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# Measurement of vector boson plus $D^*(2010)^+$ meson production in $\bar{p}p$ collisions at $\sqrt{s} = 1.96$ TeV

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A measurement of vector boson ( $V$ ) production in conjunction with a  $D^*(2010)^+$  meson is presented. Using a data sample corresponding to  $9.7 \text{ fb}^{-1}$  of proton-antiproton collisions at center-of-mass energy  $\sqrt{s} = 1.96 \text{ TeV}$  produced by the Fermilab Tevatron, we reconstruct  $V + D^{*+}$  samples with the CDF II detector. The  $D^{*+}$  is fully reconstructed in the  $D^*(2010)^+ \rightarrow D^0(\rightarrow K^-\pi^+)\pi^+$  decay mode. This technique is sensitive to the associated production of vector boson plus charm or bottom mesons. We measure the ratio of production cross sections  $\sigma(W + D^*)/\sigma(W) = [1.75 \pm 0.13(\text{stat}) \pm 0.09(\text{stat})]\%$  and  $\sigma(Z + D^*)/\sigma(Z) = [1.5 \pm 0.4(\text{stat}) \pm 0.2(\text{stat})]\%$  and perform a differential measurement of  $d\sigma(W + D^*)/dp_T(D^*)$ . Event properties are utilized to determine the fraction of  $V + D^*(2010)^+$  events originating from different production processes. The results are in agreement with the predictions obtained with the PYTHIA program, limiting possible contribution from non-standard-model physics processes.

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In high-energy proton-antiproton ( $p\bar{p}$ ) collisions, production of a vector boson,  $V$ , either a  $W^\pm$  or  $Z^0$  boson, can occur in conjunction with one or more heavy quarks,  $Q$ , where  $Q = c, b$ . Standard model processes that contribute to  $V + Q$  final states include direct (nonresonant) production as well as  $V + Q$  final states originating from decays of top quarks and Higgs bosons. In addition, high mass, non-standard-model states could also contribute, and therefore a detailed study of  $V + Q$  processes could be sensitive to these contributions.

In the standard model, production and decay of  $V + Q$  states is governed by a combination of the strong and weak interactions. For example, the  $p\bar{p} \rightarrow W + c$  process is sensitive to the magnitude of the Cabibbo-Kobayashi-Maskawa matrix element  $V_{cs}$  as well as the strange quark distribution function of the proton [1,2]. Measurements of this process, along with the related processes  $V + b\bar{b}$  and  $V + c\bar{c}$ , can be used in conjunction with our knowledge of the weak interaction in order to test our understanding of the strong interaction at high energy [3]. Study of these processes also helps to provide better modeling of the backgrounds relevant to non-standard-model searches at the Tevatron and LHC.

Previous Tevatron analyses of  $p\bar{p} \rightarrow V + c$  production have measured the absolute and relative cross sections  $\sigma(V + c)/\sigma(V + \text{jets})$  [4–7]. Restrictions on jet acceptance along with the heavy flavor identification techniques limited these analyses to high-momentum ( $p_T(\text{jet}) \gtrsim 20 \text{ GeV}/c$ ) charm and bottom hadrons. In addition, systematic uncertainties arose in the inclusive tagging techniques used to separate light-quark, gluon, and heavy-quark contributions. Similar measurements have been carried out at the LHC [8–10].

Here, we present a complementary technique in which the charm decay is fully reconstructed as a  $D^*(2010)^+$  meson (denoted as  $D^{*+}$ ) through the decay chain  $D^{*+} \rightarrow D^0\pi_s^+$  followed by  $D^0 \rightarrow K^-\pi^+$  [11]. The subscript on  $\pi_s$  is used to denote a “soft” pion with low average momentum compared to the other pion arising from the  $D^0$  decay. This full charm reconstruction provides additional information (charge correlation between decay products, impact parameter) to further classify signal contributions. With this identification technique, an improved measurement range of transverse momentum  $p_T(D^{*+}) > 3 \text{ GeV}/c$  is achieved, with an average value of  $10 \text{ GeV}/c$ . In addition to  $V + c$ , the sample identified with this technique contains contributions from  $V + c\bar{c}$  and charm from the sequential decay of the  $V + b\bar{b}$  process. Full reconstruction of charmed final states in  $W$  boson events has been carried out at the LHC [9].

In reconstructing low-momentum charm hadrons in vector boson events, this work isolates a data sample that has been previously unexplored. To date, anomalous charm hadron production in  $p\bar{p}$  collisions, particularly at low  $p_T(c)$ , has not been ruled out. By isolating and quantifying a unique set of  $V + c$  events, this analysis constitutes a new test of the standard model.

