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## Intersectional and Entangled Risks: **An Empirical Analysis of Disasters** and Landfills

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74 Landfills are environmental hazards linked to harms, such as the production of 75 greenhouse gases and the accumulation of toxins in natural and human systems. 76 77 Although environmental justice research has established such unwanted land uses as hazardous waste sites occur in poor communities and communities of color, less is known about the relationship between landfills and gender. As a driver of global climate change, there is also limited research into the relationships among disasters, landfills, and 81 climate-related risks. To fill this gap, the current study uses an intersectional approach 82 83 to theorize and empirically analyze relationships among landfills, disasters, race, class, and gender. We employ negative binomial regression to analyze a unique U.S. dataset of landfill counts, total number of disasters, and socio-demographic characteristics, 86 including the use of two-way interactions among race, sex, and socioeconomic status 87 variables, and number of federally-declared disasters that influence landfill counts. 88 Findings suggest that intersecting axes of social location (specifically gender and race) are not multiplicative when it comes to landfills or the environmental risks they pose, 91 but we argue may be entangled-that is related in non-linear and complicated ways. Using intersectionality theory, we interpret the findings to indicate that women of 93 color are agents of resistance enacting their own forms of power against dominant structural arrangements that produce and maintain environmental injustices. Conclusions and implications for environmental justice, intersectionality, and climate risks are 97 further discussed. 98

Keywords: gender, intersectionality theory, environmental justice, environmental risk, landfills, climate impacts, social disparities in climate

### INTRODUCTION

One underexplored area of great concern are the risks related to disasters and landfills particularly to women and racially minoritized groups. As global climate change leads to more frequent and catastrophic disasters, these disasters generate exponentially more waste than in a typical year. Disasters create a dual risk to residents of both the immediate effects of the disaster itself and the potentially harmful effects of disaster-generated waste. Moreover, such risk is unevenly distributed, with landfills, the eventual home of such waste, disproportionately located in communities of color and poor communities across the U.S. (e.g., Mohai and Saha, 2015a,b; Cannon, 2020). But how do these disparities of climate related impacts affect communities across intersecting axes of social location including along lines of gender, race, and socioeconomic status? To answer this question, the current paper integrates environmental justice and intersectionality theories to

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develop a theoretical framework to analyze the factors that
influence landfill counts. This theoretical integration informs
a quantitative analysis of landfills across the United States by
examining crucial socio-demographic characteristics, disasters,
and other relevant variables garnered from prior research.

Previous scholarship has found disasters have an indirect 120 relationship on communities through waste generation 121 (McKinney et al., 2015) and that landfills themselves pose 122 a risk to neighboring residents through the production of 123 greenhouse gasses and the accumulation of toxins in human 124 and natural systems (Elliott and Frickel, 2013). To build on this 125 prior work, the current paper uses an intersectional approach 126 (e.g., Crenshaw, 1989; Cho et al., 2013; Collins and Bilge, 127 2020)-that interlocking axes of social location exacerbates 128 social inequalities-to quantitatively assess disparities related to 129 disasters and landfills. Building on previous scholarship that uses 130 quantitative methods to advance intersectionality theory of social 131 inequality (McCall, 2005), this paper extends this framework 132 to further our understanding of environmental injustice, in 133 the form of landfills, and the disparate risks related to waste 134 generation from disasters. In doing so, we consider best practices 135 from prior research on the inclusion of measures of gender in 136 our analyses. 137

To accomplish these aims, we first review relevant 138 environmental justice research of waste, the environmental risks 139 it poses, and disasters. Next, we engage with leading research into 140 gender inequality, climate change, and intersectionality theory to 141 inform promising practices of conceptualizing and empirically 142 assessing unique effects of gender related to the environment. 143 Finally, we apply an intersectional environmental justice 144 approach to our empirical analysis of a unique dataset to yield 145 insights into climate-related environmental risks to landfills. 146

#### LITERATURE REVIEW

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#### Landfills, Climate Risks, and Disasters

Environmental justice scholarship has a rich tradition of 151 analyzing the disproportionate siting of environmental hazards, 152 in particular waste sites and their attendant risks, in poor 153 communities and communities of color across the U.S. (e.g., 154 UCC, 1987; Bullard, 1990; Saha and Mohai, 2005; Bullard 155 et al., 2007; Elliott and Frickel, 2013; Taylor, 2014; Mohai 156 and Saha, 2015a,b). To date, much of the environmental 157 justice research into these disparities has focused on federally-158 regulated sites, such as hazardous waste landfills or other facilities 159 recorded in the US EPA's Toxics Release Inventory. However, 160 little research has investigated the effects of other landfills-161 construction and demolition (C&D), municipal, and industrial-162 that make up 93% of all landfills in the U.S. (Cannon, 2020). 163 Much of the waste that our modern societies produce are 164 not classified or regulated as hazardous waste though much 165 of it may be hazardous posing significant risks to neighboring 166 communities (EPA (United States Environmental Protection 167 Agency), 2008). Research suggests that such waste sites pose 168 risks to human and community health and well-being, such 169 as poor air quality, water contamination, and other noxious 170 outcomes (World Health Organization, 2007; Mattiello et al., 171

172 **Q14** 2013). Moreover, landfills themselves emit greenhouse gases (GHG, i.e., methane) contributing to the acceleration of the 173 global climate emergency and associated-climate risks (EPA, 174 175 Q19 2006, 2013a). While hazardous waste is regulated federally, individual states are responsible for establishing criteria for how 176 these other landfills operate and for monitoring compliance. 177 As understudied sources of climate risks, these landfills are an 178 important site for further investigation (Cannon, 2020). 179

An important gap in the literature are the relationships among 180 climate change-related disasters, landfills, and environmental 181 risks. The ongoing environmental risks, such as increased rates 182 of cancer, poor air quality, increased asthma rates and increased 183 rates of diabetes (Ma et al., 2007; Fazzo et al., 2008; Gensburg 184 et al., 2009; Koshy et al., 2009) posed by both hazardous and 185 non-hazardous landfills (i.e., C&D, industrial, and municipal) 186 may be further exacerbated by disasters, whose frequency and 187 severity are all but certain to increase with the ongoing and 188 unfolding climate crisis (IPCC, 2018). To further understand the 189 complex linkages among landfills, disasters, climate change, and 190 environmental risk disparities, environmental justice scholarship 191 has also begun to investigate how disaster impacts on such waste 192 sites exacerbates these already existing socio-environmental 193 inequalities (McKinney et al., 2015). In their research on this 194 topic, McKinney et al. (2015) found that debris from disasters in 195 more urban areas created a pipeline of waste to more rural areas 196 across the Southeastern U.S. 197

Moreover, understanding the effects of disaster-generated 198 waste is a necessary site of study given the importance of its quick 199 and thorough removal post-disaster to a community's successful 200 recovery (Luther, 2010; Brown et al., 2011). Waste, whether 201 disaster-generated or not, contributes to a climate change-202 related positive feedback loop in which accumulation of waste 203 in landfills increases the production of greenhouse gasses, which 204 contribute to global climate change, which in turn generates more 205 frequent and extreme disasters producing more landfill waste 206 [EPA (United States Environmental Protection Agency), 2013; 207 McKinney et al., 2015]. This positive feedback loop generates 208 climate-related risks through GHG emissions, themselves; risks 209 from increased disaster frequency and severity; and, through the 210 debris produced by disasters. Although there are ongoing efforts 211 to capture methane releases from some landfills (e.g., Spokas 212 et al., 2006; Ghosh et al., 2019), these efforts are patchwork and 213 produce "fugitive emissions" that escape capture (e.g., Mønster 214 et al., 2019). Given these complex relationships, further research 215 is needed to uncover the key linkages among disasters, waste, and 216 socio-demographic characteristics that produce climate change-217 related environmental risk disparities. 218

