</b>	0.30 (0.13, 0.67)	0.124	0.54 (0.25 1.18)	0.58
2	(0.79, 0.05, 0.12)	8.81	<b>0.005</b>	0.65 (0.48, 0.88)	0.265	0.84 (0.62 1.14)	0.69
10	(1.05, -0.01, -0.02)	0.71	0.032	2.95 (1.10, 7.94)	0.791	1.18 (0.36 3.88)	0.07
8	(0.44, 0.30, 0.24)	1.27	0.373	1.18 (0.82, 1.69)	0.318	1.22 (0.82 1.82)	0.22
4	(0.96, -0.01, 0.04)	1.31	0.417	0.81 (0.48, 1.36)	0.887	0.96 (0.53 1.72)	0.02
6	(0.92, -0.10, 0.07)	46.70	0.419	0.83 (0.53, 1.30)	0.658	0.90 (0.57 1.43)	0.08

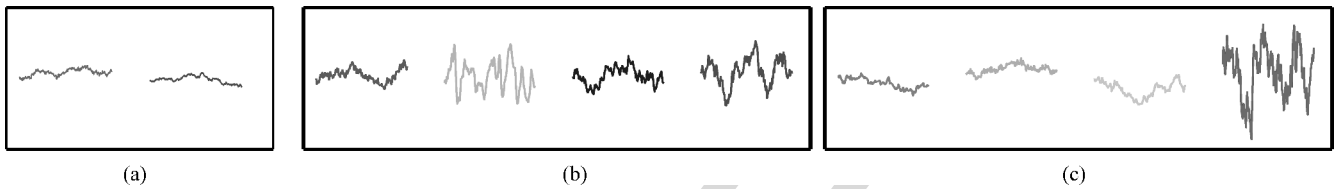


Fig. 3. Discovered dynamic modes of mean arterial BP of 453 patients during the first 24 h in the ICU. Figure shows the top ten most common dynamic modes, simulated using the AR coefficients from each dynamic mode. High-risk dynamic modes (from left to right): 3 (Magenta), 5 (Red). Low-risk dynamic modes: 1 (Violet), 9 (Cyan), 7 (Blue), and 2 (Green). Neutral dynamic modes: 10 (Brown), 8 (Orange), 4 (Light Green), 6 (Royal Blue). All modes were simulated and plotted with the same time duration (150 min) and amplitude scale. (a) High-risk modes, (b) Low-risk modes, and (c) Neutral modes.

401 during the first 24 h in the ICU are independent risk predictors  
402 of hospital mortality.

403 3) *Example Time Series of Patients With Estimated Mortality*  
404 *Risks Over Time:* Fig. 3 shows examples of low-risk and high-  
405 risk dynamical modes learned using the SVAR technique (see  
406 Table II for the odds-ratio associated with each mode). BP time  
407 series of four patients are presented in Fig. 4 panels (a) and  
408 (b). Hourly risk scores (dark green lines) were computed as the  
409 probability of death from the logistic function using a sliding  
410 window of 6 h to illustrate that these risk scores could be updated  
411 on a continuous basis for real-time monitoring purposes.

412 Panel (a) shows two of the patients with the highest risk scores  
413 (within the test set) at the end of the 24-h period; both patients  
414 died in the hospital. Panel (b) shows two patients with a decreas-  
415 ing trend in their risk scores during their first day in the ICU;  
416 both patients survived the hospital stay. All four patients were  
417 from the same test set, with mode assignment inferred based  
418 on dynamic modes learned from the corresponding training set.  
419 Note that as time progresses, patients in panel (a) tend to spend  
420 more time in the high-risk dynamic modes (mode 3 in magenta,  
421 mode 5 in red); their estimated mortality risks rise accordingly  
422 over time. In contrast, panel (b) patients show a decreasing trend  
423 in mortality risks as they transit to lower-risk dynamic modes  
424 over time.