This work utilizes the full data set collected by the CDF II detector at the Fermilab Tevatron collider, with an integrated luminosity of  $9.7 \text{ fb}^{-1}$  for  $p\bar{p}$  collisions at  $\sqrt{s} = 1.96 \text{ TeV}$ . The CDF II detector [12] is a general-purpose magnetic spectrometer surrounded by a projective-tower calorimeter and a muon detector. The CDF II central tracking volume consists of a silicon detector surrounded by a large open-cell drift chamber. The entire tracking volume is contained within a uniform axial magnetic field parallel to the proton beam direction [13].

Events are selected with a three-level online event selection system (trigger). This analysis considers events selected by an inclusive high- $p_T$  lepton trigger requiring an electron (muon) with  $E_T > 18 \text{ GeV}$  ( $p_T > 18 \text{ GeV}/c$ ) [13]. From this high- $p_T$  lepton data set, we select vector-boson events as described below.

A  $Z^0$  candidate is required to have two oppositely-charged, isolated electrons (muons), each with  $E_T > 25 \text{ GeV}$  ( $p_T > 20 \text{ GeV}/c$ ) and  $|\eta| < 1.1$ . A lepton is considered isolated if  $I < 0.1$ , where  $I$  is the isolation, defined as ratio of the total transverse energy in a cone of radius  $\Delta R \equiv \sqrt{\Delta\phi^2 + \Delta\eta^2} = 0.4$  about the lepton track (excluding the lepton) with respect to  $E_T(p_T)$  of the electron (muon). The invariant mass of the lepton pair  $m(\ell^+\ell^-)$  is required to fall between 66 and  $116 \text{ GeV}/c^2$ . Background in the  $Z^0$  candidate sample is estimated by fitting the invariant mass distribution to a double-Gaussian signal plus an exponential background hypothesis. The number of signal events is obtained by integrating the double-Gaussian distribution, yielding  $241450 \pm 580$  ( $257580 \pm 550$ )  $Z^0 \rightarrow ee(\mu\mu)$  events, where the uncertainty is statistical only. The dilepton data and fits are shown in Fig. 1.

A candidate  $W^+$  event is required to have missing transverse energy  $\cancel{E}_T > 25(20) \text{ GeV}$  [14], one isolated electron (muon) with  $E_T > 25 \text{ GeV}$  ( $p_T > 20 \text{ GeV}/c$ ) and pseudorapidity  $|\eta| < 1.1$ . Missing transverse energy is corrected for muons and other instrumental effects that may produce false  $\cancel{E}_T$ . Finally, the transverse mass of the  $W$ ,  $M_T = \sqrt{2E_T\cancel{E}_T(1 - \cos\Delta\phi)}$ , where  $E_T$  is the transverse energy of the lepton and  $\Delta\phi$  is the azimuthal angle between the lepton and missing transverse energy, must satisfy  $M_T > 20 \text{ GeV}/c^2$ .

Background to the  $W$  candidates consists of electroweak sources ( $Z \rightarrow \ell^+\ell^-$  and  $W \rightarrow \tau\nu$  events, denoted as EWK hereafter) and from hadrons displaying a similar topology as  $W$  events (multijet background). We estimate the number of all such background events by first relaxing the selection criteria on lepton  $I$  and event  $\cancel{E}_T$  and then defining four regions A, B, C, and W in the  $\cancel{E}_T/I$  plane (Table I). Region W is the signal region as defined earlier, while the other regions contain primarily multijet events mimicking the  $W$  signature. Lepton isolation versus missing transverse energy is shown in Fig. 2.