# Gender Inequality, Environmental Hazards, and Disasters

Another strand of environmental justice research has 223 investigated the disproportionate impact of environmental 224 injustice on women (for an overview of gender and 225 environmental justice research, see Sze, 2017). Increasingly, 226 gender has emerged as a key analytical focus in environmental 227 justice research as scholars increasingly examine how 228

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environmental hazards can affect women differently than 229 men (e.g., Rocheleau et al., 1996; Downey and Hawkins, 230 2008; Collins et al., 2011; Taylor, 2014; Sze, 2017; Perry and 231 Gillespie, 2019). Researchers have also begun to analyze how 232 gender intersects with other aspects of social life including age, 233 immigration status/citizenship, and indigeneity to influence 234 Q19<sub>235</sub> disproportionate exposure to environmental hazards (e.g., Q19 236 Merchant, 1980; Mies and Shiva, 1993; Linder et al., 2008; Collins et al., 2011; McKane et al., 2018; Goodling, 2019; Nirmal and 237 Rocheleau, 2019; McKee, 2020). Research along these lines, 238 using secondary data analysis, have employed female-headed 239 primary households as an indicator of gender inequality, in 240 which they have found gender inequality plays a role in cancer 241 risks associated with air pollution (Collins et al., 2011) and that 242 female-headed primary households tend to be overly represented 243 in U.S. census tracts with high concentrations of air toxics 244 (Downey and Hawkins, 2008). 245

A brief explanation on the use of the category women. 246 I recognize "women" as an expansive category and one that 247 may include non-gender binary, trans-women, intersex and 248 other people who identify as women. Empirically, given the 249 use of female-headed primary households from the American 250 Community Survey as a measure of gender inequality, I 251 acknowledge that whoever identified as female was counted in 252 this measure. In this way, I fail to capture the other ways that 253 people are discriminated by and experience oppression on the 254 basis of gender. I also acknowledge that people may identify 255 as female in the ACS who may not be identified as such on 256 their personal identification documents since the ACS asks for 257 self-reporting. However, for those whom female does not fit 258 their identity they may not be counted in this measure. More 259 work needs to be done to include the full range of identity and 260 expression in census data collection efforts. 261

Gender also plays a role in disaster experiences and recovery. 262 For instance, research suggests that gender inequalities are linked 263 to barriers to access resources and decision making around 264 healthcare services, education, finances, legal apparatuses and 265 rights for women and girls (Sultana, 2010; Ajibade et al., 2013; 266 Kimura and Katano, 2014). Research has also found that women 267 are disproportionately impacted by natural hazards, disasters, 268 and climate change (Enarson and Morrow, 1998; Fordham, 2003; 269 Sultana, 2014; McKinney and Fulkerson, 2015; Gaillard et al., 270 2017). Thus, gender plays a unique role in both environmental 271 hazards and disasters warranting further empirical analysis, while 272 recent scholarship documents that intersectionality can be a 273 powerful theory for understanding such impacts (Vinyeta et al., 274 2016). 275

#### Intersectionality Theory

278 To better assess the ways gender inequality contributes to environmental hazards and intersects with race and 279 socioeconomic status, scholars have also begun to apply 280 intersectionality theory to further understand relationships 281 among gender, race, socioeconomic status, along with other 282 283 socio-demographic characteristics, and environmental outcomes (Ducre, 2018; Malin and Ryder, 2018; Mollet and Faria, 2018; Sultana, 2021). Intersectionality, a critical theory and approach 285

that grows out of Black feminist thought, seeks to identify and 286 critique the ways various systems of power oppress specific kinds 287 of people and how oppressed peoples can enact their own forms 288 of power (e.g., Crenshaw, 1989; Collins, 1990; Cho et al., 2013; 289 290 Q14 Collins and Bilge, 2020). In foundational texts, Crenshaw (1989) and Collins (1990) assert that it is the intersection of interlocking 291 systems of domination vis-à-vis race, class, and gender that 292 oppresses people, in general, and Black, Indigenous, and women 293 of color, in particular. Moreover, intersectionality, in focusing 294 on and valuing difference, seeks to uncover the complex array 295 of identities (e.g., race, socioeconomic status, gender, sexual 296 orientation, religious, etc.) that are produced structurally and 297 culturally and occur simultaneously. Intersectionality research 298 engages with the experiences and meanings of belonging to 299 multiple social groups, examines power and inequality, and 300 understands social categories as fluid, dynamic, and emergent 301 from social contexts (Cole, 2009; Else-Quest and Hyde, 2016a). 302 Intersectional feminist theorists call for the acknowledgment and 303 investigation of structural power to illuminate how it operates 304 to create multiple and wide-ranging disparities (Ducre, 2018; 305 Mollett, 2018; Mollett and Faria, 2018; Nash, 2019). 306

Researchers have begun to use such intersectional approaches 307 to further understanding and analysis of environmental 308 disparities as it relates to disasters (e.g., Luft and Griffin, 2008; 309 Ryder, 2017), urban ecology (e.g., Di Chiro, 2006; Braun, 2015), 310 environmental risks (e.g., Olofsson et al., 2016), pollution (e.g., 311 Sze, 2006), space (e.g., Mollett and Faria, 2018), and climate 312 change (e.g., Kaijser and Kronsell, 2014; Sultana, 2021). Ducre 313 (2018), argues for the inclusion of gender in environmental 314 justice research, which has often been lacking and that it is 315 necessary to cast women experiencing environmental injustices 316 not just as victims but as agents of resistance. 317

Yet, even fewer studies have used intersectionality theory 318 to investigate waste as a form of environmental injustice 319 (Dillon, 2014; Vasudevan, 2019) or used quantitative methods 320 (for exception see McKane et al., 2018) to test hypotheses 321 derived from intersectionality theories applied to environmental 322 justice outcomes. Of the research that exists, Vasudevan (2019) 323 applies critical race theory, postcolonial studies, and feminist 324 geopolitical approaches to understand how racial capitalism as 325 a political and ecological undertaking across scales produces an 326 intimate relationship between race and waste to make metal 327 valuable. Finally, Dillon (2014) advances "waste formations" 328 as a framework for understanding and analyzing socio-329 ecological transformations that produced and reproduced urban 330 inequalities creating new geographies of waste. This dearth of 331 research creates a gap in knowledge of how race, gender, and 332 waste may be related to produce environmental risk disparities. 333

#### Empirical Approaches to Intersectionality

Just as there is a growing body of research bringing together 337 intersectionality theory and environmental issues, increasingly 338 scholars have begun to call for research to employ quantitative 339 approaches to test hypotheses generated from such rich 340 theorizations to further our understanding and analysis of 341 interlocking systems of oppression and power arrangements that