#### 425 IV. DISCUSSION AND CONCLUSION

426 We presented a SVAR framework to systematically learn and  
427 identify dynamic behaviors from vital sign time series within a  
428 patient cohort. We demonstrated that the discovered dynamics  
429 may contain prognostic values and can be used for prediction

and tracking of a patient's propensity to survive a hospital stay, 430  
as well as their 28-days survival. Interestingly, the BP time 431  
series dynamics alone had a comparable performance to that of 432  
the SAPS I score which uses age and the most extreme values of 433  
13 variables, including systolic BP, HR, temperature, respiratory 434  
rate, urinary output, blood nitrogen, hematocrit, white blood cell 435  
count, serum glucose, serum potassium, serum sodium, serum 436  
bicarbonate, and Glasgow coma score. 437

438 Additionally, our results indicate that the BP dynamics may 438  
contain complimentary information to existing acuity metrics, 439  
which assess the health of multiple organ systems based on a vari- 440  
ety of physiological and lab variables. Specifically, combining 441  
the dynamics of BP time series and SAPS I or APACHE III pro- 442  
vided a more accurate assessment of patient survival/mortality 443  
in the hospital ( $p = 0.005$  and  $p = 0.045$ ) than using SAPS I 444  
and APACHE III alone. 445

446 Association analysis of individual dynamic mode and hospi- 446  
tal mortality revealed that two of the dynamic modes (modes 447  
3 and 9) remained significant predictors of patients' outcome 448  
even after adjusting for APACHE IV scores, indicating that the 449  
proportion of time patients spent in these two dynamic modes 450  
during the first 24 h in the ICU may contain additional, inde- 451  
pendent prognostic value beyond that in the APACHE IV acuity 452  
score. Future work remains to investigate the prognostic power 453  
of these discovered dynamic modes using a larger cohort. 454

455 The dynamic features can be calculated in an online man- 455  
ner without delay, and well before the end of the first 24 h of 456  
the ICU stay as is required for the standard risk scores. One 457  
possible online deployment strategy is to construct a library of 458  
dynamic modes on archived patient data, and assign each in- 459  
coming time series sample (or a sliding window of samples) 460