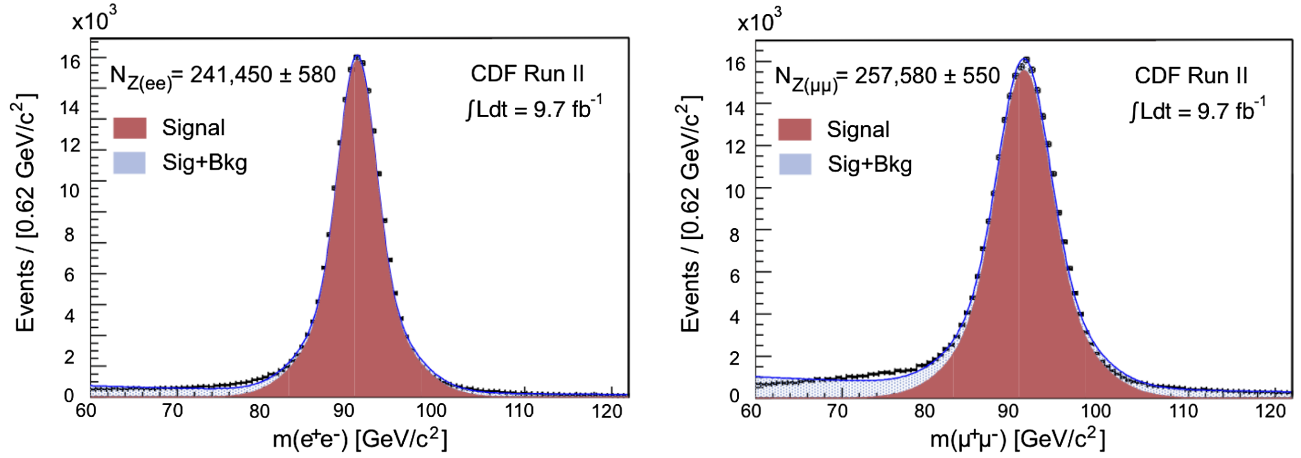


FIG. 1. Distribution of dielectron (left) and dimuon (right) invariant mass. The fit to the data is described in the text.

We estimate  $N_{\text{jet}}^W = N_{\text{jet}}^C \times N_{\text{jet}}^A / N_{\text{jet}}^B$ , where  $N_{\text{jet}}^X$  is the number of multijet events in region  $X$ , as done in Ref. [15]. The contamination arising from EWK and residual signal events in regions A, B, and C is estimated using the PYTHIA 6.2 Monte Carlo simulation [16]. The simulation is used to predict the backgrounds relative to the yield of  $W \rightarrow \ell\nu$  events in signal region W. By relating other contributions to the number of signal events, we produce a system of linear equations to be solved for the number of  $W \rightarrow \ell\nu$  signal events in region W. The solutions determine the contribution of each background in each region. Of 5081938 (5348975) candidate events in region W, 93.6 (91.5)% are  $W \rightarrow e\nu(\mu\nu)$  signal. Uncertainties in these estimates are considered later in the determination of fractions  $N(W + D^{*+})/N(W)$ .

For all  $W$  and  $Z$  candidate events, we search for  $D^{*+}$  mesons by considering all reconstructed charged particles (tracks) within 2.0 cm longitudinally of the point of closest approach of the high- $p_T$  lepton to the beamline ( $|\Delta z| < 2.0$  cm). For each possible set of three tracks with unit total net charge, we hypothesize a match to the  $D^{*+} \rightarrow D^0(\rightarrow K^-\pi^+)\pi_s^+$  decay mode and test for consistency. To ensure well-measured tracks, each track must satisfy  $|\eta| < 1.1$ , and  $p_T(K)$  and  $p_T(\pi) > 400$  MeV/ $c$ ,  $p_T(\pi_s) > 80$  MeV/ $c$ . We also require  $\Delta R < 1.1$  for each pair of tracks. As  $D^0 \rightarrow K^-\pi^+$  is CKM-favored versus  $D^0 \rightarrow K^+\pi^-$  by a factor of  $10^4$ , the two pions are required to have the same charge.

The resulting sample is subjected to a kinematic fit to reconstruct  $D^{*+}$  and  $D^0$  vertices from the  $K$ ,  $\pi$  and  $\pi_s$  track