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uphold such systems (McCall, 2005; Cho et al., 2013; Else-343 Quest and Hyde, 2016a). Notably, intersectional research offers 344 generative potential for analytical understanding of complex 345 inequities (Else-Quest and Hyde, 2016a). Although identity 346 categories as represented by quantitative measures are always 347 approximate, by broadening the intersectionality theoretical 348 perspective using more and various types of methods (other than 349 commonly used techniques such as ethnography, deconstruction, 350 and genealogy), scholars can add to how we understand 351 structural inequality and experiences of oppression (McCall, 352 2005). Following other intersectionality scholars (e.g., McCall, 353 Q27 354 2005; Cho et al., 2013; Else-Quest and Hyde, 2016), the term "social category" refers to socially constructed classes of gender, 355 race, socioeconomic status, etc. "Group" is used for groups within 356 a social category (i.e., women, men). "Location" refers to different 357 intersections of groups belonging to multiple intersecting 358 categories (i.e., Black women). This approach also represents an 359 opportunity to extend best practices for the inclusion of gender 360 and racial minority considerations into research using secondary 361 data analysis. Although such methodologies have been critiqued 362 (e.g., Browne and Misra, 2003; Stainback and Tomaskovic-Devey, 363 2009), they can be used strategically to further our knowledge 364 of interlocking systems of oppression and social (e.g., McCall, 365 2005; Grabe et al., 2015; Else-Quest and Hyde, 2016b) and 366 Q19 367 environmental inequality (McKane et al., 2018; Mollett and Faria, 2018). 368 369

Given these interlocking systems of oppression, it would be expected that membership in two disadvantaged social 370 groups (i.e., Black and women) would exacerbate the chances 371 of experiencing more disadvantage (Else-Quest and Hyde, 372 2016a,b). Mostly such insights have been garnered through 373 374 theorizations and qualitative methods (e.g., Collins, 1990; 375 Crenshaw, 1991; Collins, 1999; MacKinnon, 2013; Collins and Bilge, 2020). Applying an interpretive intersectional 376 framework to environmental justice concerns can improve 377 our understanding of environmental disparities for women, 378 generally, and women from racial and ethnic minoritized 379 groups, specifically. 380

An intersectional quantitative approach emphasizes the 381 processes by which categories (e.g., race, socioeconomic status, 382 gender, etc.) are produced and reproduced by structural 383 arrangements of power and are resisted (Cho et al., 2013; Else-384 Quest and Hyde, 2016a; Nash, 2019). Such an approach can help 385 us to explore the nature and extent of configurations of social 386 and environmental inequality. Therefore, primary focus is on 387 the nature of relationships among social groups, the nature and 388 extent of differences and inequalities, and how these relationships 389 may be changing rather than on the definition of a particular 390 391 social group (McCall, 2005, p. 1785). One analytical design 392 suggested by scholars to get at these relationships among social groups is to use interaction effects to test between and within 393 394 group configurations (McCall, 2005; Bauer, 2014; Else-Quest and Hyde, 2016b). 395

To this end, the current study uses a between-group 396 397 design (i.e., gender x race) to test hypotheses derived from intersectionality theories using traditional quantitative 398 approaches (i.e., negative binomial regression) with main 399

and interactive effects. Statistical interactions between two or more categories are one way of considering multiplicative effects (Bauer, 2014; Else-Quest and Hyde, 2016a). Multiplicative effects begin to get at the premise articulated by intersectionality theory 403 that inequality is more than the sum of its constitutive parts (e.g., 404 Crenshaw, 1989; Collins, 1990). Importantly, as Cole (2009) warns interactive effects are not sufficient for an analysis to be intersectional, however using an intersectional framework to interpret such analyses offers insights into the production of injustices across social groups (e.g., McCall, 2005; Ducre, 2018; Mollett and Faria, 2018).

#### Statement of Purpose and Hypotheses

The current study aims to build on the research reviewed 413 above to examine empirically intersectional effects of race, 414 socioeconomic status, gender, and disasters on landfills, a 415 measure of environmental hazards, and to test hypotheses 416 derived from environmental justice and intersectionality theories 417 to further understand climate risk disparities. To do this, we 418 investigate the intercategorical complexity of race, gender, and 419 socioeconomic status that contributes to inequality (McCall, 420 2005), broadly, and environmental risk disparities, specifically. 421 Hypotheses are tested to ascertain whether there are complex 422 differences and inequalities between groups with respect to 423 measures of environmental inequality. Lastly, we employ an 424 interpretive intersectional framework to quantitatively analyze 425 environmental outcomes. 426

Based on the extant literature, the following hypotheses are tested:

H1: Counties with greater landfill counts, regardless of landfill type, contain a greater number of disasters compared to counties with fewer landfill counts, holding all other variables constant.

H2: Counties with greater landfill counts, regardless of landfill type, contain greater percentages of non-white residents and lower socioeconomic status compared to counties with fewer landfill counts, holding all other variables constant.

H3: Counties with greater landfill counts, regardless of landfill type, contain greater percentages of female-headed households compared to counties with fewer landfill counts, holding all other variables constant.

H4: There are complex differences within and between social groups with respect to predicting landfill counts.

H4-A: Counties with a greater percentage of non-white 444 residents and low socioeconomic status, as measured by 445 percent of families in poverty, will have multiplicative 446 effects on landfill counts, regardless of landfill type, holding 447 all other variables constant. 448

H4-B: Counties with a greater percentage of non-white 449 residents and high gender inequality, as measured by 450 percent of female-headed primary households, will have 451 multiplicative effects on landfill counts, regardless of 452 landfill type, holding all other variables constant. 453

H4-C: Counties with higher socioeconomic status, 454 as measured by a greater percentage of percent of 455 families in poverty, and gender inequality, measured as 456

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percent of female-headed primary households, will have multiplicative effects on landfill counts, regardless of landfill type, holding all other variables constant.

#### DATA AND METHODS

In this study, we theoretically specify and empirically analyze how disasters, socioeconomic, racial, and gender status contribute to unequal risks to environmental inequalities to evaluate the understudied effects of non-hazardous waste landfills across the U.S. Negative binomial regression was used to estimate models of a unique data set of social and environmental indicators for all counties of the 48 contiguous United States. A combination of data from several sources is necessary for this research. Data are reviewed below, followed by a discussion of the analytic technique.