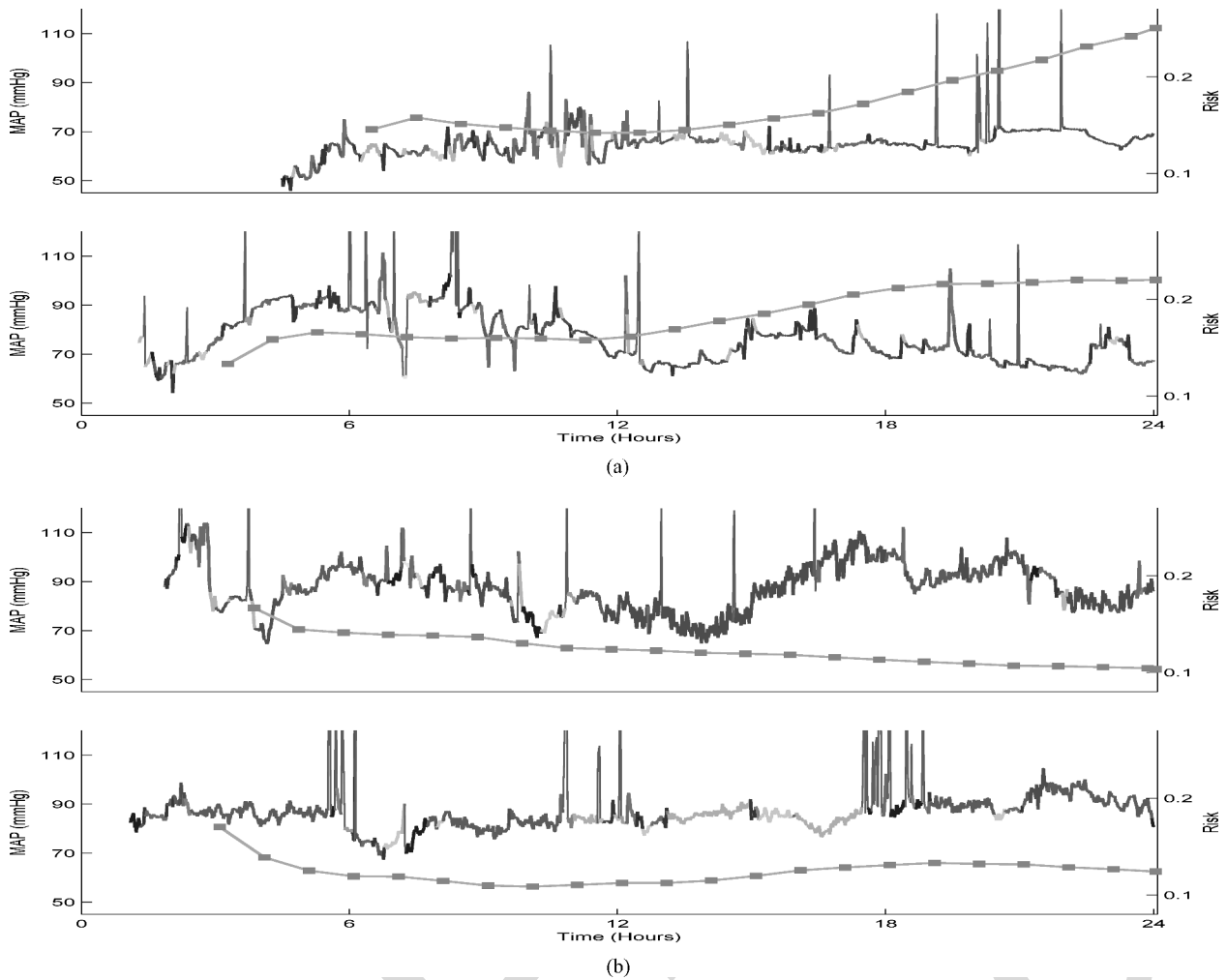


Fig. 4. Mortality risk scores and mean arterial BP of four patients during the first 24 h in the ICU. Samples are color coded by their mode assignment. Mortality risk scores, computed as the probability of death from the logistic regression, were based on mode proportions from a 6-h sliding window by stride of 1 h; estimated risks were plotted as dark green lines with scale indicated by  $y$ -axes on the right side of each graph. BP measurements plotted in original units (before detrending). All four patients were from the same test set, with dynamic modes and logistic regression parameters learned from the corresponding training set. (a) Patients with the highest ending risk scores at the end of the first day ICU stay. Patients were from MICU (top) and CCU (bottom). Both patients died in the hospital and (b) Patients with decreasing risk scores during their first day ICU stays. Patients were from CSRU (top) and CCU (bottom). Both patients survived the hospital stay.

461 to the most likely mode in the library (for instance, by  
 462 using the Viterbi algorithm [16], [22]). Recent studies suggest  
 463 that therapeutic interventions not only should aim at maintaining  
 464 the mean BP within an acceptable range, but also should direct  
 465 the patient's trajectory toward healthy dynamical regimes with  
 466 enhanced variability [10]. Thus, a real-time implementation of  
 467 the technique presented here may provide clinicians with a tool  
 468 for quantification of the effectiveness of such interventions in  
 469 the ICU.

470 We showed that changes in the dynamics of HR and BP, either  
 471 as a result of an altered underlying control system (aging-related  
 472 changes in the simulated data) or due to external perturbations  
 473 (positional changes in the tilt-table experiment), can be captured  
 474 in an automated fashion. Since the proposed framework is built  
 475 on the dynamical systems framework (which includes the class  
 476 of vector autoregressive models), the discovered modes can be  
 477 used to reveal the oscillations that are present within the indi-  
 478 vidual time series, and therefore can be used to extract useful

479 indices of HR and BP variability (assuming beat-to-beat time  
 480 series). Moreover, given beat-to-beat multivariate time-series of  
 481 vital-signs, one may use the learned dynamics to derive the di-  
 482 rectional transfer functions of the system [8] (e.g., baroreflex  
 483 control of HR and BP).