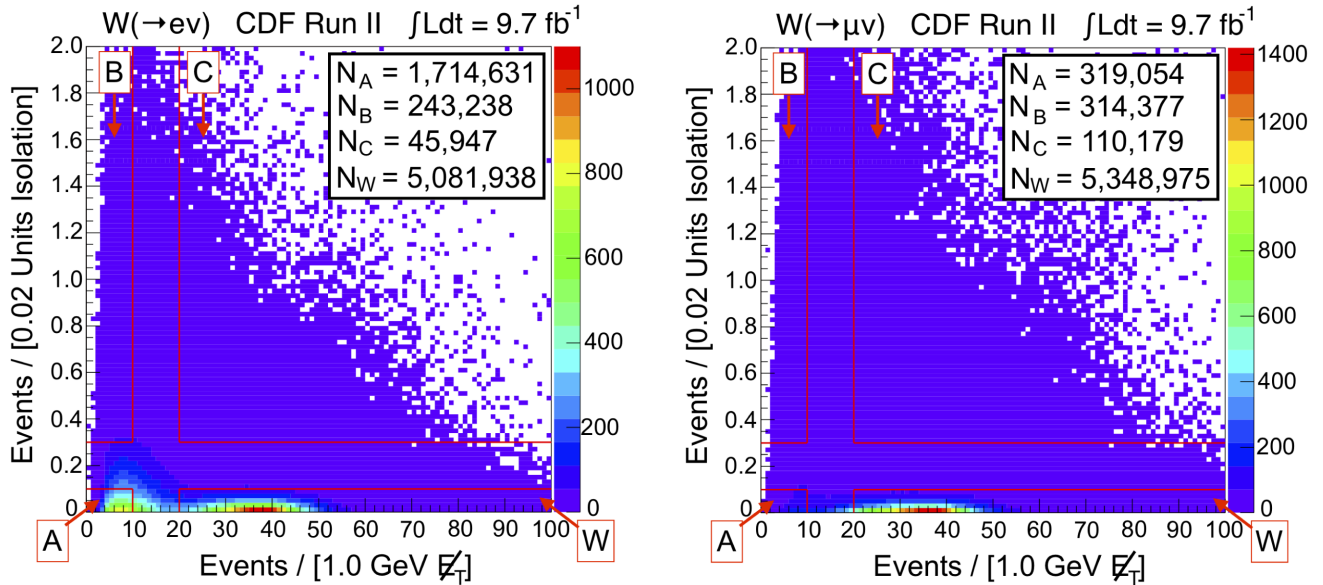


FIG. 2. Lepton isolation versus missing transverse energy for electron (left) and muon (right) events passing the selection criteria described in the text. The regions identified in Table I are shown graphically on the plots. The  $W$  signal is dominant in the region  $I < 0.05$  and  $20 < E_T < 50$  GeV/ $c$ .

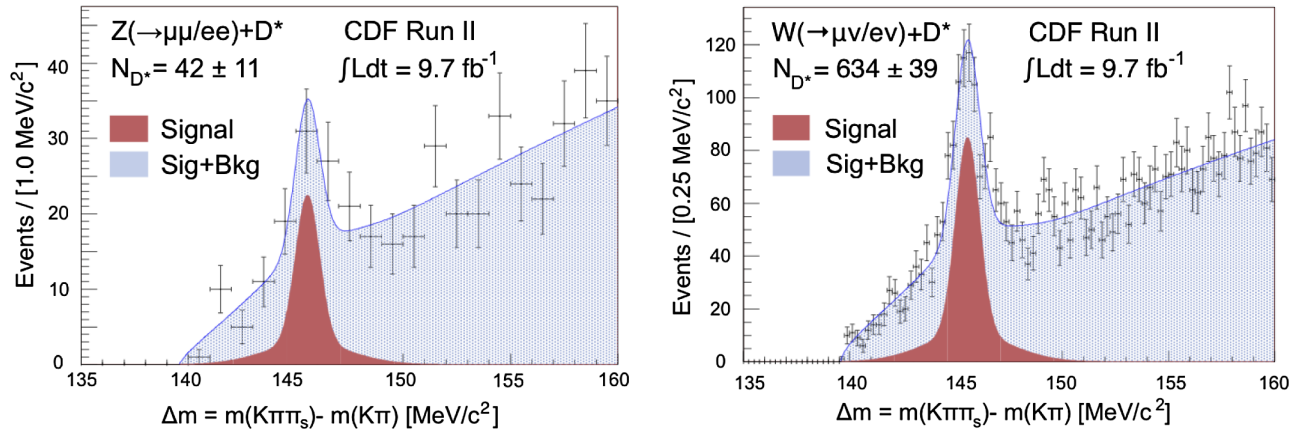


FIG. 3. Distribution of mass difference between the  $D^{*+}$  and  $D^0$  candidates in  $Z$  and  $W$  events, with fit results overlaid. The electron and muon decay modes are combined.

candidates. The  $K$  and  $\pi$  candidates must intersect to form a candidate  $D^0$  vertex, and the  $D^0$  candidate must intersect with the  $\pi_s$  to form a  $D^{*+}$  vertex. Since direct charm (e.g.,  $p\bar{p} \rightarrow Wc$ ,  $p\bar{p} \rightarrow Wc\bar{c}$ ) as well as indirect charm (e.g.,  $Wb\bar{b}$  with  $b \rightarrow c$ ) are of interest, the trajectory of the reconstructed  $D^{*+}$  is not required to point back to the primary  $\bar{p}p$  vertex. Finally, we require the  $D^0$  mass, determined by the fit, to fall within  $3\sigma = 0.033 \text{ GeV}/c^2$  of the known  $D^0$  mass value,  $m_{D^0} = 1.865 \text{ GeV}/c^2$  [17].