#### Landfill Data, 2013

Waste generated from households and through industrial, construction and demolition processes must be disposed of somewhere. Location of landfills maintained by the state are recorded by each state's environmental regulatory agency (e.g., Environmental Management, Natural Resources, Environmental Quality, Environment and Natural Resources, etc.). There is a great deal of variance across landfill records with respect to fill size, accepted materials, and address given non-hazardous waste landfills are regulated and maintained at the state level. Moreover, given such variation across state records, data collection took an extensive amount of time (3 years) and necessitated going to each state's environmental agency to build a dataset of landfill by type and county. Collecting data from individual states, although more time-intensive, proved to be more accurate and thorough than using data from national databases, such as 490 the Toxics Release Inventory (EPA, 2021), which often lacked 101 the most current and complete data available. Furthermore, 492 federal databases are not always comparable due to inconsistent 493 data collection procedures across federal agencies. As such, this 494 dataset represents the first of its kind to the author's knowledge. 495 Hazardous waste landfill data since it is regulated and monitored 496 at the federal level, were obtained from the U.S. Environmental 497 Protection Agency. Landfills in the U.S. are typically in use for 50 498 years until they are capped or closed (South Carolina Department 499 of Health Environmental Control, 2021). Because fill size is 500 not recorded across landfills, we are unable to account for the 501 differing sizes of landfills. Similar to much of the environmental 502 justice literature into waste facilities (e.g., Mohai and Saha, 503 2015b), we treat all landfills as a measure of environmental 504 inequality. All landfills that were listed as open in 2013 are used 505 in the dataset for the 3,108 counties for which there were data. 506

#### Disaster Data, 2013 508

Disaster data are taken from the US Federal Emergency 509 Management Agency (FEMA) for each county and include 510 511 all federally declared disasters for the time period 1961-2011 (i.e., tropical storms, hurricanes, floods, tornadoes, earthquakes, 512 fires, freezes, landslides, droughts, volcanoes, blizzards, water 513 shortages, and tsunamis) (Federal Emergency Management 514 **Q14** 515

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Agency, 2013). These data are only if a disaster was declared by the U.S. President as a disaster according to the policies outlined in the Robert T. Stafford Disaster Relief and Emergency 516 Assistance Act (1988). These data do not describe either the 517 spatial or temporal dimensions of a disaster but the type of 518 disaster, year, and county in which it occurred. To determine 519 key relationships between disasters and landfills, we employ a 520 total number of federally declared disasters from 1964-2011 (see 521 McKinney et al., 2015) to test research hypotheses. 522

#### Urban Rural Continuum, 2013

To control for the urban-rural spectrum of counties, the 525 rural/urban continuum coding (RUCC) from 2013 developed 526 by the United States Department of Agriculture (USDA) is 527 used. Nine classification codes designate counties by degree of 528 urbanization and proximity to metro areas, with each county 529 assigned one of the nine codes. This coding scheme allows 530 researchers to use county data to move beyond metro and non-531 metro areas and into finer residential groups, particularly in 532 analyzing trends of non-metro areas. Higher values (i.e., nine) 533 represent more rural counties. Since the RUCC codes occur at 534 the county level of analysis, this study uses the county as the level 535 of analysis. 536

While county-level analyses may have the potential for errors 537 related to aggregation, such geography is necessitated here due 538 to the requirement of socio-demographically detailed C&D, 539 industrial, and municipal landfill data. Moreover, given that 540 no research has examined the effects of these landfills at the 541 subnational scale, such level of aggregation is an important 542 contribution to our understanding of the associations between 543 social inequality and non-hazardous waste landfills. Lastly, 544 counties are administrative units which are often responsible for 545 regulating landfills. As such, they represent an important unit 546 of study. 547

Given the novelty of these data and the use of RUCC data 548 and FEMA disaster declaration summary data, the county level 549 of analysis is an important first step to testing the research 550 hypotheses elaborated above. The unit hazard coincidence 551 method, wherein analysts investigate the relationship between 552 sociodemographic characteristics of a unit (i.e., county) and 553 the occurrence of a hazard in that unit (i.e., landfill), is used 554 in this analysis. Although critiqued (e.g., Mohai and Saha, 555 2006), utilizing the unit hazard coincidence method is an 556 important initial step in analyzing these novel data and may 557 inform risk management policy given the administrative unit 558 is often the level at which such decisions are made. Given 559 the method and level of analysis, this research provides a 560 generalized view of relationships being explored, more fine-561 grained data and analyses is needed to refine the results 562 established here (see Limitations and Future Research for 563 further discussion). 564

#### Sociodemographic Data, 2013

All sociodemographic data comes from the American 567 Community Survey (ACS) 5-year estimate (2009-2013) to 568 include data for all study areas regardless of population size (U.S. 569 Census, 2021). Additionally, total population and population 570 density, or people per square mile (Smith, 2009), are used

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	Mean	SD	Min	Max	VIF
Dependent variable					
Total non-hazardous landfill counts	1.52	2.09	0	18	-
Hazardous landfill counts	0.17	0.56	0	6	-
Independent and control variables					
Population density (sq. mile)	258.39	1,724.93	0.12	69,468.42	1.15
Total population	98,479.18	314,016.51	82	9,818,605	1.3
RUCC	4.99	2.7	1	9	1.57
Race					
% non-white	16.73	16.32	0.78	97.08	3.83
Socioeconomic					
% below poverty	11.97	5.53	0	40.19	2.77
Gender					
Percent female-headed primary households	11.33	4.28	1.68	38.01	6.38

as predictors in order to statistically control for variations in 591 population size while controlling for the physical size of the 592 county. Other independent variables found in the environmental 593 justice literature (e.g., Mohai and Saha, 2015a) include percent 594 non-white as a measure of racial minoritized groups and the 595 socioeconomic variable, percent of families living below the 596 poverty line. To ascertain a unique effect of gender on landfill 597 counts, following research by Downey and Hawkins (2008) 598 and Collins et al. (2011), percent of female-headed primary 599 households is used as a measure of gender inequality. The 600 census defines this measure as "primary families maintained 601 by a female household with no husband present," (U.S. Census, 602 2021). This indicator is often used measure gender inequality in 603 environmental research. Secondary data analysis is limited by the 604 data collected and future data collection efforts should attempt 605 to collect more information on the experiences and structural 606 disadvantages that are uniquely related to gender to more fully 607 account for gender-related disparities. 608

#### 610 Analytical Approach

Negative binomial regression, a generalization of Poisson 611 regression, was estimated to predict landfill counts (i.e., non-612 hazardous waste and hazardous waste) as a function of 613 county-level socioeconomic characteristics, with a focus on 614 associations among race, socioeconomic status, and gender 615 inequality. Interactive effects were included to determine 616 multiplicative effects of key independent variables to empirically 617 test intersectionality theories of intersecting axes of oppression 618 on an environmental justice outcome (i.e., landfill counts 619 by county). These analyses follow a call by scholars to use 620 quantitative analytic strategies (specifically non-additive linear 621 modeling) to investigate whether or not and the extent to which 622 intersectionality can be identified and analyzed at a level of 623 analysis beyond the individual and to complement ongoing 624 qualitative research into intersecting forms of oppression 625 (McCall, 2005; Cho et al., 2013). Given the large sample size (i.e., 626 3,108 counties for which there were data) and this technique's 627

flexibility for handling overdispersion (caused by an excess of "false" 0s), negative binomial regression was an optimal technique to predict landfill counts. 650

Negative binomial regression is preferable to zero-inflated 651 or Poisson regression for several reasons. First, zero-inflated 652 models are used to differentiate between "true" 0s and "false" 653 Os and given the rigorous methods of data collection from 654 state sources, we assume all 0s are true. Second, negative 655 binomial regression is the preferred model because the model 656 is not under dispersed as determined by Pearson chi-square 657 which estimates dispersion and is typically a rare problem. 658 For instance, Pearson's chi-square > 1 indicates data are over 650 dispersed and if it is < 1, data are under dispersed. Furthermore, 660 dispersion parameter ( $\alpha$ ) was estimated from the data using 661 maximum likelihood since they produced better fitting models 662 than setting the dispersion parameter to 1 (Cameron and 663 Trivedi, 1998). Several commonly used measures of goodness-664 of-fit were reported [i.e., the Pearson Chi-Square, Akaike 665 information criteria (AIC), Bayesian information criterion (BIC), 666 and the Likelihood Ratio Chi-Square] and within accepted 667 ranges (Cameron, 2009). Taken together, these tests indicate 668 the likelihood of the outcome happening given the set of 669 independent variables. All models satisfied assumptions of 670 negative binomial regression. Finally, the same models reported 671 here were run using both Poisson regression and zero-inflated 672 regression, the negative binomial regression models were the best 673 fitted models. 674

#### **DISCUSSION OF RESULTS**

Summary statistics are reported in Table 1. A diagnostic678bivariate correlation table is presented in Table 2. Negative679binomial regression results are presented in Table 3.680The distribution of landfills across the U.S. are presented681in Figure 1.682

Results are organized as follows. To test research hypotheses, 683 two models, Models 1 and 2 respectively, were run with the 684

Q28 685 **TABLE 2** Bivariate correlation table of all measures included in regression models.