484 Association analysis using the minute-by-minute MIMIC-  
 485 II BP time series revealed that the high-risk modes often  
 486 correspond to less variable dynamical patterns. It is interest-  
 487 ing to note that such low-frequency variability, observed at  
 488 the minute-to-minute scale, is associated with an enhanced  
 489 chance of survival, corresponding well to the existing HR/BP  
 490 variability literature using beat-by-beat vital sign time series  
 491 [10], [12], [13], [33]. The working hypothesis of our ongo-  
 492 ing research is that the observed dynamical patterns are due  
 493 to patients' underlying physiology, patient-specific response  
 494 to clinical interventions, and measurement artifacts. Future  
 495 developments of machine-learning techniques should aim at  
 496 combining time series dynamics with contextual information

497 pertaining to clinical intervention (administration of fluids,  
498 pressors, and titration of medications) to further investigate  
499 the clinical and physiological interpretation of the discovered  
500 modes.

501 The SVAR framework allows for defining a notion of “sim-  
502 ilarity” among multivariate physiological time series based on  
503 their underlying shared dynamics. Therefore, one may consider  
504 two subjects to be similar if their underlying vital signs time se-  
505 ries exhibit similar dynamics in response to external (e.g., tilting  
506 of body) or internal perturbations (e.g., onset of blood infection).  
507 This approach provides an improvement over time series sim-  
508 ilarity measures based on trend-detection [34], wavelet-based  
509 symbolic representations [35], or Gaussian Mixture modeling  
510 [36] due to its compact representation and sharing of the model  
511 parameters within and across time series. Prior work using a fac-  
512 torial switching linear dynamical systems for patient monitoring  
513 [37] focused on detection of events associated with artifactual  
514 measurements and pathological states. Our study, in contrast,  
515 jointly models multiple time series across a large patient cohort  
516 to identify phenotypic dynamical patterns for patient outcome  
517 prediction.

518 Although we used mortality as our target outcome, there are  
519 many other physiological events of significant interest, includ-  
520 ing timely and successful discontinuation of procedures such  
521 as hemodialysis [38] or mechanical ventilation [39], as well as  
522 prediction of potentially life-threatening clinical events such as  
523 onset of severe sepsis and hypotension [13]. Other short- and  
524 long-term outcomes such as probability of readmission to hospi-  
525 tal and long-term cognitive impairment beyond ICU [40] also  
526 play an important role in closing the gap between the critical care  
527 medicine, primary care doctors, and other healthcare providers.

528 Current and ongoing work involve combining the switching  
529 linear dynamical system framework with all available clinical  
530 data, including lab tests, medication records, and nursing notes  
531 [41] to devise a comprehensive risk score, capable of integrating  
532 clinical data of diverse modality over long temporal stretches  
533 (order of hours to days). This will allow us to investigate whether  
534 continuous patient monitoring based on vital signs dynamics,  
535 and other types of sequential data, can alert clinicians to deter-  
536 iorating patient conditions at an earlier stage than the existing  
537 acuity scores, and result in improved patient care and outcome  
538 both within ICU and after hospital discharge. Such analysis is  
539 likely to provide some insight into the promise of large-scale  
540 critical care databases for the future of medicine.

#### 541 ACKNOWLEDGMENT

542 The authors would like to thank Dr. T. Heldt (Harvard-MIT  
543 Division of Health Sciences and Technology) for kindly provid-  
544 ing the tilt-table data analyzed in this study.

#### 545 REFERENCES

546 [1] M. Saeed, M. Villarroel, A. T. Reisner, G. Clifford, L. H. Lehman,  
547 G. Moody, T. Heldt, T. H. Kyaw, B. Moody, and R. G. Mark, “Multi-  
548 parameter intelligent monitoring in intensive care (MIMIC II): A public-  
549 access intensive care unit database,” *Crit. Care Med.*, vol. 39, no. 5,  
550 pp. 952–960, May 2011.

[2] J. R. Le Gall, P. Loirat, A. Alperovitch, P. Glaser, C. Granthil, D. Mathieu, 551  
P. Mercier, R. Thomas, and D. Villers, “A simplified acute physiology 552  
score for ICU patients,” *Crit. Care Med.*, vol. 12, no. 11, pp. 975–977, 553  
Nov. 1984.

[3] J. R. Le Gall, S. Lemeshow, and F. Saulnier, “A new simplified acute 554  
physiology score (SAPS II) based on a European/North American 555  
multicenter study,” *J. Amer. Med. Assoc.*, vol. 270, no. 24, pp. 2957–2963, 556  
1993.