To reduce background in the  $D^{*+}$  selection, we train an artificial neural network (ANN) to discriminate among  $D^{*+}$  candidates using 19 variables as described in chapter 5 of Ref. [18]. These variables include charged-particle momenta and opening angles; the distance between the  $p\bar{p}$  collision vertex and the reconstructed  $D^{*+}$  and  $D^0$  vertices; and track impact parameters. To train the ANN, we provide samples of signal and background events. For signal, we generate a sample of  $W + D^{*+}$  events using PYTHIA 6.2 [16] followed by a full detector simulation. For ANN training, we define the signal sample as all simulated  $W + D^{*+}$  candidates with mass difference,  $\Delta m \equiv m(K\pi\pi_s) - m(K\pi)$ , within  $3\sigma = 0.0029 \text{ GeV}/c^2$  of the known mass-difference value  $\Delta m(D^{*+} - D^0) = 0.1455 \text{ GeV}/c^2$  [17]. To model background for ANN training, we choose  $D^{*+}$  candidates in data that pass all fitting and kinematic selection criteria, but whose final-state particles do not have the proper sign correlations (e.g.,  $K^-\pi^-\pi_s^+$ ). To model background events as accurately as possible, we require these background  $D^{*+}$  candidates to fall within  $3.2\sigma$  of the peak, a compromise that allows an equal number of signal and background events for optimal ANN training. Control data samples as well as simulated samples are utilized to verify that the ANN is well behaved and unbiased [18].

We map ANN output scores to the region  $[-1.0, 1.0]$ , with more positive numbers being more signal-like. Candidates are required to have an ANN score greater than 0.0 to be classified as signal. This suppresses 80% of the background while retaining 90% of the signal.

With this algorithm, we search the selected vector-boson-candidate events for  $D^{*+}$  candidates. To measure the  $D^{*+}$  yield, we fit the  $\Delta m$  distribution to a power-law background plus double-Gaussian signal hypothesis. We use a template for the double-Gaussian signal, taken from a fit to simulated signal events, with the signal width increased by a factor of 1.1. This accounts for the slightly better mass resolution observed in the simulation [18].

We obtain the following yields in the  $W$  candidate sample:  $N(W(e\nu) + D^{*+}) = 340 \pm 30$  and  $N(W(\mu\nu) + D^{*+}) = 294 \pm 26$ , where the uncertainties are statistical only. In the  $Z$  candidate sample, we observe  $N(Z(\mu\mu) + D^{*+}) = 22 \pm 9$  and  $N(Z(ee) + D^{*+}) = 20 \pm 7$  events. When the electron and muon decay channels are combined, the yields are  $N(W(\mu\nu/ev) + D^{*+}) = 634 \pm 39$  and  $N(Z(\mu\mu/ee) + D^{*+}) = 42 \pm 11$ . The mass-difference distributions are shown in Fig. 3. Yields reported here are the number of reconstructed  $D^{*+}$  in events passing  $V$  candidate selection criteria, and therefore represent the number of reconstructed  $D^{*+}$  produced in conjunction with both signal and background  $V$  candidates. In the following, we account for the background  $V$  contribution in  $V + D^{*+}$  candidate events.

For  $Z$  candidates, we define two regions in  $\ell^+\ell^-$  mass: The signal region is defined as  $|m_{\ell^+\ell^-} - 91 \text{ GeV}/c^2| \leq 3\sigma$ , and the background region is defined as  $|m_{\ell^+\ell^-} - 91 \text{ GeV}/c^2| > 3\sigma$ , with  $\sigma = 2.0(3.0) \text{ GeV}/c^2$  for  $Z \rightarrow \mu\mu(ee)$  bounded by the  $66 < m_{\ell^+\ell^-} < 116 \text{ GeV}/c^2$  mass window. We fit the  $m_Z$  distribution to a double-Gaussian signal plus exponential background hypothesis, and integrate beneath the curves to estimate the signal and background yields in each region. For each region, we fit the  $\Delta m$  for all  $D^{*+}$  candidates using our previously defined signal plus background fitting function. This provides  $D^{*+}$  yields in the signal and background regions. Two coupled linear equations are used to solve for the rates  $f_{D^{*+}}^{Z,\text{sig}}$  ( $f_{D^{*+}}^{Z,\text{bkg}}$ ) at which signal (background)  $Z$  candidates are produced with  $D^{*+}$  mesons as described in chapter 6 of Ref. [18].