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	Total non-hazardous waste landfills	1.00							
<sup>89</sup> 2. ⊢	Hazardous waste landfills	0.291**	1.00						
<sup>90</sup> 3. P	Population density	0.009	0.055**	1.00					
<sup>91</sup> 4. T	Total population	0.352**	0.256**	0.333**	1.00				
<sup>92</sup> 5. F	RUCC	-0.211**	-0.256**	-0.178**	-0.335**	1.00			
<sup>93</sup> 6. A	All disasters (1964–2011)	-0.003	0.054**	0.025	0.173**	-0.106**	1.00		
<sup>94</sup> 7. P	Percent non-white	0.169**	0.117**	0.162**	0.203**	-0.17**	-0.027	1.00	
<sup>95</sup> 8. P	Percent families below poverty	-0.021	-0.27	0.006	-0.05**	0.175**	0.008	0.532**	1.00
96 9. P	Percent female-headed primary households	0.133**	0.124**	0.106**	0.133**	-0.24**	0.032	0.84**	0.702** 1.00

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same set of independent variables predicting total non-hazardous waste landfill counts (i.e., C&D, municipal, and industrial) and hazardous waste landfill counts. The variables included in interaction terms (i.e., percent non-white, percent female-headed primary households, percent living below the poverty line) are centered on the averages to improve interpretation of the noninteraction effects (see below Main effects) (Williams, 2015). Main effects of control and explanatory variables are reported first followed by a discussion of interactive effects. Incidence rate ratios for each variable are presented for ease of interpretation.

#### **Control Variables** 712

\*p < 0.05; \*\*p < 0.01.

Control variables, population density and total population, were 713 in the expected direction for both models. Specifically, greater 714 population density and total population are positively associated 715 716 with landfill counts regardless of landfill type (Smith, 2009). For 717 rurality as Model 1 indicates, consistent with previous research (Cannon, 2020), more urban counties are positively associated 718 with a greater number of non-hazardous waste landfills. One 719 driver of this relationship may be the expansive building boom 720 in urban areas given the relative cheap credit available after the 721 2008 Great Recession along with both private investment in 722 development and public priorities to fund housing. Moreover, we 723 would expect a greater number of non-hazardous waste landfills 724 near urban areas given the large populations concentrated in 725 those areas. Model 2 also indicates that a greater number of 726 hazardous waste landfills is positively associated with more 727 urban areas. This finding is surprising given hazardous waste 728 landfills often take waste from oil processes and other energy 729 production activities that tend to occur outside of urban areas 730 (EPA (United States Environmental Protection Agency), 2020). 731 Taken together, with the expected rise in population size in urban 732 areas, we may expect an increase in non-hazardous waste landfills 733 contributing to climate change through carbon emissions 734 and posing additional climate-related risks for an increased 735 population. With the continued dominant fossil fuel energy 736 system, we may similarly expect an increase in hazardous waste 737 landfills, which take mostly waste from extractive and refinement 738 of oil and natural gas processes, with their attendant risks. Below, 739 statistically significant estimates of main effects of independent 740 variables are reported, followed by interactive effects. 741

#### Main Effects

For Model 1, total number of disasters from 1964-2011 is negatively associated with non-hazardous waste landfill counts. This finding does not support Hypothesis 1-that counties with a greater number of non-hazardous waste landfills contain a greater number of federally declared disasters-but does support prior research that investigated the relationship between disasters and landfills in the Southeastern U.S. (McKinney et al., 2015). McKinney et al. (2015) theorized that there was waste displacement from areas that experienced disasters to other areas, extending the impacts and risks of disasters beyond immediately affected areas through waste transfer. This finding may similarly support what Dillon (2014) has referred to as "waste formations" in which waste is displaced to newly "waste-able spaces." These results can help inform how we think about climate-related risks from landfills and disasters and the downstream impacts for communities through such waste formations, waste generation and transfer. In Model 2, the relationship between disasters and hazardous waste landfills was not statistically significant.

#### Conditional Effects

For models that include interactive effects between two continuous variables, non-interaction effects (often referred to as main effects) are typically interpreted to ensure main effects and interactive effects are not confounded (Jaccard et al., 1990; Aiken and West, 1991). In such models, main effects represent conditional relationships-that is the main effect of one variable (i.e., percent families living below the poverty line) is in fact the effect of that variable when the other predictors used in the interaction terms equal zero (Williams, 2015). Put another way, the main effect of a variable on the outcome is conditioned by the other variable in the interaction term. For this reason, we have centered the variables-that is we subtracted the mean from each case—included in the interaction terms. This technique improves interpretability of our analyses such that we can interpret the non-interaction effects of each variable included in an interaction term as its average instead of as zero (Williams, 2015).

The conditional effect of race in Model 1 indicates a positive 795 association with a greater number of non-hazardous waste 796 landfills for counties with average percentages of female-headed 797 primary households and families in poverty. This result supports

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TABLE 3 | Negative binomial regression examining main and interactive effects of socio-demographic variables on counties with counts of non-hazardous and hazardous waste landfills.

	Model 1	Model 2
Independent variables	Non-hazardous	Hazardous
Population density (sq. mile)	1.00*** (5.46 x 10 <sup>-5</sup> )	1.00 (6.65 x 10 <sup>-5</sup>
Total population <sup>a</sup>	1.01*** (0.0012)	1.01*** (0.0024)
Rurality	(0.00.2)	(
RUCC code	0.95*** (0.0113)	0.77*** (0.0328)
Disasters		
Total disasters 1964-2011	0.98*** (0.0045)	0.998 (0.0126)
Socioeconomic		
% below poverty (centered)	0.96*** (0.007)	0.96 <sup>†</sup> (0.208)
Race		
% non-white (centered)	1.01*** (0.0031)	0.995 (0.0076)
Gender		
% female-headed primary households (centered)	1.03* (0.0135)	1.23*** (0.0381)
Interactive effects		
Race*Gender	0.999* (0.0004)	0.998 <sup>†</sup> (0.0012)
Gender*SES	1.00 (0.0014)	0.98***
Race*SES	1.00	1.003 <sup>*</sup> (0.0014)
Constant	2.17* (0.0736)	0.4*** (0.1909)
Goodness-of-fit	(0.000)	()
Likelihood Ratio Chi-Square (df)	433.43(10)***	487.61(10)**
Pearson Chi-Square (Value/df)	1.17	1.05
AIC	9,994.95	2,480.41
BIC N = 3,108	10,067.32	2,552.67

p < 0.05, p < 0.01, p < 0.01, p < 0.001, p < 0.01

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Q13 a Total population is reported in the hundreds of thousands.