[4] W. A. Knaus, D. P. Wagner, E. A. Draper, J. E. Zimmerman, M. Bergner, 557  
P. G. Bastos, C. A. Sirio, D. J. Murphy, T. Lotring, A. Damiano, 558  
and F. Harrell, “The APACHE III prognostic system,” *Chest*, vol. 100, 559  
pp. 1619–1636, 1991.

[5] J. E. Zimmerman, A. A. Kramer, D. S. McNair, and F. M. Malila, “Acute 560  
physiology and chronic health evaluation (APACHE) IV: Hospital mor- 561  
tality assessment for today’s critically ill patients,” *Crit. Care Med.*, 562  
vol. 34, no. 5, pp. 1297–1310, May 2006.

[6] P. C. Ivanov, L. A. Amaral, A. L. Goldberger, S. Havlin, M. G. Rosenblum, 563  
Z. R. Struzik, and H. E. Stanley, “Multifractality in human heartbeat 564  
dynamics,” *Nature*, vol. 399, pp. 461–465, 1999.

[7] M. Costa, A. L. Goldberger, and C. K. Peng, “Multiscale entropy analysis 565  
of complex physiologic time series,” *Phys. Rev. Lett.*, vol. 89, no. 6, 566  
p. 068102, 2002.

[8] S. Nemati, B. A. Edwards, S. A. Sands, P. J. Berger, A. Wellman, G. C. 567  
Verghese, A. Malhotra, and J. P. Butler, “Model-based characterization 568  
of ventilatory stability using spontaneous breathing,” *J. Appl. Physiol.*, 569  
vol. 111, no. 1, pp. 55–67, 2011.

[9] G. Mancia, “Short-and long-term blood pressure variability present and 570  
future,” *Hypertension*, vol. 60, no. 2, pp. 512–517, 2012.

[10] G. Parati, J. E. Ochoa, C. Lombardi, and G. Bilo, “Assessment and 571  
management of blood-pressure variability,” *Nat. Rev. Cardiol.*, vol. 10, 572  
pp. 143–155, 2013.

[11] T. G. Buchman, “Nonlinear dynamics, complex systems, and the patho- 573  
biology of critical illness,” *Current Opinion Crit. Care*, vol. 10, no. 5, 574  
pp. 378–382, 2004.

[12] S. Saria, A. K. Rajani, J. Gould, D. Koller, and A. Penn, “Integration 575  
of early physiological responses predicts later illness severity in preterm 576  
infants,” *Sci. Translational Med.*, vol. 2, no. 48, pp. 48–65, 2010.

[13] J. R. Moorman, J. B. Delos, A. A. Flower, H. Cao, B. P. Kovatchev, 577  
J. S. Richman, and D. E. Lake, “Cardiovascular oscillations at the bed- 578  
side: Early diagnosis of neonatal sepsis using heart rate characteris- 579  
tics monitoring,” *Physiol. Meas.*, vol. 32, no. 11, pp. 1821–1832, Nov. 580  
2011.

[14] L. H. Lehman, S. Nemati, R. P. Adams, and R. G. Mark, “Discovering 581  
shared dynamics in physiological signals: Application to patient monitor- 582  
ing in ICU,” in *Proc. IEEE Eng. Med. Biol. Soc.*, 2012, pp. 5939–5942. 583

[15] J. Wiens, E. Horvitz, and J. V. Guttag, “Patient risk stratification for 584  
hospital-associated C. diff as a time-series classification task,” in *Adv. 585  
Neural Inf. Process. Syst.*, 2012, pp. 476–484.

[16] L. H. Lehman, S. Nemati, R. P. Adams, G. Moody, A. Malhotra, and R. G. 586  
Mark, “Tracking progression of patient state of health in critical care using 587  
inferred shared dynamics in physiological time series,” in *Proc. IEEE Eng. 588  
Med. Biol. Soc.*, 2013, pp. 7072–7075.