After measuring these rates, the procedure is repeated for various choices of signal and background  $m(\ell^+\ell^-)$  boundaries. We fit the set of all  $f_{D^{*+}}^{Z,\text{sig}}$  values found using these definitions to a constant value hypothesis, and take the variation in this fit to be the systematic uncertainty due to the unknown rate of  $D^{*+}$  mesons produced in association with  $V$  backgrounds. This is done separately for  $Z \rightarrow e^+e^-$  and  $\mu^+\mu^-$  decay channels, giving a systematic uncertainty in  $f_{D^{*+}}^{Z,\text{sig}}$  of 20% for the  $Z \rightarrow \mu\mu$  channel, and of 11% for the  $Z \rightarrow ee$  channel.

To determine the number of real  $V$  plus real  $D^{*+}$  events, we follow a procedure similar to that outlined above to determine the background contributions to the  $W$  sample. We split candidate events into  $E_T/I$  regions A, B, C, and W, as defined in Table I. Each region contains  $D^{*+}$  mesons produced with signal  $W$ , as well as  $D^{*+}$  produced with three background sources:  $W \rightarrow \tau\nu$ ,  $Z \rightarrow \ell^+\ell^-$ , and multijet events. The linear equations are solved simultaneously for the rates  $f_{D^{*+}}^{W,\text{sig}}$  ( $f_{D^{*+}}^{W,\text{bkg}}$ ) at which signal (background)  $W$  candidates are produced with  $D^{*+}$  mesons (see Sec. VI.3 of Ref. [18].) The associated systematic uncertainty is determined by repeating the procedure for several definitions of regions A, B, C and W, first keeping the  $I$  boundaries fixed and varying the  $E_T$  boundaries, and then keeping the  $E_T$  boundaries fixed and varying the  $I$  boundaries. This is done separately for the electron and muon  $W$  decay channels. The resulting systematic uncertainty in  $f_{D^{*+}}^W$  is 10% in the  $W \rightarrow e\nu$  channel, and 2% in the  $W \rightarrow \mu\nu$  channel. The larger uncertainty for the  $W(\rightarrow e\nu)$  case is due to a larger fraction of multijet background, relative to the  $W \rightarrow \mu\nu$  case.

The full results of the sample composition studies are reported in chapter 6 of Ref. [18]. We observe the rate of  $D^{*+}$  mesons produced in events with falsely reconstructed  $V$  bosons to be considerably higher than the rate of  $D^{*+}$  mesons observed in real  $V$  events. This arises through Quantum Chromodynamic (QCD) processes producing falsely reconstructed leptons.

We unfold tagged signal fractions,  $f_{D^{*+}}^V$ , to ratios  $\sigma(V + D^{*+})/\sigma(V)$ , where  $V = (W, Z)$ , by following the prescription

TABLE I. Boundaries used to define regions A, B, C, and W for  $W$  candidates in the  $E_T$ - $I$  plane. Regions A, B, and C consist primarily of multijet background and are used to estimate the multijet background content of signal region W. Numbers in parentheses are for  $W^+ \rightarrow \mu^+\nu$  when different from the  $W^+ \rightarrow e^+\nu$  boundaries.

Region	$E_T$ range	$I$ range
A	$< 10$ GeV	$< 0.1$
B	$< 10$ GeV	$> 0.3$
C	$> 25(20)$ GeV	$> 0.3$
W	$> 25(20)$ GeV	$< 0.1$

$$\frac{\sigma(V + D^{*+})}{\sigma(V)} = \frac{f_{D^{*+}}^V (A \cdot \epsilon)_V}{(A \cdot \epsilon)_{V+D^{*+}} \mathcal{B}(D^{*+} \rightarrow D^0(\rightarrow K\pi)\pi_s)} \quad (1)$$

Here, the branching ratio  $\mathcal{B}(D^{*+} \rightarrow D^0(\rightarrow K\pi)\pi_s)$  is  $0.0263 \pm 0.0004$  [17], and  $(A \cdot \epsilon)_V$ ,  $(A \cdot \epsilon)_{V+D^{*+}}$  is our total (geometric plus kinematic) acceptance for  $V$  and  $V + D^{*+}$  events, respectively.

To determine the  $V$  acceptances, we apply the  $W/Z$  tagging algorithms over inclusive simulated samples generated using PYTHIA 6.2. For the  $V + D^{*+}$  acceptances, we determine the values of  $(A \cdot \epsilon)$  both differentially as a function of  $p_T(D^{*+})$ , and for the inclusive set of all events with  $p_T(D^{*+}) > 3$  GeV, the threshold below which the analysis acceptance vanishes. In the differential case, bins are chosen such that the expected number of tagged  $D^{*+}$  candidates in each bin is approximately constant according to simulation. Bins are as shown in Fig. 4; refer to chapter 8 of Ref. [18] for specific values.