Exponentiated coefficients (i.e., incidence rate ratio) are reported with standard errors in parentheses.

Hypothesis 2—that counties with greater non-hazardous waste
landfills contain greater percentages of non-white residents
compared to counties with fewer landfills—and is expected given
prior research into environmental inequalities across other kinds
of waste containment (e.g., Bullard et al., 2007; Smith, 2009; Pais
et al., 2014; Mohai and Saha, 2015a,b).

The conditional effect of socioeconomic status in Model 1 and Model 2 both indicate a negative association with a greater number of non-hazardous waste landfills and hazardous waste landfills, respectively, for counties with average percentages of non-white residents and female-headed primary households. This result, which does not support hypothesis 2, is somewhat

surprising given environmental justice research into toxic waste 856 that indicates it tends to be concentrated in poorer areas in the 857 U.S. (e.g., Downey, 2005; Mohai and Saha, 2015b), but does 858 offer some support for Kosmicki and Long's (Kosmicki and 859 Long, 2016) finding that nuclear power plants tend to occur in 860 census tracts with higher median household incomes than coal 861 hosting tracts in the U.S. This finding suggests the continued 862 need to study different forms of environmental injustices and 863 waste containment to better understand key relationships among 864 environmental and social inequality. 865

The conditional effect of gender in Models 1 and 2 both 866 indicate there is a positive association between a greater number 867 of non-hazardous waste landfills and hazardous waste landfills, 868 respectively, and percent female-headed primary households for 869 counties with average percentages of non-white residents and 870 families in poverty. This result supports hypothesis 3-that 871 counties with greater counts of non-hazardous waste landfills 872 are positively associated with gender inequality-and adds to 873 the growing body of research that uses empirical analyses with 874 indicators of gender inequality to investigate environmental 875 hazards (Downey and Hawkins, 2008; Collins et al., 2011). It also 876 could demonstrate that risks related to the landfills are often born 877 by women. 878

One possible explanation for this finding may be that women, 879 on average, make less money than men (81.6 cents to a man's 880 dollar in the U.S.) (Gharehgozli and Atal, 2020), while accounting 881 for more workers in the U.S. workforce (Kelly, 2020), they 882 subsequently may have less time to mount a political resistance 883 to unwanted land uses such as landfills and their potentially 884 pernicious effects. Research has further shown that women 885 experience time shortages due to the gender wage gap since 886 they have to work more hours or more jobs to increase their 887 wages (e.g., Hyde et al., 2020). Such an explanation advances 888 the path of political least resistance theory in environmental 889 justice research (e.g., Bullard, 1990; Downey, 2005) by including 890 gender disparities. 891

And yet, much environmental justice research has shown the 892 importance of women, generally, and Black, Indigenous, and 893 women of color, in particular, in advancing environmental justice 894 movements (e.g., Bullard, 1993a,b; Taylor, 2000; Cole and Foster, 895 2001; Schlosberg, 2003, 2004; Bell and Braun, 2010; Mann, 2011; 896 McGregor, 2015; Martinez-Alier et al., 2016; Pulido and De Lara, 897 2018; Perkins, 2021). Although, Black, Indigenous, and women of 898 color earn less money and have less time, they are key organizers 899 in fighting environmental injustices in their communities. A key 900 insight of intersectionality theory is that women of color are 901 not just survivors or victims of environmental injustices they 902 also enact resistance to the systems of oppression that perpetrate 903 these injustices (Ducre, 2018). Taking an intersectional approach 904 to environmental justice, it is likely that although women of 905 color experience greater risks from landfills, they also tend to 906 organize resistance for healthier environments, a benefit that 907 extends beyond their immediate communities. This insight is 908 aligned with Ducre's (2018, p. 33) argument regarding Black 909 women's liberation and its possible wide-ranging effects, when 910 she writes "their [Black women] liberation from oppression and 911 rights to clean air and water would ensure that all others would 912



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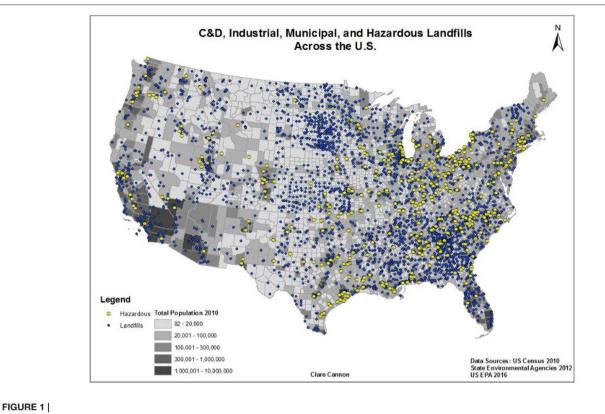
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have access to the same." Similarly, given our analyses suggest there are downstream consequences from disasters on landfills in areas not immediately impacted by such disasters, future research may discover that these sites of resistance also have downstream impacts.

It could also be the case that since women tend to make less money than men, they have less choice in where they can 948 afford to live (e.g., Gotham, 2014). Previous environmental 949 justice research has argued that environmental inequalities are 950 disproportionately located in racial minority neighborhoods 951 because racial housing discrimination constrains residential 952 choices of racial minorities and confines them to neighborhoods 953 with high levels of environmental hazards (Bullard, 1993a,b; 05/ Mohai and Saha, 2006; Pais et al., 2014). Following this research 955 and applying an intersectional interpretive framework, one 956 explanation could be that Black, Indigenous, and women 957 of color's housing choice is further constrained for a 958 multitude of factors both related directly to gender as well 959 as intersecting systems that affect women's residential choices 960 (i.e., discriminatory practices, both overt and covert, related to 961 racial, gender, and socioeconomic statuses, including the gender 962 wage gap). Just as women of color are fighting for healthier 963 environments to reduce risks and environmental injustices, 964 women of color are also leaders in the movement for affordable 965 housing (i.e., Moms 4 Housing). Such insights suggest that just 966 as experiences of injustices are intersectional so is the resistance 967 to them. More research is necessary to understand both why 968 female-headed primary households are disproportionately 969

located in areas with landfills and how women are organizing to fight these injustices. Given the current and existential threat due to the climate crisis, future research should particularly focus on the role of women of color within social mobilizations confronting climate change.

#### Interactive Effects

To test hypotheses elucidated above drawn from intersectionality and environmental justice scholarship, three two-way interaction terms—race x gender, gender x socioeconomic status, and socioeconomic status x race-were included in both models with respective environmental outcomes, non-hazardous and hazardous waste landfill counts (see Table 3). Results of statistically significant interactions are graphed in Figures 1, 2 for Models 1 and 2, respectively, and elaborated below.