[17] L. Mayaud, P. S. Lai, G. D. Clifford, L. Tarassenko, L. A. Celi, and 589  
D. Annane, “Dynamic data during hypotensive episode improves mortality 590  
predictions among patients with sepsis and hypotension,” *Crit. Care Med.*, 591  
vol. 41, no. 4, pp. 954–962, 2013.

[18] M. Blount, M. R. Ebling, J. M. Eklund, A. G. James, C. McGregor, 592  
N. Percival, K. P. Smith, and D. Sow, “Real-time analysis for intensive 593  
care: Development and deployment of the artemis analytic system,” *IEEE 594  
Eng. Med. Biol. Mag.*, vol. 29, no. 2, pp. 110–118, Mar–Apr. 2010.

[19] A. Porta, T. Bassani, V. Bari, E. Tobaldini, A. C. M. Takahashi, A. M. 595  
Catai, and N. Montano, “Model-based assessment of baroreflex and cardiopulmonary 596  
couplings during graded head-up tilt,” *Comput. Biol. Med.*, 597  
vol. 42, no. 3, pp. 298–305, Mar. 2012.

[20] R. Mukkamala, J. M. Mathias, T. J. Mullen, R. J. Cohen, and R. Freeman, 598  
“System identification of closed-loop cardiovascular control mechanisms: 599  
Diabetic autonomic neuropathy,” *Amer. J. Physiol.-Regulatory, Integrative 600  
Comparative Physiol.*, vol. 276, no. 3, pp. R905–R912, 1999.

[21] I. Korhonen, “Multivariate closed-loop model for analysis of cardiovas- 601  
cular dynamics,” *Methods Inf. Med.*, vol. 36, pp. 264–267, 1997.

[22] K. P. Murphy, “Switching Kalman filter,” Compaq Cambridge Res. Lab., 602  
Cambridge, CA, USA, Tech. Rep. 98-10, 1998.

[23] S. Nemati, L. H. Lehman, R. P. Adams, and A. Malhotra, “Discovering 603  
shared cardiovascular dynamics within a patient cohort,” in *Proc. IEEE 604  
Eng. Med. Biol. Soc.*, 2012, pp. 6526–6529. 605  
606  
607  
608  
609  
610  
611  
612  
613  
614  
615  
616  
617  
618  
619  
620  
621  
622  
623  
624  
625