Systematic uncertainties on the production cross sections arise from sample yield and purity estimates, acceptance, efficiency, and simulation [18]. Sources of uncertainty that

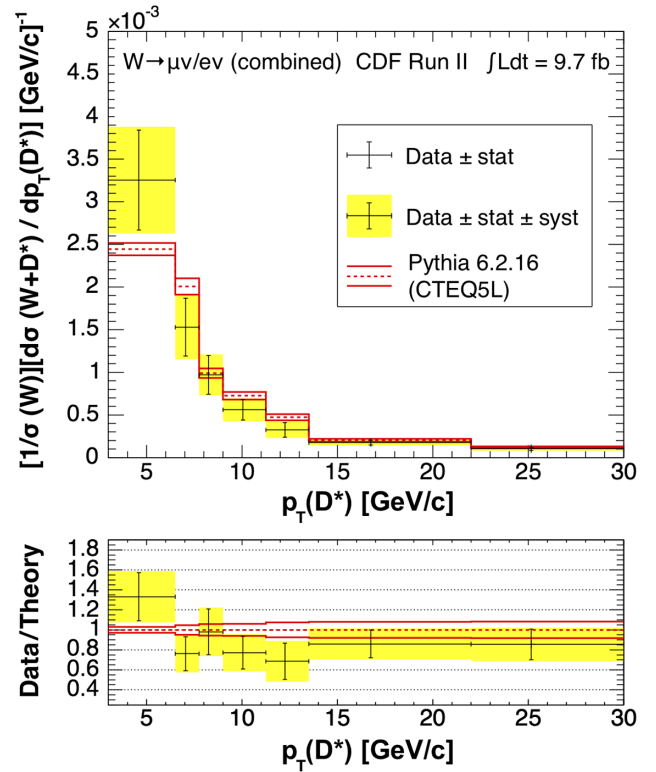


FIG. 4. Ratio of cross sections  $\sigma(W + D^{*+})/\sigma(W)$  as a function of  $p_T(D^{*+})$ , for combined  $W \rightarrow e\nu$  and  $W \rightarrow \mu\nu$  results. Error bars show the statistical uncertainty; the sum in quadrature of the statistical and systematic errors is shown as a yellow error band. The dotted red line shows the prediction of PYTHIA 6.2 obtained using the CTEQ5L PDF, with solid red lines showing PDF uncertainty in this prediction. The ratio of the simulated distribution to data is shown in the lower plot.



TABLE II. Ratio of cross sections  $\sigma(V + D^{*+})/\sigma(V)$  for the inclusive sample  $p_T(D^{*+}) > 3$  GeV/ $c$ , and predictions of PYTHIA 6.2.16 simulation using the CTEQ5L PDF. For results from data, the first uncertainty is statistical, the second is systematic. For results from simulation, uncertainty listed is statistical.

Sample	$\sigma(V + D^{*+})/\sigma(V)$ (in %)	
	Data	PYTHIA 6.2.16
$W \rightarrow e\nu$	$1.74 \pm 0.21 \pm 0.17$	
$W \rightarrow \mu\nu$	$1.75 \pm 0.17 \pm 0.05$	$1.77 \pm 0.07$
$W$ combined	$1.75 \pm 0.13 \pm 0.09$	
$Z^0 \rightarrow ee$	$1.0 \pm 0.6 \pm 0.2$	
$Z^0 \rightarrow \mu\mu$	$1.8 \pm 0.5 \pm 0.2$	$1.36 \pm 0.05$
$Z^0$ combined	$1.5 \pm 0.4 \pm 0.2$	

are common in measurements of total cross section (e.g., luminosity and trigger efficiency uncertainties), cancel when taking the ratio  $\sigma(V + D^{*+})/\sigma(V)$ .

The systematic uncertainty in  $(A \cdot \epsilon)$  is dominated by the uncertainty in the CTEQ [19] parton-distribution function (PDF) of the proton chosen for this analysis. Uncertainties due to the PDFs are estimated using 90% confidence-level (C.L.) variations on the 20 CTEQ eigenvectors. To investigate possible further dependencies on the chosen PDF, the MSTW2008 [20] central value is also checked. When the difference between CTEQ and MSTW2008 is smaller than the uncertainty from variations on the 20 eigenvectors, no additional uncertainty is taken. When the difference is larger, that difference is added in quadrature to the eigenvector uncertainty. Uncertainty in  $\alpha_s(M_Z)$  at 90% C.L. is taken as an additional systematic uncertainty, and added in quadrature to the above.