Model 1 indicates that as the effect of percent of nonwhite residents of a county increases there is a corresponding increase in the number of non-hazardous waste landfills, but 1016 that effect decreases as percent of female-headed primary 1017 households increases (see Figure 2). Model 2 indicates that as the 1018 percentage of female-headed primary households increases there 1019 is a corresponding increase in the number of hazardous waste 1020 landfills, but that effect decreases as percentage of non-white 1021 residents increases (Figure 3A). Taken together, these findings do 1022 not support hypothesis 4-A that interactions between gender and 1023 race have a multiplicative effect on landfills. 1024

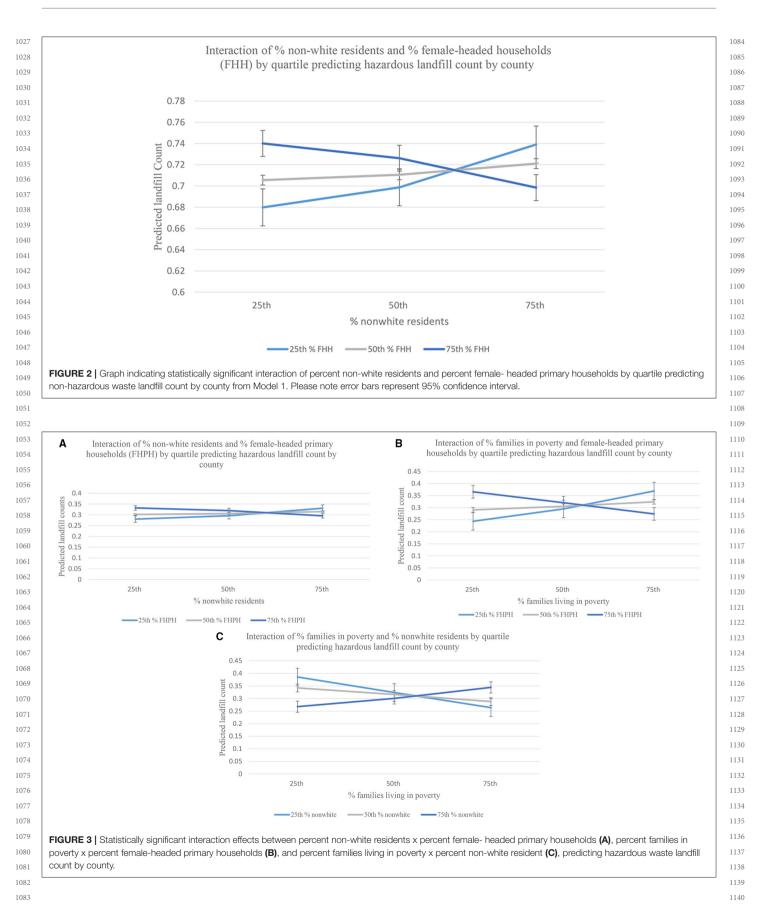
This finding is unexpected given research into both 1025 environmental hazards (Downey and Hawkins, 2008; Collins 1026

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et al., 2011) and intersectionality (McCall, 2005) that would 1141 suggest racial and gender inequality would have multiplicative 1142 effects on increased environmental hazards. This finding does 1143 not support hypothesis 4-B. Given these unexpected findings, 1144 models were re-run multiple times regionally to ascertain if there 1145 were differences across regions. For instance, the same models 1146 were run by US EPA region (i.e., Northeast, Southeast, West, 1147 and Midwest) and the same results held across each region. 1148 Models were also run by RUCC code. Counties were subsetted 1149 by metro, non-metro, and completely rural counties according 1150 to RUCC codes. The findings reported here also held across 1151 these different groupings (Results are available upon request). Considering this and drawing on more recent quantitative intersectional research that has found intersectional risks to 1154 be more than the sum of their parts (i.e., not multiplicative) 1155 (e.g., Evans et al., 2018), we argue that intersectional risks to 1156 environmental injustices may not be multiplicative so much 1157 as entangled. 1158

Entanglement is understood both as a complicated or 1159 compromising relationship and as what Einstein referred to 1160 as "spooky action at a distance"-that two objects affect each 1161 other in hard to predict and non-linear ways over space and 1162 time (e.g., Gilder, 2009). To say the least, socially constructed 1163 categories of race and gender, not to mention data measures 1164 of these constructs (e.g., D'Ignazio and Klein, 2020), are 1165 complicated relationships in themselves and to each other 1166 that may put residents in compromising positions because of 1167 the ways that structural racism and heteropatriarchy play out 1168 across space and time (e.g., formal and informal practices 1169 of redlining). Rather than understanding the multiplicative 1170 ways that race and gender can and do disadvantage some, intersectional theory pushes us to think deeper about the 1172 arrangements of power that constrain people's lives and life 1173 chances. Applying an intersectional theoretical framework, we 1174 suggest that race and gender interlock in entangled ways and not 1175 just multiplicative ones. This may mean that though we would 1176 expect, for example, Black women to be doubly constrained 1177 by structural racism and heteropatriarchy, we also find that 1178 they have access to forms of power and resistance that push 1179 back on these forces and tells a more complicated story than 1180 one solely of multiple forms of oppressions and attendant 1181 injustices (e.g., Ducre, 2018). The non-linear findings suggest 1182 that just as structural arrangements of power that produce and 1183 maintain oppressions are complicated so are arrangements of 1184 resistance. The pressing social and environmental crises due 1185 to climate change are likely to both exacerbate the effects 1186 of such power arrangements as well as inspire opposition to 1187 it. Future research should investigate and test this suggestion 1188 particularly using qualitative methods to gain greater insights 1189 into the lived experiences of Black, Indigenous, and women of 1190 color in both their experiences of environmental injustice and 1191 forms of their resistance to it, especially as it relates to climate 1192 change risks. 1193

<sup>1194</sup> In Model 2, the two-way interaction between gender <sup>1195</sup> inequality and poverty indicates that as the percentage <sup>1196</sup> of female headed primary households increases there is <sup>1197</sup> a corresponding increase in the number of hazardous 1198 waste landfills, but that effect declines as percentage of 1199 families in poverty increases (**Figure 3B**). This result 1200 does not support hypothesis 4-C that gender and 1201 income inequality would have multiplicative effects on 1202 environmental hazards.

Adding to the point above, it may be the case that the 1204 interlocking systems of oppression (i.e., heteropatriarchy and 1205 structural racism) that constrain women's life chances do so 1206 in non-linear ways. For instance, take the gender wage gap 1207 referenced above. When that gap is broken down further by 1208 race and ethnicity, it tells a more specific story that black 1209 women earn 66%, while Indigenous women make 60%, and 1210 Hispanic women earn 58% of a white man's pay for the same 1211 work (Hegewisch and Tesfaselassie, 2018; Asante Muhammad 1212 et al., 2019; Hegewisch and Barsi, 2020). It could be that 1213 this lack of wages not only constrains residential choice 1214 putting women of color in areas with more landfills and 1215 thus in areas with increased environmental risks as posited 1216 above, but also may inform relationships between gender, race, environmental justice, and resistance in ways not yet theorized or 1218 fully understood. 1219