- 626 [24] A. Fowler and M. McGuinness, "A delay recruitment model of the cardiovascular control system," *J. Math. Biol.*, vol. 51, no. 5, pp. 508–526, 2005. 664
- 627
- 628
- 629 [25] P. McSharry, M. McGuinness, and A. Fowler, "Confronting a cardiovascular system model with heart rate and blood pressure data," in *Proc. Comput. Cardiol.*, 2005, pp. 587–590. 665
- 630
- 631
- 632 [26] T. Heldt, M. B. Oefinger, M. Hoshiyama, and R. G. Mark, "Circulatory response to passive and active changes in posture," in *Proc. Comput. Cardiol.*, 2003, vol. 30, pp. 263–266. 666
- 633
- 634
- 635 [27] T. Heldt, "Computational models of cardiovascular response to orthostatic stress," Ph.D. dissertation, Massachusetts Inst. Technol., Cambridge, MA, USA, Sep. 2004. 667
- 636
- 637
- Q3 638 [28] A. Camm, M. Malik, J. Bigger, G. Breithardt, S. Cerutti, R. Cohen, P. Coumel, E. Fallen, H. Kennedy, R. Kleiger *et al.*, "Heart rate variability: Standards of measurement, physiological interpretation and clinical use. Task force of the european society of cardiology and the north american society of pacing and electrophysiology," *Circulation*, vol. 93, no. 5, pp. 1043–1065, 1996. 668
- 639
- 640
- 641
- 642
- 643
- Q4 644 [29] A. L. Goldberger, L. A. N. Amaral, L. Glass, J. M. Hausdorff, P. C. Ivanov, R. G. Mark, J. E. Mietus, G. B. Moody, C.-K. Peng, and H. E. Stanley. (2000 Jun. 13) PhysioBank, PhysioToolkit, and PhysioNet: Components of a new research resource for complex physiologic signals. *Circulation*. [Online]. 101(23), pp. e215–e220. Available: <http://circ.ahajournals.org/cgi/content/full/101/23/e215> PMID:1085218; doi: 10.1161/01.CIR.101.23.e215 669
- 645
- 646
- 647
- 648
- 649
- 650
- 651 [30] G. Schwarz, "Estimating the dimension of a model," *Ann. Statist.*, vol. 6, no. 2, pp. 461–464, 1978. 670
- 652
- 653 [31] E. R. DeLong, D. M. DeLong, and D. L. Clarke-Pearson, "Comparing the areas under two or more correlated receiver operating characteristic curves: A nonparametric approach," *Biometrics*, vol. 44, pp. 837–845, 1988. 671
- 654
- 655
- 656
- 657 [32] Y. Benjamini and Y. Hochberg, "Controlling the false discovery rate: A practical and powerful approach to multiple testing," *J. Roy. Statist. Soc.*, vol. 57, no. 1, pp. 289–300, 1995. 672
- 658
- 659
- 660 [33] W. P. Riordan Jr., P. R. Norris, J. M. Jenkins, and J. A. Morris Jr., "Early loss of heart rate complexity predicts mortality regardless of mechanism, anatomic location, or severity of injury in 2178 trauma patients," *J. Surg. Res.*, vol. 156, no. 2, pp. 283–289, 2009. 673
- 661
- 662
- 663
- 664 [34] R. K. Avent and J. D. Charlton, "A critical review of trend-detection methodologies for biomedical monitoring systems," *Crit. Rev. Biomed. Eng.*, vol. 17, no. 6, pp. 621–659, 1990. 665
- 666
- 667 [35] M. Saeed and R. G. Mark, "A novel method for the efficient retrieval of similar multiparameter physiologic time series using wavelet-based symbolic representations," in *Proc. AMIA Annu. Symp.*, 2006, pp. 679–683. 668
- 669
- 670
- 671 [36] L. H. Lehman, M. Saeed, G. Moody, and R. G. Mark, "Similarity-based searching in multi-parameter time series databases," in *Proc. Comput. Cardiol.*, 2008, pp. 653–656. 672
- 673
- 674 [37] J. A. Quinn, C. K. Williams, and N. McIntosh, "Factorial switching linear dynamical systems applied to physiological condition monitoring," *IEEE Trans. Pattern Anal. Mach. Intell.*, vol. 31, no. 9, pp. 1537–1551, Sep. 2009. 675
- 676
- 677
- 678 [38] N. Gibney, E. Hoste, E. A. Burdmann, T. Bunchman, V. Kher, R. Viswanathan, R. L. Mehta, and C. Ronco, "Timing of initiation and discontinuation of renal replacement therapy in AKI: Unanswered key questions," *Clin. J. Amer. Soc. Nephrol.*, vol. 3, no. 3, pp. 876–880, 2008. 679
- 680
- 681
- 682 [39] A. Mikhno and C. M. Ennett, "Prediction of extubation failure for neonates with respiratory distress syndrome using the MIMIC-II clinical database," in *Proc. IEEE Eng. Med. Biol. Soc.*, 2012, pp. 5094–5097. 683
- 684
- 685 [40] A. E. Wolters, A. J. Slooter, A. W. van der Kooij, and D. van Dijk, "Cognitive impairment after intensive care unit admission: A systematic review," *Intensive Care Med.*, vol. 39, no. 3, pp. 376–386, 2013. 686
- 687
- 688 [41] L. H. Lehman, M. Saeed, W. Long, J. Lee, and R. G. Mark, "Risk stratification of ICU patients using topic models inferred from unstructured progress notes," in *Proc. AMIA Annu. Symp Amer. Med. Informat. Assoc.*, 2012, pp. 505–511. 689
- 690
- 691
- 692 Authors', photographs and biographies not available at the time of publication. 693
- 693

- 695 Q1. Author: Please verify whether the affiliations of all the authors are OK as set.  
696 Q2. Author: Please provide complete page range in Ref. [7].  
697 Q3. Author: Please provide the department name in Ref. [27].  
698 Q4. Author: Please provide names of all the authors in Ref. [28].

IEEE  
Proof