For both the  $W$  and  $Z$  samples, we find the ratio  $\sigma(V + D^{*+})/\sigma(V)$  for the inclusive set of events  $p_T(D^{*+}) > 3$  GeV/ $c$ . For the  $W$  samples only, we unfold  $f_{D^{*+}}^W$  differentially as a function of  $p_T(D^{*+})$ . Results from the electron and muon decay channels are combined using a best-linear-uncertainty estimator, assuming that systematic uncertainties are fully correlated across  $W$  or  $Z$  decay modes. The final results are shown in Table II and Fig. 4, and in all cases agree with the PYTHIA predictions within uncertainties.

The  $W + D^{*+}$  signal events are expected to come from three different production processes:  $s(d) + g \rightarrow W + c$  ( $Wc$ ),  $q + \bar{q}' \rightarrow W + g(\rightarrow c\bar{c})$  ( $Wcc$ ), and  $q + \bar{q}' \rightarrow W + g(\rightarrow b\bar{b})$  ( $Wbb$ ). Due primarily to a difference in  $p_T(D^{*+})$  spectra, a neural network may generally identify  $D^{*+}$  from different sources ( $Wcc$ ,  $Wc$ , and  $Wbb$ ) with different efficiencies. This enables the fraction of signal from each production process to be determined.

We train two tiers of neural networks; the first consists of networks that are trained to identify one type of signal ( $Wcc$ ,  $Wc$ , and  $Wbb$ ) versus the  $D^{*+}$  background sample described earlier; the second consists of networks that are

trained to identify one type of signal over another ( $Wcc$  vs  $Wbb$ ,  $Wc$  vs  $Wcc$ , and  $Wbb$  vs  $Wc$ ). Each first-tier neural network is paired with a second-tier neural network, in order to select  $D^{*+}$  candidates from one production source preferentially, over all other  $D^{*+}$  candidates as described in chapter 10 of Ref. [18]. We apply each ANN-pair to simulated signal events to determine the efficiencies with which each pair identifies  $D^{*+}$  from each production source, obtaining a matrix of efficiencies. We finally apply each ANN-pair to the data and determine three yields. We solve the resulting linear equations to find fractions  $X^{Wc}$ ,  $X^{Wcc}$ , and  $X^{Wbb}$ , where  $X^Y$  is the fraction of  $W + D^{*+}$  signal that comes from production process  $Y$ . After repeating the technique by varying the efficiencies for each process, we determine that the uncertainty is dominantly statistical.

This method gives good separation for  $X^{Wbb}$ , but is not very effective at determining the fraction  $X^{Wc}$  owing to similarities in the kinematic properties of  $D^{*+}$  from  $Wcc$ , and from  $Wc$ . In the case of  $Wc$  production, conservation of charge requires that the  $W$  and  $D^{*+}$  are produced with opposite signs. In  $Wcc$  or  $Wbb$  production, however, we are equally likely to identify oppositely charged  $W$  and  $D^{*+}$ ,  $(W + D^{*+})_{OS}$ , or same-signed  $W$  and  $D^{*+}$ ,  $(W + D^{*+})_{SS}$ . We therefore determine the number of  $Wc$  events in the  $W + D^{*+}$  signal by taking the difference  $(W + D^{*+})_{OS} - (W + D^{*+})_{SS}$ . We divide by the sum  $(W + D^{*+})_{OS} + (W + D^{*+})_{SS}$  to get the fraction  $X^{Wc}$ . The uncertainty in this measurement of  $X^{Wc}$  is statistical only and treated as uncorrelated with the uncertainty in the measurement of  $X^{Wbb}$ , due to the use of a different technique for finding each fraction. Normalized to unity, the production-process fractions in the overall  $W + D^{*+}$  signal are obtained as  $X^{Wcc} = 73 \pm 8\%$ ,  $X^{Wbb} = 13 \pm 5\%$ ,  $X^{Wc} = 14 \pm 6\%$ . The relative contribution of these processes is momentum dependent. In particular, our ability to identify low-momentum  $D^{*+}$  candidates enhances the  $Wcc$  contribution compared to analyses performed at higher momenta [18].

In conclusion, we present the first measurement of  $V + D^{*+}$  production at hadron colliders in the regime  $p_T(D^{*+}) > 3$  GeV/ $c$ . The expected rate of  $D^{*+}$  production in  $V$  events is as predicted by PYTHIA 6.2, both for the integrated sample  $p_T(D^{*+}) > 3$  GeV/ $c$ , and differentially as a function of  $p_T(D^{*+})$ .

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