That is to suggest, gender, race, class, and environmental 1220 injustice-here in the form of environmental risks posed by 1221 landfills-are non-linear and may be entangled. For example, given environmental justice research we would expect that as 1223 the percent of non-white residents increases and the percent 1224 of female-headed primary households increases so does the number of landfills. However, we find that this effect decreases 1226 as the number of female-headed primary households increases. One explanation of this result could be that women, and in 1228 particular Black, Indigenous, and women of color, though they 1229 earn less money for the same work, are often the driving force 1230 behind organizing and coalition building to stop unwanted 1231 land uses (i.e., landfills) (e.g., Taylor, 2000; Cole and Foster, 1232 2001; Bullard and Wright, 2012). Using an intersectional 1233 interpretive framework to understand these analyses, we argue 1234 that such findings could evidence sites of resistance and 1235 forms of power by Black, Indigenous, and women of color to 1236 environmental injustices and risks (Mann, 2011; Ducre, 2018; 1237 Perkins, 2021) which may explain the unexpected finding. Such 1238 an interpretation could have implications for climate risks both 1239 in how landfills contribute to global warming through methane 1240 emissions and in how landfills are not "climate ready"-that 1241 is they are not designed to withstand shocks anticipated from 1242 increased disaster frequency and severity due to the global 1243 climate crisis (e.g., Weber et al., 2011; Sinnathamby et al., 2014; 1244 Wille, 2018; Beaven et al., 2020). 1245

Model 2 also shows that as percentage of non-white 1246 residents in a county increases there is a corresponding 1247 increase in the number of hazardous waste landfills in 1248 a county, especially as percentage of families in poverty 1249 increases (Figure 3C). This finding supports Hypothesis 4-1250 A that membership in both a racial minoritized group and 1251 low socioeconomic status group has multiplicative effects on 1252 the likelihood of experiencing environmental inequality. It 1253

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also may signal the complicated relationship among structural 1255 power arrangements in that there is a multiplicative effect of 1256 race and class on presence of environmental risks, confirming 1257 insights from previous environmental justice research (e.g., 1258 Brulle and Pellow, 2006; Mohai and Saha, 2015b; Cannon, 1259 2020) and intersectionality research (e.g., McCall, 2005; Else-1260 Quest and Hyde, 2016b). While belonging to two disadvantaged 1261 social groups exacerbates environmental injustices, there is still 1262 much to learn about the ways that gender, race, class, and 1263 environmental justice intersect to produce and reproduce both 1264 environmental risks and resistance. This finding also highlights 1265 the importance of including gender in theoretically developing 1266 and empirically analyzing environmental outcomes, an often-1267 overlooked dimension in quantitative environmental justice 1268 scholarship, as it tells a more complicated story than one of just 1269 race and class. 1270

To summarize, we find some support for the idea that 1271 intersecting social axes do not result in just multiplicative 1272 effects but that the linkages among race, class, gender, and 1273 waste are more complex. Applying an intersectional framework 1274 to environmental justice theories of waste distribution, such 1275 as the path of least political resistance and constrained 1276 housing choices referenced above, informs us not only of 1277 the complexities that gender brings to socio-environmental 1278 analysis that has thus far predominantly focused on race 1279 and class, but also generates novel insights into structural 1280 arrangements of power. In our intersectional reading, it is not 1281 simply that women of color most likely experience greater 1282 environmental injustices due to structural power arrangements 1283 that make intersecting identities matter, but that they most 1284 likely are actively resisting such injustices. Such insights have 1285 implications for climate-related risks not just from landfills 1286 but also for social mobilizations against major contributors to 1287 climate change (e.g., fossil fuel industries, industrial agriculture, 1288 regulatory and compliance failures). Additional research is 1289 needed to test the explanation presented here and to better 1290 articulate the complex relationships among waste, power, 1291 and resistance. 1292

#### 1294 Limitations and Future Research

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There are limitations to this study that suggest important 1295 areas for future research. First, the current study employs 1296 a cross-sectional analysis of secondary data (for critiques, 1297 see Mohai and Saha, 2015a). As such, this research is 1298 unable to determine causality, future research should include 1299 historic census data, municipality data, and landfill permitting 1300 dates. More data is needed at an even greater detailed level 1301 to focus on at-risk areas for disaster. Such finer grained 1302 data is necessary to more accurately identify and define 1303 key relationships among sociodemographic characteristics, 1304 disaster, and landfills especially for those most vulnerable 1305 to harmful impacts. Future research should consider using 1306 disaster data that include both spatial and temporal indicators 1307 to investigate relationships across time, space, and social 1308 inequality intersectionally. Moreover, future research should 1309 consider investigating the relationship between disaster type and 1310 social inequality. Secondly, the sub-national level scope and 1311

the county scale might affect estimated relationships (Baden et al., 2007). Similar to many geographic units of analysis, 1313 there are limitations in accounting for differences across the 1314 unit of analysis (i.e., counties) (Ringquist, 2005). For example, 1315 there may be more within-county variation for certain variables 1316 (i.e., percentages of non-white residents) than other variables 1317 (i.e., RUCC). Future research should further test the identified 1318 relationships here among sociodemographic characteristics, 1319 disasters, and landfills at a more fine-grained level of analysis 1320 (i.e., using census tracts, at-risk areas for disaster). Additionally, 1321 environmental justice research has shown the effectiveness of 1322 using distance-based methods (i.e., geographically weighted 1323 regression) to ascertain key relationships among race, class, and 1324 hazardous waste sites (Mohai and Saha, 2015a). This research 1325 has demonstrated the importance of including measures of 1326 gender inequality in environmental analyses. However, analyses 1327 are limited by the secondary data collected. Subsequently, 1328 additional data must be collected to more fully measure and 1329 account for gender inequality. Additionally, more research 1330 is needed to address other vulnerable groups to disaster 1331 including the elderly and children. Similarly, intersectional 1332 data should be collected to further test the findings and 1333 theories presented here. Finally, future research should employ 1334 qualitative approaches to further extend intersectionality theory 1335 within environmental justice and to advance our understanding 1336 of underlying mechanisms of general trends presented here 1337 (e.g., Malin and Ryder, 2018). 1338

#### CONCLUSIONS

Quantitative environmental justice research into waste 1343 containment has rarely theoretically developed or analytically 1344 included measures of gender or gender inequality. This paper 1345 advances our understanding of intersectional environmental 1346 risks through an analysis of both non-hazardous and hazardous 1347 waste landfills, disasters, race, class, and gender across the U.S. 1348 In doing so, we highlight the risk disparities related to landfills. 1349 We advance intersectionality studies by empirically testing 1350 research hypotheses derived from intersectionality theory and 1351 applied to environmental outcomes. Doing so we improve 1352 understanding of disparate environmental impacts on women 1353 and racially minoritized groups and reflect on best practices 1354 for including gender and racial minority considerations into 1355 research design and analysis. We build on these insights to 1356 suggest that relationships among gender, race, class, and waste 1357 are entangled-that they are related in non-linear and spooky 1358 ways-and as such reflect the complicated ways that minoritized 1359 women are agents of power not just encountering environmental 1360 injustices but also fighting them. Our theory can inform policies 1361 related to risk management around landfills, as well-strategies 1362 of resistance to environmental injustices, especially in light 1363 of the pressing and existential challenges driven by global 1364 climate change, particularly for marginalized groups. This 1365 theory necessitates further testing and refinement through both 1366 quantitative and qualitative means. In sum, the theoretical 1367 framework and analysis presented here extends environmental 1368

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justice research through a synthesis of intersectionality theory to 1369 develop a theoretical framework of intersectional entanglement 1370 empirically analyze climate-related risk 1371 to in the U.S. 1372

#### DATA AVAILABILITY STATEMENT

data supporting the conclusions of this The raw article will be made available by the authors, without undue reservation.

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The author confirms being the sole contributor of this work and has approved it for publication.

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